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# DEALING WITH HIDDEN STATE SEPARATION IN HIDDEN MARKOV MODELS VIA PENALIZED MAXIMUM LIKELIHOOD ESTIMATION

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# Outline

- 1 The hidden Markov model
- 2 Penalized maximum likelihood approach
- 3 Simulation study
- 4 Application
- 5 Conclusions
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# Hidden Markov model: notation and formulation

- **Univariate binary response variables**  $Y_i^{(t)}$ , observed for unit  $i$  at time occasion  $t$
- **Hidden process**  $U_i^{(t)}$ , following a first-order Markov chain with state-space  $\{1, \dots, k\}$
- **Covariates**  $\mathbf{x}_i^{(t)}$ , representing the vector of observed individual covariates for unit  $i$  at time  $t$
- Composed of two sub-models:
  - 1 **Measurement model**: conditional distribution of the response variable  $Y_i^{(t)}$  given the latent variable  $U_i^{(t)}$  and the possible influence of individual time-varying and time-fixed covariates  $\mathbf{x}_i^{(t)}$
  - 2 **Latent model**: non-parametric distribution of the latent process; accounts for the **unobserved heterogeneity** between individuals

# Hidden Markov model: parameters

- **Initial probabilities:**  $\pi_u = p(U_i^{(1)} = u)$
- **Transition probabilities:**  $\pi_{u|\bar{u}} = p(U_i^{(t)} = u | U_i^{(t-1)} = \bar{u})$
- **Conditional response probabilities:**

$$\phi_{y|ux}^{(t)} = p(Y_i^{(t)} = y | U_i^{(t)} = u, \mathbf{X}_i^{(t)} = \mathbf{x}),$$

based on the following **global logit parameterization**:

$$\log \frac{\phi_{y|ux}^{(t)} + \dots + \phi_{c-1|ux}^{(t)}}{\phi_{0|ux}^{(t)} + \dots + \phi_{y-1|ux}^{(t)}} = \alpha_u + \mathbf{x}'\boldsymbol{\beta},$$

with:

- $\boldsymbol{\alpha} = (\alpha_1, \dots, \alpha_k)$ : support points of the latent variables
- $\boldsymbol{\beta} = (\beta_1, \dots, \beta_p)$ : regression parameters for the covariates

# Maximum likelihood estimation

- The **expectation-maximization** (EM) algorithm is employed to perform **maximum likelihood estimation**
- It maximizes the observed-data log-likelihood function  $\ell(\boldsymbol{\theta})$  relying on the **complete-data log-likelihood** function  $\ell^*(\boldsymbol{\theta})$
- It alternates the following steps until convergence:
  - **E-step: compute the conditional expected value** of  $\ell^*(\boldsymbol{\theta})$  given the value of the parameters at the previous step and the observed data
  - **M-step: update the model parameters** by maximizing the expected value of  $\ell^*(\boldsymbol{\theta})$ :
    - explicit solutions are available for  $\pi_u$  and  $\pi_{u|\bar{u}}$
    - a Newton-Raphson algorithm is used for updating  $\alpha$  and  $\beta$
- Standard errors of the estimates are obtained as minus the second derivative of the expected value of the complete-data log-likelihood

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# Motivation

- **separation problem**: the support points  $\alpha_u$  may be very large and far apart, leading to **widely separated latent states**
- This may result in:
  - **excessively higher relevance of one or more latent states** than others
  - **reduced importance of the available covariates** whose estimated effects may become negligible and insignificant
  - **instability of the estimates** and **large standard errors**

# Penalized maximum likelihood estimation

- Proposed **penalty**: to **reduce latent states separation**:

$$\mathcal{A} = \sum_{u=1}^k (\alpha_u - \bar{\alpha})^2,$$

where  $\bar{\alpha} = \frac{1}{k} \sum_{u=1}^k \alpha_u$

- Applied to** both the **observed-data log-likelihood**  $\ell(\boldsymbol{\theta})$  and the **complete-data log-likelihood**  $\ell^*(\boldsymbol{\theta})$ :

$$\tilde{\ell}(\boldsymbol{\theta}) = \ell(\boldsymbol{\theta}) - \lambda \mathcal{A} \quad \text{and} \quad \tilde{\ell}^*(\boldsymbol{\theta}) = \ell^*(\boldsymbol{\theta}) - \lambda \mathcal{A},$$

where  $\lambda \in \mathbb{R}^+$  is a **tuning parameter** controlling the penalization

- Penalized estimation through the **EM algorithm**:
  - the **E-step** remains unaltered
  - the **M-step** requires to **revise the Newton-Raphson** iteration for  $\alpha$

# Cross-validated log-likelihood

- A **cross validation (CV)** approach is employed to **jointly select** the **penalization parameter**  $\lambda$  and the **number of states**  $k$  of the hidden chain
- We consider  $M$  **partitions of the data**  $D: (D \setminus S_m, S_m)_{m=1, \dots, M}$
- For the  $m$ -th partition:
  - the model is estimated on the data subset  $D \setminus S_m$ , providing parameters estimates  $\hat{\theta}^{(k, \lambda)}(D \setminus S_m)$
  - $\ell(\hat{\theta}^{(k, \lambda)}(D \setminus S_m) \mid S_m)$  denotes the (possibly penalized) log-likelihood function where the model parameters are estimated on the training data  $D \setminus S_m$  but the log-likelihood is evaluated on the test data  $S_m$
  - the **cross-validated likelihood** is defined as

$$\ell_{\text{CV}} = \frac{1}{M} \sum_{m=1}^M \ell(\theta^{(k, \lambda)}(D \setminus S_m) \mid S_m).$$

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# Simulation study

- **Different scenarios** (48) to explore the performance of the proposal with  $k = 3$  hidden states: **sample size** ( $n = 250, 500$ ), **number of time occasions** ( $T = 10, 20$ ), hidden **state persistence** (high or low) and **separation level** (five different behaviors:  $\alpha^j, j = 1, \dots, 6$ )
- **Four covariates** also including the **lagged response variable**; the corresponding vector of regression coefficients is  $\beta = (1, -1, 1, 1)'$
- Extensive **Monte Carlo simulation study**; for each scenario:
  - we randomly draw 50 samples
  - we estimate the HM model using both standard and penalized approaches

# Comparison criteria

- **Percentage variation** in the following quantities for the penalized estimation approach with respect to the standard one:
  - ① **root mean squared relative error between real ( $\beta_j$ ) and estimated ( $\hat{\beta}_{js}$ ) model parameters**, defined as:

$$\text{MSE}(\beta_j) = \frac{1}{50} \sum_{s=1}^{50} (\beta_j - \hat{\beta}_{js})^2, \quad j = 1, 2, 3, 4$$

- ② **standard errors of the covariate regression parameters**
- ③ **computational time**

## Fixed tuning parameter ( $\lambda = 0.01$ ) - MSE

- In the majority of scenarios (40 out of 48), the **penalized estimation method yields lower MSE** values for all regression coefficients  $\beta_j$  (**higher estimation accuracy**)
- Scenarios characterized by highly separated states show the **greatest improvement**; the percentage decreases using penalized estimation often exceeds 90%
- The penalization approach is **less effective** in scenario characterized by **closely spaced states** and in cases with a **high value of  $T$**  (20 time occasions)
- In the **few instances** where the **penalized estimation does not improve** accuracy, the **increase in MSE is extremely low** (typically below 1%)

# Fixed tuning parameter ( $\lambda = 0.01$ ) - Other results

- **Standard errors:**

- In most scenarios, **penalization reduces the estimated standard errors** (greater precision of the parameter estimates)
- The proposed approach is **less effective** in scenarios characterized by **poorly separated latent states**
- In all other cases, the **percentage decrease is significant**, reaching very high values (over **90%** in scenarios with **highly separated** states)

- **Computational time:**

- Estimation with the **penalty** approach often **reduces the average computational time**
- Benefits are **particularly evident** (percentage reduction up to 50%) when **hidden states are widely separated**

# Cross validation

- Settings:
  - 10-fold cross validation
  - Grid search:  $\lambda \in \{0.00, 0.01, \dots, 0.05\}$
- Results:
  - Consistent with the fixed- $\lambda$  analysis, the CV approach yields lower MSE, standard errors, and computational times compared to standard estimation method in most scenarios
  - **High separation** scenarios: CV results are similar to those obtained with  $\lambda$  fixed at 0.01  $\rightarrow$  Penalization, regardless of precise tuning, is sufficient to improve estimation
  - **Low separation** scenarios: CV results outperform those obtained with the fixed penalty approach  $\rightarrow$  Selecting the correct tuning parameter becomes crucial

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# Hypotension during spinal anesthesia

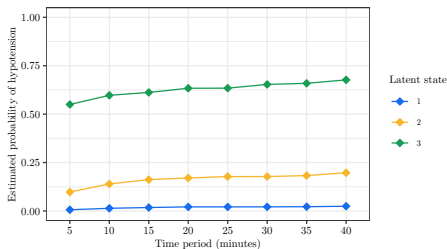
- Data<sup>1</sup> refer to **375 patients** undergoing spinal anesthesia during a surgery; they cover the period **from January 2008 to January 2011**
- Measurements are taken 8 times, at equally spaced intervals over a period of 40 minutes
- Variable  $Y_i^{(t)}$  indicates whether or not patient  $i$  has experienced **hypotension** (decrease in mean systolic blood pressure) at time  $t$
- Approximately **25% ( $n = 94$ ) of patients** recorded **at least one hypotensive status**
- Time-fixed covariates, time-varying covariates, and lagged response
- Using a **CV approach**, we select  **$k = 3$  hidden states** and a **penalization parameter of  $\lambda = 0.02$**  as optimal

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<sup>1</sup>Data are freely available at <https://peerj.com/articles/648/>.

# Estimated conditional hypotension probability

- Patients in the **first hidden state** ( $\hat{\alpha}_1 = -0.821$ ) have an **almost negligible probability of hypotension** during the surgery
- Patients in the **second hidden state** ( $\hat{\alpha}_2 = 3.147$ ) experience a **low probability of hypotension**, ranging approximately from 0.10 to 0.20
- Patients in the **third hidden state** ( $\hat{\alpha}_3 = 7.359$ ) have a **high probability of hypotension** during surgery, ranging from 0.54 to 0.68



# Estimated regression coefficients

Covariate	$\hat{\beta}$	$\hat{se}$	$p$ -value
<b>Gender (Female)</b>	<b>1.526 *</b>	0.752	0.042
Position (Supine)	0.797	0.508	0.116
Operation (Urogoly)	0.374	1.073	0.727
Operation (General surgery)	-0.058	0.897	0.949
EKG (Normal)	-0.858	0.893	0.337
Age (year)	0.036 *	0.019	0.042
DBP	-0.166 **	0.022	0.000
Pulse rate	-0.002	0.011	0.858
Marcain-heavy	-0.049	0.040	0.220
Midazolam	0.242 *	0.115	0.036
Chirocaine	-0.049	0.036	0.180
Fentanyl	1.210	2.350	0.608
Hypotension ( $t - 1$ )	2.678 **	0.011	0.000

- Gender** (female) has a **significant positive effect** on the response variable, indicating that the conditional **probability of experimenting hypotension** given the latent state is **higher for females**

# Estimated regression coefficients

Covariate	$\hat{\beta}$	$\hat{se}$	$p$ -value
Gender (Female)	1.526 *	0.752	0.042
Position (Supine)	0.797	0.508	0.116
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Hypotension ( $t - 1$ )	2.678 **	0.011	0.000

- Older individuals exhibit higher log-odds of being diagnosed with hypotension compared to younger individuals

# Estimated regression coefficients

Covariate	$\hat{\beta}$	$se$	$p$ -value
Gender (Female)	1.526 *	0.752	0.042
Position (Supine)	0.797	0.508	0.116
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Hypotension ( $t - 1$ )	2.678 **	0.011	0.000

- **Diastolic blood pressure** has a significant **negative effect** on the log-odds of hypotension: lower pressure is associated with higher probabilities of experiencing hypotension

# Estimated regression coefficients

Covariate	$\hat{\beta}$	$\hat{se}$	$p$ -value
Gender (Female)	1.526 *	0.752	0.042
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Hypotension ( $t - 1$ )	2.678 **	0.011	0.000

- **Midazolam** has a significant **positive effect**, indicating that higher concentration of this drug in the blood is associated with increased odds of experiencing hypotension during surgery. For the other drugs, the estimated coefficients are not significant

# Estimated regression coefficients

Covariate	$\hat{\beta}$	$\hat{se}$	$p$ -value
Gender (Female)	1.526 *	0.752	0.042
Position (Supine)	0.797	0.508	0.116
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Fentanyl	1.210	2.350	0.608
<b>Hypotension (<math>t - 1</math>)</b>	<b>2.678 **</b>	0.011	0.000

- The **lagged response** has a significant **positive effect on hypotension** indicating serial correlation

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## Limitations and future works

- Implementation of the estimation procedure in **C++** to **improve computational speed**
- **Parallel computation** of the CV approach, distributing the grid search across multiple cores
- **Adoption of genetic algorithms** as an alternative to the exhaustive grid search in CV
- **Evaluation of the model's predictive performance** and comparison with machine learning methods, also in connection with the use of HM models as early warning systems
- **Theoretical investigation** to better understand the circumstances under which latent state separation occurs

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# State separation: a motivating example

- Data generation:
  - one binary response variable for  $n = 250$  units over  $T = 10$  time points
  - four covariates (including the lagged response), with  $\beta = [1, -1, 1, 1]'$
  - $k = 3$  latent states with support points  $\alpha = [-20, -5, 5]'$
  - 1<sup>st</sup> time point: 88 units in the 1<sup>st</sup> state, 71 in the 2<sup>nd</sup>, and 91 in the 3<sup>rd</sup>
- Model estimation using standard maximum likelihood:

	$\hat{\alpha}_1$	$\hat{\alpha}_2$	$\hat{\alpha}_3$	$\hat{\beta}_1$	$\hat{\beta}_2$	$\hat{\beta}_3$	$\hat{\beta}_4$
Estimate	-250.13	146.80	457.96	24.14	-80.21	56.69	103.48
Standard error	-	-	-	23.28	69.43	50.91	90.14
<i>p</i> -value	-	-	-	0.30	0.25	0.27	0.25

- poor accuracy in parameter estimation for both  $\alpha_u$  and  $\beta_j$
- 1<sup>st</sup> time point: 159 units in the 1<sup>st</sup> state and 91 in the 3<sup>rd</sup>
- conditional response probabilities equal to 0 (1<sup>st</sup> state) or 1 (3<sup>rd</sup> state)