

THE TWO-DIMENSIONAL KPZ EQUATION IN THE ENTIRE SUBCRITICAL REGIME

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We consider the KPZ equation in space dimension 2 driven by space-time white noise. We showed in previous work that if the noise is mollified in space on scale ε and its strength is scaled as $\hat{\beta}/\sqrt{|\log \varepsilon|}$, then a transition occurs with explicit critical point $\hat{\beta}_c = \sqrt{2\pi}$. Recently Chatterjee and Dunlap showed that the solution admits subsequential scaling limits as $\varepsilon \downarrow 0$, for sufficiently small $\hat{\beta}$. We prove here that the limit exists in the entire subcritical regime $\hat{\beta} \in (0, \hat{\beta}_c)$ and we identify it as the solution of an additive stochastic heat equation, establishing so-called Edwards–Wilkinson fluctuations. The same result holds for the directed polymer model in random environment in space dimension 2.

1. Introduction and main results. We present first our results for the two-dimensional KPZ equation, and then similar results for its discrete analogue, the directed polymer model in random environment in dimension $2 + 1$. We close the Introduction with an outline of the rest of the paper.

1.1. *KPZ in two dimensions.* The KPZ equation is a stochastic PDE, formally written as

$$(1.1) \quad \partial_t h(t, x) = \frac{1}{2} \Delta h(t, x) + \frac{1}{2} |\nabla h(t, x)|^2 + \beta \xi(t, x), \quad t \geq 0, x \in \mathbb{R}^d,$$

where $\xi(t, x)$ is the space-time white noise, and $\beta > 0$ governs the strength of the noise. It was introduced by Kardar, Parisi and Zhang [31] as a model for random interface growth, and has since been an extremely active area of research for both physicists and mathematicians. The equation is ill-posed due to the singular term $|\nabla h|^2$ which is undefined, because ∇h is expected to be a distribution (generalized function).

In spatial dimension $d = 1$, these difficulties can be bypassed by considering the so-called *Cole–Hopf solution* $h := \log u$, where u is defined as the solution of the multiplicative stochastic heat equation $\partial_t u = \frac{1}{2} \Delta u + \beta \xi u$, which is linear and well-posed in dimension $d = 1$, by classical Itô theory. On large space-time scales, the Cole–Hopf solution exhibits the same fluctuations as many exactly solvable one-dimensional interface growth models, all belonging to the so-called KPZ universality class. See the surveys [14, 42] for reviews on the extensive literature. Few results are known in higher dimensions (see below).

Along a different line, intense research has been carried out in recent years to make sense of the solutions of the KPZ equation and other singular stochastic PDEs. A robust theory was lacking until the seminal work by Hairer [27] and his subsequent theory of regularity structures [28]. Since then, a few alternative approaches have been developed, including the theory of paracontrolled distributions by Gubinelli, Imkeller and Perkowski [26], the theory

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of energy solutions by Gonçalves and Jara [23] and the renormalization approach by Kupiainen [32]. All of these approaches are only applicable to KPZ in space dimension $d = 1$, where the equation is so-called *subcritical*, in the sense that the nonlinearity vanishes in the small scale limit with a scaling that preserves the linear and the noise terms in the equation. In the language of renormalization groups, the KPZ equation in $d = 1$ is *super-renormalizable* (see, e.g., [32]), while regarded as a disordered system, it would be called *disorder relevant* (see, e.g., [22, 29] and [6, 7]).

In this paper, we focus on $d = 2$, which for KPZ is the *critical dimension* (the *renormalizable* or *disorder marginal* case). To define a solution to (1.1), we follow the standard approach and consider a spatially mollified version $\xi^\varepsilon := j_\varepsilon * \xi$ of the noise, where $\varepsilon > 0$, $j \in C_c(\mathbb{R}^2)$ is a probability density on \mathbb{R}^2 with $j(x) = j(-x)$, and $j_\varepsilon(x) := \varepsilon^{-2}j(x/\varepsilon)$. The key question is whether it is possible to replace $\beta\xi$ in (1.1) by $\beta_\varepsilon\xi^\varepsilon - C_\varepsilon$, for suitable constants $\beta_\varepsilon, C_\varepsilon$, such that the corresponding solution h^ε converges to a nontrivial limit as $\varepsilon \downarrow 0$.

It turns out that in space dimension $d = 2$ the right way to tune the noise strength is

$$(1.2) \quad \beta_\varepsilon := \hat{\beta} \sqrt{\frac{2\pi}{\log \varepsilon^{-1}}} \quad \text{for some } \hat{\beta} \in (0, \infty),$$

and to consider the following *mollified KPZ equation* (with $\|j\|_2^2 := \int_{\mathbb{R}^2} j(x)^2 dx$):

$$(1.3) \quad \partial_t h^\varepsilon = \frac{1}{2} \Delta h^\varepsilon + \frac{1}{2} |\nabla h^\varepsilon|^2 + \beta_\varepsilon \xi^\varepsilon - C_\varepsilon \quad \text{where } C_\varepsilon := \beta_\varepsilon^2 \varepsilon^{-2} \|j\|_2^2.$$

For simplicity, we take $h^\varepsilon(0, \cdot) \equiv 0$ as initial datum. If we define

$$(1.4) \quad u^\varepsilon(t, x) := e^{h^\varepsilon(t, x)},$$

then, by Itô’s formula, u^ε solves the *mollified multiplicative Stochastic Heat Equation* (SHE):

$$(1.5) \quad \partial_t u^\varepsilon = \frac{1}{2} \Delta u^\varepsilon + \beta_\varepsilon u^\varepsilon \xi^\varepsilon, \quad u^\varepsilon(0, \cdot) \equiv 1.$$

In [7], we investigated the *finite-dimensional distributions* of the mollified KPZ solution h^ε as $\varepsilon \downarrow 0$. In particular, we discovered in [7], Section 2.3, that there is a transition in the *one-point distribution* as $\hat{\beta}$ varies, with critical value $\hat{\beta}_c := 1$: For any $t > 0$,

$$(1.6) \quad h^\varepsilon(t, x) \xrightarrow{\varepsilon \downarrow 0} \begin{cases} \sigma_{\hat{\beta}} Z - \frac{1}{2} \sigma_{\hat{\beta}}^2 & \text{if } \hat{\beta} < 1, \\ -\infty & \text{if } \hat{\beta} \geq 1 \end{cases} \quad \text{with } \sigma_{\hat{\beta}}^2 := \log \frac{1}{1 - \hat{\beta}^2}, Z \sim N(0, 1).$$

(Note that the limiting distribution does not depend on $t > 0$.) This can be viewed as a *weak disorder to stronger disorder transition*, where we borrow terminology from the directed polymer model (see Section 1.2). It was also shown in [7] that in the subcritical regime $\hat{\beta} < \hat{\beta}_c := 1$ the k -point distribution of h^ε asymptotically factorizes: for any finite set of distinct points $(x_i)_{1 \leq i \leq k}$, the random variables $(h^\varepsilon(t, x_i))_{1 \leq i \leq k}$ converge as $\varepsilon \downarrow 0$ to *independent Gaussians*.

It is natural to investigate the fluctuations of h^ε , regarded as a random field, as $\varepsilon \downarrow 0$. This is what Chatterjee and Dunlap recently addressed in [10]. They actually considered a variant of the mollified KPZ equation (1.3), where β_ε is placed in front of the nonlinearity instead of the noise, namely,

$$(1.7) \quad \partial_t \tilde{h}^\varepsilon = \frac{1}{2} \Delta \tilde{h}^\varepsilon + \frac{1}{2} \beta_\varepsilon |\nabla \tilde{h}^\varepsilon|^2 + \xi^\varepsilon.$$

However, there is a simple relation between \tilde{h}^ε in (1.7) and h^ε in (1.3) (see Appendix A):

$$(1.8) \quad \tilde{h}^\varepsilon(t, x) - \mathbb{E}[\tilde{h}^\varepsilon(t, x)] = \frac{1}{\beta_\varepsilon} (h^\varepsilon(t, x) - \mathbb{E}[h^\varepsilon(t, x)]),$$

therefore, working with \tilde{h}^ε or h^ε is equivalent.

The main result in [10] is that for any fixed $t > 0$, when $\hat{\beta}$ is sufficiently small, the centered solution $\tilde{h}^\varepsilon(t, \cdot) - \mathbb{E}[\tilde{h}^\varepsilon(t, \cdot)]$, viewed as a random distributions on \mathbb{R}^2 , admits nontrivial weak subsequential limits as $\varepsilon \downarrow 0$ (in a negative Hölder space). As a matter of fact, [10] considered the KPZ equation (1.7) on the two-dimensional torus \mathbf{T}^2 , for technical reasons, but it is reasonable to believe that their results should also hold on \mathbb{R}^2 .

The perturbative approach followed by Chatterjee and Dunlap [10] is limited to $\hat{\beta}$ sufficiently small, and it does not prove the existence of a unique limiting random field. Our main result shows that such a limit indeed exists, in the *entire subcritical regime* $\hat{\beta} \in (0, 1)$, and identifies it as the solution of an *additive SHE* with a nontrivial noise strength (that depends explicitly on $\hat{\beta}$). This is commonly called Edwards–Wilkinson fluctuations [18].

THEOREM 1.1 (Edwards–Wilkinson fluctuations for 2-dimensional KPZ). *Let h^ε be the solution of the mollified KPZ equation (1.3), with β_ε as in (1.2) and $\hat{\beta} \in (0, 1)$. Denote*

$$(1.9) \quad \mathfrak{h}^\varepsilon(t, x) := \frac{h^\varepsilon(t, x) - \mathbb{E}[h^\varepsilon(t, x)]}{\beta_\varepsilon} = \frac{\sqrt{\log \varepsilon^{-1}}}{\sqrt{2\pi} \hat{\beta}} (h^\varepsilon(t, x) - \mathbb{E}[h^\varepsilon(t, x)]),$$

where the centering satisfies $\mathbb{E}[h^\varepsilon(t, x)] = -\frac{1}{2}\sigma_{\hat{\beta}}^2 + o(1)$ as $\varepsilon \downarrow 0$; see (1.6).

For any $t > 0$ and $\phi \in C_c(\mathbb{R}^2)$, the following convergence in law holds:

$$(1.10) \quad \langle \mathfrak{h}^\varepsilon(t, \cdot), \phi(\cdot) \rangle = \int_{\mathbb{R}^2} \mathfrak{h}^\varepsilon(t, x) \phi(x) \, dx \xrightarrow[\varepsilon \downarrow 0]{d} \langle v^{(c_{\hat{\beta}})}(t, \cdot), \phi(\cdot) \rangle,$$

where $v^{(c)}(t, x)$ is the solution of the two-dimensional additive stochastic heat equation

$$(1.11) \quad \begin{cases} \partial_t v^{(c)}(t, x) = \frac{1}{2} \Delta v^{(c)}(t, x) + c \xi(t, x), \\ v^{(c)}(0, x) \equiv 0, \end{cases} \quad \text{where } c := c_{\hat{\beta}} := \sqrt{\frac{1}{1 - \hat{\beta}^2}}.$$

REMARK 1.2. For the version (1.7) of KPZ, Chatterjee and Dunlap showed in [10] that any subsequential limit of $\tilde{h}^\varepsilon - \mathbb{E}[\tilde{h}^\varepsilon]$ as $\varepsilon \downarrow 0$ does not coincide with the solution of the additive SHE obtained by simply dropping the nonlinearity $\beta_\varepsilon |\nabla \tilde{h}^\varepsilon|^2$ in (1.7). Here, we show that the limit of $\tilde{h}^\varepsilon - \mathbb{E}[\tilde{h}^\varepsilon]$ actually coincides with the solution of the additive SHE with a *strictly larger noise strength* $c = c_{\hat{\beta}} > 1$. In other words, the nonlinearity in (1.7) produces an independent nonzero noise term in the limit, even though its strength $\beta_\varepsilon \rightarrow 0$.

Our proof of Theorem 1.1 is based on an analogous fluctuation result we proved in [7] for the solution of the SHE (1.5). The independent noise can be seen to arise from the second and higher order chaos expansions of the solution, supported on microscopic scales.

REMARK 1.3. We can view $\mathfrak{h}^\varepsilon(t, \cdot)$ as a *random distribution on \mathbb{R}^2* , that is, a random element of the space of distributions \mathcal{D}' , the dual space of $\mathcal{D} = C_c^\infty(\mathbb{R}^2)$. Our results show that $\mathfrak{h}^\varepsilon(t, \cdot)$ converges in law to $v^{(c_{\hat{\beta}})}(t, \cdot)$ as random distributions. This is because convergence in law on \mathcal{D}' is equivalent to the pointwise convergence of the characteristic functional [20], Th. III.6.5 (see also [4], Corollary 2.4, for an analogue for tempered distributions):

$$\forall \phi \in \mathcal{D} = C_c^\infty(\mathbb{R}^2): \quad \mathbb{E}[e^{i \langle \mathfrak{h}^\varepsilon(t, \cdot), \phi(\cdot) \rangle}] \xrightarrow[\varepsilon \downarrow 0]{} \mathbb{E}[e^{i \langle v^{(c_{\hat{\beta}})}(t, \cdot), \phi(\cdot) \rangle}]$$

and this clearly follows by (1.10).

REMARK 1.4. For simplicity, we only formulated the convergence of $\mathfrak{h}^\varepsilon(t, \cdot)$ to $v^{(c_{\hat{\beta}})}(t, \cdot)$ as a random distribution in space for each fixed t . However, our proof can be easily adapted to prove the convergence of $\mathfrak{h}^\varepsilon(\cdot, \cdot)$ to $v^{(c_{\hat{\beta}})}(\cdot, \cdot)$ as a random distribution in space and time.

REMARK 1.5. The solution $v^{(c)}(t, \cdot)$ of the additive SHE (1.11), also known as the Edwards–Wilkinson equation [18], is the random distribution on \mathbb{R}^2 formally given by

$$(1.12) \quad v^{(c)}(t, x) = c \int_0^t \int_{\mathbb{R}^2} g_{t-s}(x - z) \xi(s, z) \, ds \, dz \quad \text{with } g_t(x) = \frac{1}{2\pi t} e^{-\frac{|x|^2}{2t}}.$$

For any $\phi \in C_c(\mathbb{R}^2)$, we have that $\langle v^{(c)}(t, \cdot), \phi \rangle := \int_{\mathbb{R}^2} v^{(c)}(t, x) \phi(x) \, dx$ is a Gaussian random variable with zero mean and variance

$$(1.13) \quad \text{Var}[\langle v^{(c)}(t, \cdot), \phi \rangle] = c^2 \sigma_\phi^2, \quad \sigma_\phi^2 := \langle \phi, K_t \phi \rangle = \int_{(\mathbb{R}^2)^2} \phi(x) K_t(x, y) \phi(y) \, dx \, dy,$$

where the covariance kernel is given by

$$(1.14) \quad K_t(x, y) := \int_0^t \frac{1}{4\pi u} e^{-\frac{|x-y|^2}{4u}} \, du = \frac{1}{4\pi} \int_{\frac{|x-y|^2}{4t}}^\infty \frac{e^{-z}}{z} \, dz.$$

In [7], we also proved Edwards–Wilkinson fluctuations for the solution u^ε of the 2-dimensional multiplicative SHE (1.5). More precisely, if similarly to (1.9) we set

$$(1.15) \quad u^\varepsilon(t, x) := \frac{1}{\beta_\varepsilon} (u^\varepsilon(t, x) - \mathbb{E}[u^\varepsilon(t, x)]) = \frac{\sqrt{\log \varepsilon^{-1}}}{\sqrt{2\pi} \hat{\beta}} (u^\varepsilon(t, x) - 1),$$

then as $\varepsilon \downarrow 0$ we have the convergence in law $\langle u^\varepsilon(t, \cdot), \phi(\cdot) \rangle \rightarrow \langle v^{(c\hat{\beta})}(t, \cdot), \phi(\cdot) \rangle$ as in (1.10) in the entire subcritical regime $\hat{\beta} \in (0, 1)$; see [7], Theorem 2.17 (which is formulated for space-time fluctuations, but its proof is easily adapted to space fluctuations).

Since $u^\varepsilon(t, x) = \exp(h^\varepsilon(t, x))$, it is tempting to relate (1.15) and (1.9) via Taylor expansion. This is nonobvious, because the one-point distributions of $h^\varepsilon(t, x)$ do not vanish as $\varepsilon \downarrow 0$ (see (1.6)), so we cannot approximate $h^\varepsilon(t, x) \approx u^\varepsilon(t, x) - 1$. We will show in Section 2 that the approximation of $h(t, x)$ is highly nontrivial, and the main contribution actually comes from specific parts of the expansion of $u(t, x)$ which are negligible relative to $u(t, x)$.

For future work, the goal will be to understand the scaling limit of the KPZ solution $h^\varepsilon(t, x)$ at or above the critical point $\hat{\beta}_c = 1$. To our best knowledge, this remains a mystery also for physicists (even the weak to strong disorder transition (1.6) discovered in [7] seems not to have been noted previously in the physics literature). Also the scaling limit of the SHE solution $u^\varepsilon(t, x)$ at or above the critical point is not completely known, even though we recently made some progress at the critical point [8], improving the study initiated in [3] (where the regime (1.2), with $\hat{\beta}$ close to 1, was first studied).

We conclude this subsection with an overview of related results. In space dimension $d = 1$, the Cole–Hopf solution $h(t, x) := \log u(t, x)$ of the KPZ equation (1.1) is well defined as a random function, for any $\beta \in (0, \infty)$, and there is no phase transition in the one-point distribution as β varies. Edwards–Wilkinson fluctuations for $h(t, x)$ and $u(t, x)$ are easily established as $\beta \downarrow 0$, combining Wiener chaos and Taylor expansion (because $u(t, x) \rightarrow 1$).

In space dimensions $d \geq 3$, the right way to scale the disorder strength is $\beta_\varepsilon = \hat{\beta} \varepsilon^{\frac{d-2}{2}}$. It was shown in [12, 39] that the mollified SHE solution $u^\varepsilon(t, x)$ of (1.5) undergoes a weak to stronger disorder transition, similar to the directed polymer model [13]: there is a critical value $\hat{\beta}_c \in (0, \infty)$ such that $u^\varepsilon(t, x)$ converges in law as $\varepsilon \downarrow 0$ to a strictly positive limit when $\hat{\beta} < \hat{\beta}_c$, while it converges to zero if $\hat{\beta} > \hat{\beta}_c$. The KPZ solution $h^\varepsilon(t, x) = \log u^\varepsilon(t, x)$ is thus qualitatively similar to the 2-dimensional case (1.6): $h^\varepsilon(t, x)$ converges in law to a finite limit for $\hat{\beta} < \hat{\beta}_c$, while it converges to $-\infty$ for $\hat{\beta} > \hat{\beta}_c$. The value of $\hat{\beta}_c$ is unknown.

Edwards–Wilkinson fluctuations for the KPZ solution $h^\varepsilon(t, x)$ in dimension $d \geq 3$ have been established recently by Magnen and Unterberger [36], assuming that the noise strength $\hat{\beta}$ is sufficiently small. The corresponding result for the SHE solution $u^\varepsilon(t, x)$ was proved in

[12, 25]. The approaches in these papers do not allow to cover the entire subcritical regime, as we do in dimension 2.

We should also mention that in space dimension $d = 2$, Edwards–Wilkinson fluctuations are believed to hold (and verified in some cases; see, e.g., [43]) also for models in the *anisotropic* KPZ class, where anisotropy means that the term $|\nabla h|^2$ in the KPZ equation (1.1) is replaced by $\langle \nabla h, A \nabla h \rangle$ for some matrix A with $\det(A) \leq 0$.

Shortly after we posted our paper, Dunlap et al. [17] gave an alternative proof (to [36]) of Edwards–Wilkinson fluctuations for the KPZ equation in dimension $d \geq 3$ when $\hat{\beta}$ is sufficiently small. Using the same techniques (Clark–Ocone formula and second-order Poincaré inequality), Gu [24] proved the same Edwards–Wilkinson fluctuation as in our Theorem 1.1 for the KPZ equation in dimension $d = 2$, except his result is restricted to $\hat{\beta}$ small instead of covering the entire subcritical regime.

1.2. The directed polymer model. In this subsection, we state our result for the partition function of the directed polymer model in dimension $2 + 1$. See [11] for an overview of the directed polymer model. In the language of disordered systems, space dimension 2 is critical for this model, where disorder is *marginally relevant*. For further background on the notion of disorder relevance/irrelevance (which corresponds to subcriticality/supercriticality in the context of singular SPDEs); see, for example, [6, 22, 29].

The directed polymer model is defined as a change of measure for a random walk, depending on a random environment (disorder). Let S be the simple symmetric random walk on \mathbb{Z}^2 . If S starts at $x \in \mathbb{Z}^2$, then we denote its law by P_x with expectation E_x , and we omit x when $x = 0$. We set

$$(1.16) \quad q_n(x) := P(S_n = x).$$

Denoting by \tilde{S} an independent copy of S , we define the expected overlap by

$$(1.17) \quad R_N := \sum_{n=1}^N P(S_n = \tilde{S}_n) = \sum_{n=1}^N \sum_{x \in \mathbb{Z}^2} q_n(x)^2 = \sum_{n=1}^N q_{2n}(0) = \frac{\log N}{\pi} + O(1).$$

We fix $\hat{\beta} \in (0, \infty)$ and define $(\beta_N)_{N \in \mathbb{N}}$ by

$$(1.18) \quad \beta_N := \frac{\hat{\beta}}{\sqrt{R_N}} = \frac{\sqrt{\pi} \hat{\beta}}{\sqrt{\log N}} \left(1 + \frac{O(1)}{\log N} \right).$$

Disorder is given by i.i.d. random variables $(\omega(n, x))_{n \in \mathbb{N}, x \in \mathbb{Z}^2}$ with law \mathbb{P} , such that

$$(1.19) \quad \mathbb{E}[\omega] = 0, \quad \mathbb{E}[\omega^2] = 1, \quad \lambda(\beta) := \log \mathbb{E}[e^{\beta \omega}] < \infty \quad \forall \beta > 0 \text{ small enough.}$$

For technical reasons, we require that the law of ω satisfies a concentration inequality. Recall that a function $f : \mathbb{R}^n \rightarrow \mathbb{R}$ is called 1-Lipschitz if $|f(x) - f(y)| \leq |x - y|$ for all $x, y \in \mathbb{R}^n$, with $|\cdot|$ the Euclidean norm. We assume the following:

$$(1.20) \quad \exists \gamma > 1, C_1, C_2 \in (0, \infty) : \text{for all } n \in \mathbb{N} \text{ and } f : \mathbb{R}^n \rightarrow \mathbb{R} \text{ convex and 1-Lipschitz}$$

$$\mathbb{P}(|f(\omega_1, \dots, \omega_N) - M_f| \geq t) \leq C_1 \exp\left(-\frac{t^\gamma}{C_2}\right),$$

where M_f denotes a median of $f(\omega_1, \dots, \omega_N)$. (By changing C_1, C_2 , one can equivalently replace M_f by $\mathbb{E}[f(\omega_1, \dots, \omega_N)]$; see [35], Proposition 1.8.) Condition (1.20) is satisfied if ω is bounded, or if it is Gaussian, or more generally if it has a density $\exp(-V(\cdot) + U(\cdot))$, with V uniformly strictly convex and U bounded. See [35] for more details.

Given $\omega, N \in \mathbb{N}$, and β_N as defined in (1.18), we define the Hamiltonian by

$$(1.21) \quad H_N := \sum_{n=1}^N (\beta_N \omega(n, S_n) - \lambda(\beta_N)) = \sum_{n=1}^N \sum_{y \in \mathbb{Z}^2} (\beta_N \omega(n, y) - \lambda(\beta_N)) \mathbb{1}_{\{S_n=y\}}.$$

We will be interested in the family of partition functions

$$(1.22) \quad \begin{aligned} Z_N(x) &= Z_{N, \beta_N}(x) = \mathbb{E}_x[e^{H_N}], \quad N \in \mathbb{N}, x \in \mathbb{Z}^2, \\ Z_N(x) &:= Z_{\lfloor N \rfloor}(\lfloor x \rfloor), \quad N \in [0, \infty), x \in \mathbb{R}^2. \end{aligned}$$

We will write $Z_N := Z_N(0)$ for simplicity. Note that the law of $Z_N(x)$ does not depend on $x \in \mathbb{Z}^2$, and we have $\mathbb{E}[Z_N(x)] = \mathbb{E}[Z_N] = 1$.

The partition function $Z_N(x)$ is a discrete analogue (modulo a time reversal) of the SHE solution $u^\varepsilon(t, x)$ in (1.5), as can be seen from its Feynman–Kac formula (5.1) below (see also [1]). Then $\log Z_N(x)$ is a discrete analogue of the KPZ solution $h^\varepsilon(t, x)$ in (1.3). In fact, we proved in [7], Theorem 2.8, that for β_N as in (1.18), the random variable $\log Z_N(x)$ converges in distribution to the same limit as in (1.6), with critical value $\hat{\beta}_c = 1$. It is not surprising that here we can also prove the following analogue of Theorem 1.1.

THEOREM 1.6 (Edwards–Wilkinson fluctuations for directed polymer). *Let $Z_{N, \beta_N}(x)$ be the family of partition functions defined as in (1.22), with β_N as in (1.18) with $\hat{\beta} \in (0, 1)$, and the disorder ω satisfying assumptions (1.19) and (1.20). Denote*

$$(1.23) \quad \begin{aligned} \mathfrak{h}_N(t, x) &:= \frac{\log Z_{tN}(x\sqrt{N}) - \mathbb{E}[\log Z_{tN}]}{\beta_N} \\ &= \frac{\sqrt{\log N}}{\sqrt{\pi} \hat{\beta}} (\log Z_{tN}(x\sqrt{N}) - \mathbb{E}[\log Z_{tN}]). \end{aligned}$$

For any $t > 0$ and $\phi \in C_c^\infty(\mathbb{R}^2)$, the following convergence in law holds, with $c_{\hat{\beta}}$ as in (1.11):

$$(1.24) \quad \langle \mathfrak{h}_N(t, \cdot), \phi(\cdot) \rangle = \int_{\mathbb{R}^2} \mathfrak{h}_N(t, x) \phi(x) \, dx \xrightarrow[N \rightarrow \infty]{d} \langle v^{(\sqrt{2}c_{\hat{\beta}})}(t/2, \cdot), \phi(\cdot) \rangle,$$

where $v^{(c)}(s, x)$ is the solution of the two-dimensional additive SHE as in (1.11).

REMARK 1.7. Here, the limit $v^{(\sqrt{2}c_{\hat{\beta}})}(t/2, \cdot)$ differs from $v^{(c_{\hat{\beta}})}(t, \cdot)$ in Theorem 1.1 because the increment of the simple symmetric random walk on \mathbb{Z}^2 has covariance matrix $\frac{1}{2}I$.

We will in fact prove Theorem 1.6 first, since the structure is more transparent in the discrete setting, and then outline the changes needed to prove Theorem 1.1 for KPZ.

1.3. *Outline.* The rest of the paper is organized as follows:

- In Section 2, we present the proof steps and describe the main ideas.
- In Section 3, we give bounds on positive and negative moments for the directed polymer partition function, based on concentration inequalities and hypercontractivity.
- In Section 4, we prove our main result Theorem 1.6 for directed polymer.
- In Section 5, we explain how the proof for directed polymer can be adapted to prove our main result Theorem 1.1 for KPZ.

We will conclude with a few appendices which might be of independent interest, where we prove some results needed in the proofs.

- Appendix A establishes scaling relations for KPZ with different parameters.
- Appendix B recalls and refines known hypercontractivity results for suitable functions (polynomial chaos) of i.i.d. random variables.
- Appendix C formulates a concentration of measure result for the left tail of convex functions that are not globally Lipschitz, defined on general Gaussian spaces.
- Lastly, in Appendix D, we discuss linearity and measurability properties of stochastic integrals, which are needed in the proof in Section 5.

2. Outline of proof steps and main ideas. In this section, we outline the proof steps for Theorems 1.1 and 1.6 and describe the basic setup. We focus on the directed polymer partition function (the case of KPZ follows the same steps). The two main ideas are a *decomposition of the partition function* Z_N which allows us to “linearize” $\log Z_N$ (see Section 2.1), and a representation of Z_N as a *polynomial chaos expansion* in the disorder (see Section 2.2). The “linearization” of $\log Z_N$ essentially reduces Theorem 1.6 to an analogous result for Z_N which we proved in [7], Theorem 2.13.

2.1. *Decomposition and linearization.* Given a subset $\Lambda \subseteq \mathbb{N} \times \mathbb{Z}^2$, we denote by $Z_{\Lambda,\beta}(x)$ the partition function where disorder is only sampled from within Λ , that is,

$$(2.1) \quad Z_{\Lambda,\beta}(x) := E_x[e^{H_{\Lambda,\beta}}] \quad \text{where } H_{\Lambda,\beta} := \sum_{(n,x) \in \Lambda} (\beta\omega_{n,x} - \lambda(\beta)) \mathbb{1}_{\{S_n=x\}}.$$

The original partition function $Z_{N,\beta}(x)$ in (1.21)–(1.22) corresponds to $\Lambda = \{1, \dots, N\} \times \mathbb{Z}^2$.

In our previous study in [7], a key observation was that for $\hat{\beta} \in (0, 1)$ the partition function $Z_{N,\beta_N}(x)$ essentially depends only on disorder in a space-time window around the starting point $(0, x)$ that is negligible on the diffusive scale (N, \sqrt{N}) . This motivates us to approximate $Z_{N,\beta_N}(x)$ by a partition function $Z_{N,\beta_N}^A(x)$ with disorder present only in such a space-time window A_N^x . More precisely, we define a scale parameter a_N tending to zero as

$$(2.2) \quad a_N := \frac{1}{(\log N)^{1-\gamma}} \quad \text{with } \gamma \in (0, \gamma^*),$$

where $\gamma^* > 0$ depends only on $\hat{\beta}$ in Theorem 1.6 and its choice will be clear from the estimate in (4.4) later on. We now introduce the space-time window

$$(2.3) \quad A_N^x := \{(n, z) \in \mathbb{N} \times \mathbb{Z}^2 : n \leq N^{1-a_N}, |z - x| < N^{\frac{1}{2}-\frac{a_N}{4}}\},$$

and define $Z_{N,\beta}^A(x)$ as the partition function which only samples disorder in A_N^x , that is,

$$(2.4) \quad Z_{N,\beta}^A(x) := Z_{\Lambda,\beta}(x) \quad \text{with } \Lambda = A_N^x.$$

We then decompose the original partition function $Z_{N,\beta}(x)$ as follows:

$$(2.5) \quad Z_{N,\beta}(x) = Z_{N,\beta}^A(x) + \hat{Z}_{N,\beta}^A(x),$$

where $\hat{Z}_{N,\beta}^A(x)$, defined by the previous relation, is a “remainder.” In a sense that we will make precise later (see (3.4)), it holds that for any fixed x , $\hat{Z}_{N,\beta_N}^A(x) \ll Z_{N,\beta_N}^A(x)$, and thus

$$(2.6) \quad \begin{aligned} \log Z_{N,\beta_N}(x) &= \log Z_{N,\beta_N}^A(x) + \log \left(1 + \frac{\hat{Z}_{N,\beta_N}^A(x)}{Z_{N,\beta_N}^A(x)} \right) \\ &\approx \log Z_{N,\beta_N}^A(x) + \frac{\hat{Z}_{N,\beta_N}^A(x)}{Z_{N,\beta_N}^A(x)}. \end{aligned}$$

More precisely, if we define the error $O_N(x)$ via

$$(2.7) \quad \log Z_{N,\beta_N}(x) = \log Z_{N,\beta_N}^A(x) + \frac{\hat{Z}_{N,\beta_N}^A(x)}{Z_{N,\beta_N}^A(x)} + O_N(x),$$

then we will show the following.

PROPOSITION 2.1. *Let $O_N(\cdot)$ be defined as in (2.7), then for any $\phi \in C_c(\mathbb{R}^2)$*

$$(2.8) \quad \sqrt{\log N} \frac{1}{N} \sum_{x \in \mathbb{Z}^2} (O_N(x) - \mathbb{E}[O_N(x)]) \phi\left(\frac{x}{\sqrt{N}}\right) \xrightarrow[N \rightarrow \infty]{L^2(\mathbb{P})} 0.$$

Remarkably, even though $\log Z_{N,\beta_N}^A(x)$ gives the dominant contribution to $\log Z_{N,\beta_N}(x)$ for any fixed x , it does not contribute to the fluctuations of $\log Z_{N,\beta_N}(x)$ when averaged over x , that is, we have the following.

PROPOSITION 2.2. *Let $Z_{N,\beta_N}^A(\cdot)$ be defined as in (2.4), then for any $\phi \in C_c(\mathbb{R}^2)$*

$$(2.9) \quad \sqrt{\log N} \frac{1}{N} \sum_{x \in \mathbb{Z}^2} (\log Z_{N,\beta_N}^A(x) - \mathbb{E}[\log Z_{N,\beta_N}^A(x)]) \phi\left(\frac{x}{\sqrt{N}}\right) \xrightarrow[N \rightarrow \infty]{L^2(\mathbb{P})} 0.$$

As a consequence, the fluctuations of $\log Z_{N,\beta_N}(\cdot)$ are determined by the “normalized remainder” $\hat{Z}_{N,\beta_N}^A(\cdot)/Z_{N,\beta_N}^A(\cdot)$. To determine the fluctuations of this term, we define the set

$$(2.10) \quad B_N^{\geq} := ((N^{1-9a_N/40}, N] \cap \mathbb{N}) \times \mathbb{Z}^2,$$

and we let $Z_{N,\beta_N}^{B^{\geq}}(x)$ be the partition function where disorder is sampled only from B_N^{\geq} , that is,

$$(2.11) \quad Z_{N,\beta_N}^{B^{\geq}}(x) := Z_{\Lambda,\beta_N}(x) \quad \text{with } \Lambda = B_N^{\geq}.$$

Note that $\mathbb{E}[Z_{N,\beta_N}^{B^{\geq}}(x)] = 1$, so $(Z_{N,\beta_N}^{B^{\geq}}(x) - 1)$ is a centered random variable. The key point, and the more involved step, will be to show that

$$(2.12) \quad \hat{Z}_{N,\beta_N}^A(x) \approx Z_{N,\beta_N}^A(x) (Z_{N,\beta_N}^{B^{\geq}}(x) - 1),$$

in the following sense.

PROPOSITION 2.3. *Let $Z_{N,\beta_N}^A(\cdot)$, $\hat{Z}_{N,\beta_N}^A(\cdot)$, $Z_{N,\beta_N}^{B^{\geq}}(\cdot)$ be defined as in (2.4), (2.5), (2.11). Then for any $\phi \in C_c(\mathbb{R}^2)$*

$$(2.13) \quad \sqrt{\log N} \frac{1}{N} \sum_{x \in \mathbb{Z}^2} \left(\frac{\hat{Z}_{N,\beta_N}^A(x)}{Z_{N,\beta_N}^A(x)} - (Z_{N,\beta_N}^{B^{\geq}}(x) - 1) \right) \phi\left(\frac{x}{\sqrt{N}}\right) \xrightarrow[N \rightarrow \infty]{L^1(\mathbb{P})} 0.$$

It remains to identify the fluctuations of $Z_{N,\beta_N}^{B^{\geq}}(\cdot)$. This falls within the scope of Theorem 2.13 in [7], which we will recall in Section 4.4 and which will show that the fluctuations of $Z_{N,\beta_N}^{B^{\geq}}(\cdot)$ converges to the solution $v^{(\sqrt{2c\beta})}(1/2, \cdot)$ of the two-dimensional additive SHE, as in Theorem 1.6. The proof is based on polynomial chaos expansions of the partition function, which we will recall in the next subsection.

PROPOSITION 2.4. *Let $Z_{N,\beta_N}^{B\geq}(\cdot)$ be defined as in (2.11). Then*

$$(2.14) \quad \frac{\sqrt{\log N}}{\sqrt{\pi \hat{\beta}}} \frac{1}{N} \sum_{x \in \mathbb{Z}^2} (Z_{N,\beta_N}^{B\geq}(x) - 1) \phi\left(\frac{x}{\sqrt{N}}\right) \xrightarrow[N \rightarrow \infty]{d} \langle v^{(\sqrt{2}c_{\hat{\beta}})}(1/2, \cdot), \phi \rangle,$$

where $v^{(e)}(s, x)$ is the solution of the two-dimensional additive SHE as in (1.11).

Theorem 1.6 is a direct corollary of the decomposition (2.7) and Propositions 2.1–2.4. Regarding the centering, it suffices to note that $\mathbb{E}[\log Z_{N,\beta_N}(x)] = \mathbb{E}[\log Z_{N,\beta_N}^A(x)] + \mathbb{E}[O_N(x)]$, because the random variable $\hat{Z}_{N,\beta_N}^A(x)/Z_{N,\beta_N}^A(x)$ has zero mean, which follows from the polynomial chaos expansions of partition functions that we now present.

2.2. *Polynomial chaos expansions.* Our analysis of the partition functions $Z_{N,\beta_N}^A(\cdot)$, $\hat{Z}_{N,\beta_N}^A(\cdot)$, $Z_{N,\beta_N}^B(\cdot)$ is based on multilinear expansions, known as polynomial chaos expansions, which have also been used extensively in [6, 7].

Recall the definition (1.18) of β_N and our assumptions (1.19) on the disorder, and note that by (1.19) and Taylor expansion, $\lambda(2\beta) - 2\lambda(\beta) \sim \beta^2$ as $\beta \rightarrow 0$. We introduce the sequence

$$(2.15) \quad \sigma_N := \sqrt{e^{\lambda(2\beta_N) - 2\lambda(\beta_N)} - 1} \underset{N \rightarrow \infty}{\sim} \beta_N,$$

where we agree that $a_N \sim b_N$ means $\lim_{N \rightarrow \infty} a_N/b_N = 1$, and we define the random variables

$$(2.16) \quad \xi_{n,x}^{(N)} := \sigma_N^{-1} (e^{\beta_N \omega(n,x) - \lambda(\beta_N)} - 1).$$

We will suppress the dependence of $\xi_{n,x}^{(N)}$ on N , for notational simplicity. Note that $(\xi_{n,x})_{n \in \mathbb{N}, x \in \mathbb{Z}^2}$ are i.i.d. with $\mathbb{E}[\xi_{n,x}] = 0$ and $\mathbb{E}[\xi_{n,x}^2] = 1$.

Recall the definition (1.21)–(1.22) of the partition function $Z_{N,\beta_N}(x)$ of the polymer that starts at time zero from location x . This can be written as

$$(2.17) \quad \begin{aligned} Z_{N,\beta_N}(x) &= \mathbb{E}_x \left[\prod_{1 \leq n \leq N, y \in \mathbb{Z}^2} (1 + \sigma_N \xi_{n,y} \mathbb{1}_{\{S_n=y\}}) \right] \\ &= 1 + \sum_{k=1}^N \sigma_N^k \sum_{\substack{0=n_0 < n_1 < \dots < n_k \leq N \\ x_0=x, x_1, \dots, x_k \in \mathbb{Z}^2}} \prod_{i=1}^k q_{n_i - n_{i-1}}(x_i - x_{i-1}) \xi_{n_i, x_i}, \end{aligned}$$

where $q_n(x) := P(S_n = x)$. Note that the terms in the sum are orthogonal to each other in L^2 , and when $\hat{\beta} \in (0, 1)$ the dominant contribution to $Z_{N,\beta_N}(x)$ comes from disorder $\xi_{\cdot, \cdot}$ in a space-time window that is negligible on the diffusive scale. More precisely, a second moment calculation (see (3.4) below) shows that $Z_{N,\beta_N}(x)$ is close in L^2 to the partition function $Z_{N,\beta_N}^A(x)$ which only samples disorder from within A_N^x (recall (2.3)).

It will be convenient to introduce a concise representation for the expansion (2.17) as follows: given a point (n_0, x_0) and a finite subset $\tau := \{(n_1, x_1), \dots, (n_{|\tau|}, x_{|\tau|})\}$ of $\mathbb{N}_0 \times \mathbb{Z}^2$ with $n_0 < n_1 < \dots < n_{|\tau|}$, we introduce the notation

$$\mathbf{q}^{(n_0, x_0)}(\tau) := \prod_{i=1}^{|\tau|} q_{n_i - n_{i-1}}(x_i - x_{i-1}) \quad \text{and} \quad \boldsymbol{\xi}(\tau) := \prod_{i=1}^{|\tau|} \xi_{n_i, x_i}.$$

For $\tau = \emptyset$, we define $\mathbf{q}^{(n_0, x)}(\tau) = \boldsymbol{\xi}(\tau) := 1$. In this way, we can write concisely the chaos expansion of $Z_{N,\beta_N}(x)$ as

$$(2.18) \quad Z_{N,\beta_N}(x) = \sum_{\tau \subset \{1, \dots, N\} \times \mathbb{Z}^2} \sigma_N^{|\tau|} \mathbf{q}^{(0, x)}(\tau) \boldsymbol{\xi}(\tau).$$

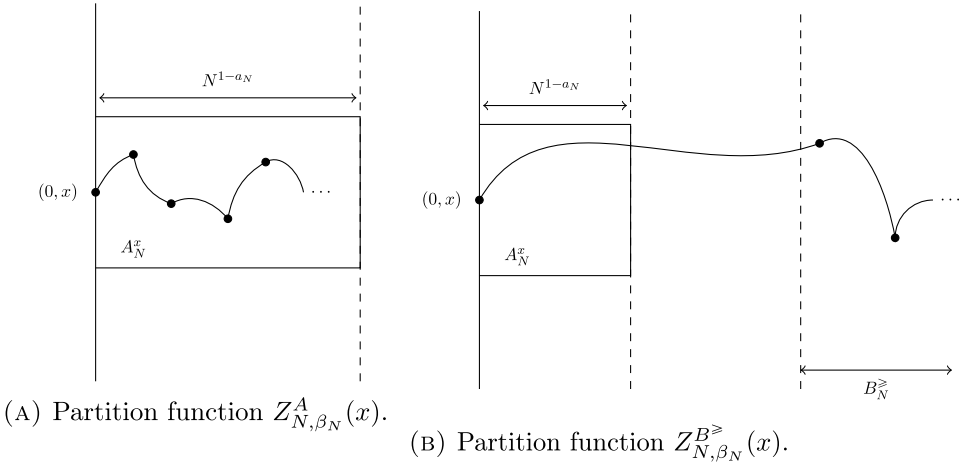


FIG. 1. The above figures depict the chaos expansions of $Z_{N, \beta_N}^A(x)$ and $Z_{N, \beta_N}^{B \geq}(x)$. The disorder sampled by $Z_{N, \beta_N}^A(x)$ is restricted to the set A_N^x , while that of $Z_{N, \beta_N}^{B \geq}(x)$ is restricted to B_N^{\geq} .

Similarly, for the partition functions $Z_{N, \beta}^A(x)$, $Z_{N, \beta}^{B \geq}(x)$ in (2.4), (2.11) we can write

$$(2.19) \quad Z_{N, \beta_N}^A(x) = \sum_{\tau \subset A_N^x} \sigma_N^{|\tau|} \mathbf{q}^{(0,x)}(\tau) \xi(\tau), \quad Z_{N, \beta_N}^{B \geq}(x) = \sum_{\tau \subset B_N^{\geq}} \sigma_N^{|\tau|} \mathbf{q}^{(0,x)}(\tau) \xi(\tau).$$

A graphical illustration of $Z_{N, \beta}^A(x)$ and $Z_{N, \beta}^{B \geq}(x)$ appears in Figure 1.

These polynomial chaos expansions are discrete analogues of Wiener–Itô chaos expansions. They are especially suited for variance calculations and provide important insight. For instance, the partition function $\hat{Z}_{N, \beta}^A(x) := Z_{N, \beta}(x) - Z_{N, \beta}^A(x)$ (see (2.5)) is obtained by restricting the sum in (2.18) to τ which include space-time points outside the set A_N^x , and hence $\hat{Z}_{N, \beta}^A(x)/Z_{N, \beta}^A(x)$ has zero mean due to the independence between the disorder inside and outside A_N^x . Similarly, the centered partition function $(Z_{N, \beta}^{B \geq}(x) - 1)$, which appears in (2.13)–(2.14), is the contribution to (2.19) given by configurations τ that contain only points (and at least one point) in B_N^{\geq} .

3. Moment bounds. In this section, we collect some moment bounds that will be used in the proof.

3.1. *Second moment.* We bound the second moment of $Z_{N, \beta_N}(x)$, $Z_{N, \beta_N}^A(x)$, $\hat{Z}_{N, \beta_N}^A(x)$. We start from $Z_{N, \beta_N}(x)$. It follows by (2.17) and (1.17) that

$$(3.1) \quad \begin{aligned} \mathbb{E}[Z_{N, \beta_N}(x)^2] &= \sum_{\tau \subset \{1, \dots, N\} \times \mathbb{Z}^2} (\sigma_N^2)^{|\tau|} \mathbf{q}^{(0,x)}(\tau)^2 \\ &= 1 + \sum_{k=1}^N (\sigma_N^2)^k \sum_{\substack{0 =: n_0 < n_1 < \dots < n_k \leq N \\ x_0 =: x, x_1, \dots, x_k \in \mathbb{Z}^2}} \prod_{i=1}^k q_{n_i - n_{i-1}}(x_i - x_{i-1})^2 \\ &= 1 + \sum_{k=1}^N (\sigma_N^2)^k \sum_{0 =: n_0 < n_1 < \dots < n_k \leq N} q_{2(n_i - n_{i-1})}(0). \end{aligned}$$

If we let each increment $n_i - n_{i-1}$ vary freely in $\{1, 2, \dots, N\}$, by (1.17) we get the bound $\mathbb{E}[Z_{N,\beta_N}(x)^2] \leq \sum_{k \geq 0} (\sigma_N^2 R_N)^k = (1 - \sigma_N^2 R_N)^{-1}$. Recalling (1.18) and (2.15), we obtain

$$(3.2) \quad \forall \hat{\beta} \in (0, 1) \exists C_{\hat{\beta}} < \infty \text{ such that } \forall N \in \mathbb{N}: \quad \mathbb{E}[Z_{N,\beta_N}(x)^2] \leq C_{\hat{\beta}},$$

where $C_{\hat{\beta}}$ will denote a generic constant depending on $\hat{\beta}$.

Next, we look at $Z_{N,\beta_N}^A(x)$. The polynomial chaos expansion for $Z_{N,\beta_N}^A(x)$ is a subset of the one for $Z_{N,\beta_N}(x)$, hence the same bound (3.2) applies:

$$(3.3) \quad \forall \hat{\beta} \in (0, 1) \exists C_{\hat{\beta} < \infty} \text{ such that } \forall N \in \mathbb{N}: \quad \mathbb{E}[Z_{N,\beta_N}^A(x)^2] \leq C_{\hat{\beta}}.$$

We turn to $\hat{Z}_{N,\beta_N}^A(x)$. The bound (3.2) can again be applied, but it is quite poor. In fact, the following much better bound holds (recall that a_N is defined in (2.2)):

$$(3.4) \quad \forall \hat{\beta} \in (0, 1) \exists C_{\hat{\beta}} < \infty \text{ such that } \forall N \in \mathbb{N}: \quad \mathbb{E}[\hat{Z}_{N,\beta_N}^A(x)^2] \leq C_{\hat{\beta}} a_N.$$

The proof, given below, is elementary but slightly technical (see Section 3.4).

We conclude with an alternative viewpoint on the bound (3.2). If we denote by S and \tilde{S} two independent copies of the random walk, by (1.21)–(1.22) we can compute

$$(3.5) \quad \mathbb{E}[Z_{N,\beta_N}^2] = \mathbb{E}\mathbb{E}[e^{H_{N,\beta_N}(S)+H_{N,\beta_N}(\tilde{S})}] = \mathbb{E}[e^{(\lambda(2\beta_N)-2\lambda(\beta_N))\mathcal{L}_N(S,\tilde{S})}],$$

where $\mathcal{L}_N(S, \tilde{S})$ is the overlap of the two copies S, \tilde{S} up to time N , defined by

$$(3.6) \quad \mathcal{L}_N(S, \tilde{S}) := \sum_{n=1}^N \mathbb{1}_{\{S_n = \tilde{S}_n\}} = |S \cap \tilde{S} \cap (\{1, \dots, N\} \times \mathbb{Z}^2)|.$$

Since $\lambda(\beta) \sim \frac{1}{2}\beta^2$ as $\beta \downarrow 0$, see (1.19), we get

$$(3.7) \quad \mathbb{E}[Z_{N,\beta_N}^2] = \mathbb{E}[e^{(1+\varepsilon_N)\beta_N^2 \mathcal{L}_N(S,\tilde{S})}] \quad \text{where } \lim_{N \rightarrow \infty} \varepsilon_N = 0.$$

Note that $\frac{\pi}{\log N} \mathcal{L}_N(S, \tilde{S})$ converges in law to a mean 1 exponential random variable; see, for example, [19]. This matches with $\lim_{N \rightarrow \infty} \mathbb{E}[Z_{N,\beta_N}^2] = (1 - \hat{\beta}^2)^{-1}$ for β_N as in (1.18).

3.2. Positive moments via hypercontractivity. We will bound higher positive moments of our partition functions using the *hypercontractivity* of polynomial chaos [38], which we recall (with some strengthening) in Appendix B.

By (2.17), each partition function $Z_{N,\beta_N}(x)$, $Z_{N,\beta_N}^A(x)$, $\hat{Z}_{N,\beta_N}^A(x)$ can be expressed as a series

$$(3.8) \quad \sum_{k=0}^{\infty} X_k^{(N)}$$

(actually a finite sum) where $X_k^{(N)}$ is a multilinear polynomial of degree k in the i.i.d. random variables $(\xi_{n,x}^{(N)})_{(n,x) \in \mathbb{N} \times \mathbb{Z}^2}$, which have zero mean and unit variance; see (2.16). These random variables have uniformly bounded higher moments,

$$(3.9) \quad \forall p \in (2, \infty): \quad \sup_{N \in \mathbb{N}} \mathbb{E}[|\xi_{n,x}^{(N)}|^p] < \infty,$$

as one can check directly from (2.16) and (1.19) (see [6], equation (6.7)).

Under these conditions, hypercontractivity ensures that, for every $p \in (2, \infty)$, the p th moment of the series (3.8) can be bounded in terms of second moments,

$$(3.10) \quad \mathbb{E}\left[\left|\sum_{k=0}^{\infty} X_k^{(N)}\right|^p\right] \leq \left(\sum_{k=0}^{\infty} (c_p^k)^2 \mathbb{E}[(X_k^{(N)})^2]\right)^{p/2},$$

where $c_p \in (1, \infty)$ is a constant, uniform in N , which only depends on the laws of the $\xi_{n,x}^{(N)}$. This is proved in [38], Section 3.2 (extending [30]), where a nonoptimal value of c_p is given. We will recall these results in Appendix B, where we will prove that the optimal c_p satisfies

$$(3.11) \quad \lim_{p \downarrow 2} c_p = 1.$$

This result, which is of independent interest, is crucial in order to apply (3.10) to our partition functions $Z_{N,\beta_N}(x)$, $Z_{N,\beta_N}^A(x)$, $\hat{Z}_{N,\beta_N}^A(x)$, because for any subcritical $\hat{\beta} < 1$ we can fix $p > 2$ such that $c_p \hat{\beta} < 1$ is still subcritical. More precisely, note that multiplying $X_k^{(N)}$ by c_p^k amounts to replacing σ_N by $c_p \sigma_N$ (see (2.17)), and this corresponds asymptotically to replacing $\hat{\beta}$ by $c_p \hat{\beta}$; see (2.15) and (1.18). Then, by (3.2)–(3.4), we obtain

$$(3.12) \quad \forall \hat{\beta} \in (0, 1) \exists p = p_{\hat{\beta}} \in (2, \infty) \exists C'_{\hat{\beta}} < \infty \text{ such that } \forall N \in \mathbb{N} \\ \mathbb{E}[Z_{N,\beta_N}(x)^p] \leq C'_{\hat{\beta}}, \quad \mathbb{E}[Z_{N,\beta_N}^A(x)^p] \leq C'_{\hat{\beta}}, \quad \mathbb{E}[|\hat{Z}_{N,\beta_N}^A(x)|^p] \leq C'_{\hat{\beta}}(a_N)^{p/2}.$$

3.3. *Negative moments via concentration.* We give bounds on the *negative* moments of partition functions $Z_{N,\beta_N}(x)$ and $Z_{N,\beta_N}^A(x)$ (see (1.22), (2.17) and (2.3),(2.19)). We work with the general partition function $Z_{\Lambda,\beta}(x)$ defined in (2.1), which coincides with $Z_{N,\beta_N}(x)$, respectively, $Z_{N,\beta_N}^A(x)$, for $\Lambda = \{1, \dots, N\} \times \mathbb{Z}^2$, respectively, $\Lambda = A_N^x$.

For fixed (say bounded) $\Lambda \subseteq \mathbb{N} \times \mathbb{Z}^2$, it is not difficult to show that the log partition function $\log Z_{\Lambda,\beta}$ is a convex and Lipschitz function of the random variables $(\omega(n, y) : (n, y) \in \Lambda)$. However, if $\beta = \beta_N$ and the subset Λ grows with N , its Lipschitz constant can diverge as $N \rightarrow \infty$, hence we cannot directly apply the concentration inequality (1.20). However, it turns out that, for any $\Lambda = \Lambda_N \subseteq \{1, \dots, N\} \times \mathbb{Z}^2$, the Lipschitz constant is *tight* as $N \rightarrow \infty$. This yields the following estimate for the *left tail* of $\log Z_{\Lambda_N,\beta_N}$, proved below.

PROPOSITION 3.1 (Left tail). *For any $\hat{\beta} \in (0, 1)$, there exists $c_{\hat{\beta}} \in (0, \infty)$ with the following property: for every $N \in \mathbb{N}$ and for every choice of $\Lambda \subseteq \{1, \dots, N\} \times \mathbb{Z}^2$, one has*

$$(3.13) \quad \forall t \geq 0: \quad \mathbb{P}(\log Z_{\Lambda,\beta_N} \leq -t) \leq c_{\hat{\beta}} e^{-t^\gamma/c_{\hat{\beta}}},$$

where $\gamma > 1$ is the same exponent appearing in assumption (1.20).

As a corollary, for every $p \in (0, \infty)$ we can estimate, *uniformly in* $\Lambda \subseteq \{1, \dots, N\} \times \mathbb{Z}^2$,

$$\mathbb{E}[(Z_{\Lambda,\beta_N})^{-p}] = \mathbb{E}[e^{-p \log Z_{\Lambda,\beta_N}}] = p \int_{-\infty}^{\infty} e^{pt} \mathbb{E}[\mathbb{1}_{\{t < -\log Z_{\Lambda,\beta_N}\}}] dt \\ \leq 1 + p \int_0^{\infty} e^{pt} c_{\hat{\beta}} \exp(-t^\gamma/c_{\hat{\beta}}) dt =: C_{p,\hat{\beta}} < \infty.$$

Choosing $\Lambda = \{1, \dots, N\} \times \mathbb{Z}^2$ or $\Lambda = A_N^x$, we finally obtain the bounds

$$(3.14) \quad \forall \hat{\beta} \in (0, 1) \forall p \in (0, \infty) \exists C_{p,\hat{\beta}} < \infty: \quad \sup_{N \in \mathbb{N}} \mathbb{E}[Z_{N,\beta_N}(x)^{-p}] \leq C_{p,\hat{\beta}} < \infty,$$

$$(3.15) \quad \sup_{N \in \mathbb{N}} \mathbb{E}[Z_{N,\beta_N}^A(x)^{-p}] \leq C_{p,\hat{\beta}} < \infty.$$

For later use, let us also state the following consequence:

$$(3.16) \quad \forall \hat{\beta} \in (0, 1) \forall p \in (0, \infty) \exists C_{p,\hat{\beta}} < \infty: \quad \sup_{N \in \mathbb{N}} \mathbb{E}[|\log Z_{N,\beta_N}^A(x)|^p] \leq C_{p,\hat{\beta}} < \infty.$$

The proof of this fact is simple: we can bound $|\log y| \leq C_p(y^{1/p} + y^{-1/p})$ for all $y > 0$ and for suitable $C_p < \infty$ (just distinguish $y \geq 1$ and $y < 1$). This leads to $\mathbb{E}[|\log Z_{N,\beta_N}^A(x)|^p] \leq C_p(\mathbb{E}[Z_{N,\beta_N}^A(x)] + \mathbb{E}[Z_{N,\beta_N}^A(x)^{-1}]) = C_p(1 + \mathbb{E}[Z_{N,\beta_N}^A(x)^{-1}])$, so (3.16) follows by (3.15).

It remains to prove Proposition 3.1. To this goal, we follow the strategy developed in [9] for the pinning model, which generalizes [37]. We need the following result, which is [9], Proposition 3.4, inspired by [35], Proposition 1.6.

PROPOSITION 3.2. *Assume that disorder ω has the concentration property (1.20). There exist constants $c_1, c_2 \in (0, \infty)$ such that, for every $n \in \mathbb{N}$ and for every differentiable convex function $f : \mathbb{R}^n \rightarrow \mathbb{R}$, the following bound holds for all $a \in \mathbb{R}$ and $t, c \in (0, \infty)$:*

$$(3.17) \quad \mathbb{P}(f(\omega) \leq a - t)\mathbb{P}(f(\omega) \geq a, |\nabla f(\omega)| \leq c) \leq c_1 \exp\left(-\frac{(t/c)^\gamma}{c_2}\right),$$

where $\omega = (\omega_1, \dots, \omega_n)$ and $|\nabla f(\omega)| := \sqrt{\sum_{i=1}^n (\partial_i f(\omega))^2}$ is the norm of the gradient.

We can deduce the bound (3.13) from (3.17) applied to the function $f = f_N$ given by

$$(3.18) \quad f_N(\omega) = \log Z_{\Lambda, \beta_N}.$$

We only need to bound from below the second probability in the left-hand side of (3.17). This is provided by the next lemma, which completes the proof of Proposition 3.1.

LEMMA 3.3. *For any $\hat{\beta} \in (0, 1)$, there exist $c_{\hat{\beta}} \in (0, \infty)$ and $\vartheta_{\hat{\beta}} \in (0, 1)$ such that*

$$(3.19) \quad \inf_{N \in \mathbb{N}} \inf_{\Lambda \subseteq \{1, \dots, N\} \times \mathbb{Z}^2} \mathbb{P}(f_N(\omega) \geq -\log 2, |\nabla f_N(\omega)| \leq c_{\hat{\beta}}) \geq \vartheta_{\hat{\beta}} > 0.$$

PROOF. We set $a = -\log 2$. For any $c > 0$, we have

$$(3.20) \quad \mathbb{P}(f_N(\omega) \geq a, |\nabla f_N(\omega)| \leq c) = \mathbb{P}(f_N(\omega) \geq a) - \mathbb{P}(f_N(\omega) \geq a, |\nabla f_N(\omega)| > c).$$

The first probability can be estimated using the Paley–Zygmund inequality:

$$(3.21) \quad \mathbb{P}(f_N(\omega) \geq a) = \mathbb{P}\left(Z_{\Lambda, \beta_N} \geq \frac{1}{2}\right) = \mathbb{P}\left(Z_{\Lambda, \beta_N} \geq \frac{1}{2}\mathbb{E}[Z_{\Lambda, \beta_N}]\right) \geq \frac{\mathbb{E}[Z_{\Lambda, \beta_N}]^2}{4\mathbb{E}[(Z_{\Lambda, \beta_N})^2]}.$$

Note that $\mathbb{E}[Z_{\Lambda, \beta_N}] = 1$. For $\Lambda \subseteq \{1, \dots, N\} \times \mathbb{Z}^2$ we have $\mathbb{E}[(Z_{\Lambda, \beta_N})^2] \leq \mathbb{E}[(Z_{N, \beta_N})^2] \leq C_{\hat{\beta}}$ (see (3.2)), hence

$$(3.22) \quad \mathbb{P}(f_N(\omega) \geq a) \geq \frac{1}{4C_{\hat{\beta}}} =: 2\vartheta_{\hat{\beta}}.$$

We now proceed to estimate the second term in (3.20). First, we compute for $n \in \mathbb{N}, x \in \mathbb{Z}^2$,

$$\frac{\partial f_N(\omega)}{\partial \omega_{n,x}} = \frac{1}{Z_{\Lambda, \beta_N}} \mathbb{E}[\beta_N \mathbb{1}_{(n,x) \in S \cap \Lambda} e^{H_{\Lambda, \beta_N}(S)}] \quad \text{and}$$

$$|\nabla f_N(\omega)|^2 = \sum_{(n,x) \in \mathbb{N} \times \mathbb{Z}^2} \left(\frac{\partial f_N}{\partial \omega_{n,x}}\right)^2 = \frac{1}{(Z_{\Lambda, \beta_N})^2} \mathbb{E}[\beta_N^2 |S \cap \tilde{S} \cap \Lambda| e^{H_{\Lambda, \beta_N}(S) + H_{\Lambda, \beta_N}(\tilde{S})}],$$

where S and \tilde{S} are two independent copies of the random walk, and with some abuse of notation, we also denote by S the random subset $\{(n, S_n)\}_{n \in \mathbb{N}} \subseteq \mathbb{N} \times \mathbb{Z}^2$.

For $\Lambda \subseteq \{1, \dots, N\} \times \mathbb{Z}^2$, we have $|S \cap \tilde{S} \cap \Lambda| \leq \mathcal{L}_N(S, \tilde{S})$ (see (3.6)), where $\mathcal{L}_N(S, \tilde{S})$ denotes the overlap up to time N of the two trajectories S and \tilde{S} . On the event that $f_N(\omega) \geq a = -\log 2$, that is, $Z_{\Lambda, \beta_N} \geq 1/2$, we can thus bound

$$|\nabla f_N(\omega)|^2 \leq 4\mathbb{E}[\beta_N^2 \mathcal{L}_N(S, \tilde{S}) e^{H_{\Lambda, \beta_N}(S) + H_{\Lambda, \beta_N}(\tilde{S})}],$$

and note that, arguing as in (3.5)–(3.7), for every $\delta > 0$ we have, for all N large enough,

$$\begin{aligned} \mathbb{E}\mathbb{E}[\beta_N^2 \mathcal{L}_N(S, \tilde{S}) e^{H_{\Lambda, \beta_N}(S) + H_{\Lambda, \beta_N}(\tilde{S})}] &\leq \mathbb{E}[\beta_N^2 \mathcal{L}_N(S, \tilde{S}) e^{(1+\delta)\beta_N^2 \mathcal{L}_N(S, \tilde{S})}] \\ &\leq \frac{1}{\delta} \mathbb{E}[e^{(1+2\delta)\beta_N^2 \mathcal{L}_N(S, \tilde{S})}], \end{aligned}$$

where we used the bound $x \leq \frac{1}{\delta} e^{\delta x}$. Thus, for all N large enough we have

$$\mathbb{P}(f_N(\omega) \geq a, |\nabla f_N(\omega)| > c) \leq \frac{1}{c^2} \mathbb{E}[|\nabla f_N(\omega)|^2 \mathbb{1}_{\{f_N(\omega) \geq a\}}] \leq \frac{4}{c^2} \frac{1}{\delta} \mathbb{E}[e^{(1+2\delta)\beta_N^2 \mathcal{L}_N(S, \tilde{S})}].$$

Let us now define $\hat{\beta}' := \frac{1+\hat{\beta}}{2}$, so that $\hat{\beta} < \hat{\beta}' < 1$, and define $\beta'_N := \hat{\beta}'/\sqrt{R_N}$; see (1.18). Then we can fix $\delta = \delta_{\hat{\beta}} > 0$ small enough so that $(1 + 2\delta)\beta_N^2 < \lambda(2\beta'_N) - 2\lambda(\beta'_N)$ (note that $\lambda(2\beta) - 2\lambda(\beta) \sim \beta^2$ as $\beta \rightarrow 0$), hence by (3.5) and (3.2),

$$\mathbb{P}(f_N(\omega) \geq a, |\nabla f_N(\omega)| > c) \leq \frac{4}{c^2} \frac{1}{\delta_{\hat{\beta}}} C_{\hat{\beta}'}$$

Choosing $c = c_{\hat{\beta}}$ large enough, we can make the right-hand side smaller than $\vartheta_{\hat{\beta}}$; see (3.22). Looking back at (3.20), we see that (3.19) is proved. \square

3.4. *Proof of equation (3.4).* The quantity $\mathbb{E}[\hat{Z}_{N, \beta_N}^A(x)^2]$ admits a representation similar to the first line of (3.1), without the constant term 1 and with the inner sum restricted to space-time points such that $(n_i, x_i) \notin A_N^x$ for some $i = 1, \dots, k$, that is, either $n_i > N^{1-a_N}$ or $|x_i - x| \geq N^{\frac{1}{2} - \frac{a_N}{4}}$. Since there are k space-time points, for some $j = 1, \dots, k$ we must have either $n_j - n_{j-1} > \frac{1}{k} N^{1-a_N}$ or $|x_j - x_{j-1}| \geq \frac{1}{k} N^{\frac{1}{2} - \frac{a_N}{4}}$ (we recall that $n_0 = 0$ and $x_0 = x$). Defining the new variables $\ell_i := n_i - n_{i-1}$ and $z_i := x_i - x_{i-1}$, and enlarging the range $0 < n_1 < \dots < n_k \leq N$ to $\ell_1, \dots, \ell_k \in \{1, \dots, N\}$, we can then bound

$$\begin{aligned} \mathbb{E}[\hat{Z}_{N, \beta_N}^A(x)^2] &\leq \sum_{k=1}^N (\sigma_N^2)^k \sum_{\substack{\ell_1, \dots, \ell_k \in \{1, \dots, N\} \\ z_1, \dots, z_k \in \mathbb{Z}^2}} \sum_{j=1}^k (\mathbb{1}_{\{\ell_j > \frac{1}{k} N^{1-a_N}\}} \\ &\quad + \mathbb{1}_{\{\ell_j \leq \frac{1}{k} N^{1-a_N}, |z_j| \geq \frac{1}{k} N^{\frac{1}{2} - \frac{a_N}{4}}\}}) \prod_{i=1}^k q_{\ell_i}(z_i)^2. \end{aligned} \tag{3.23}$$

We now switch the sum over j with the double sum over ℓ_i, z_i 's. We can sum over all variables z_j 's with $i \neq j$, replacing each kernel $q_{\ell_i}(z_i)^2$ by $q_{2\ell_i}(0)$ (see (1.17)), and then sum $q_{2\ell_i}(0)$ for all ℓ_i 's with $i \neq j$, which gives $(R_N)^{k-1}$ (see again (1.17)). This yields

$$\mathbb{E}[\hat{Z}_{N, \beta_N}^A(x)^2] \leq \sum_{k=1}^N (\sigma_N^2)^k R_N^{k-1} k \sum_{\substack{\ell \in \{1, \dots, N\} \\ z \in \mathbb{Z}^2}} (\mathbb{1}_{\{\ell > \frac{1}{k} N^{1-a_N}\}} + \mathbb{1}_{\{\ell \leq \frac{1}{k} N^{1-a_N}, |z| \geq \frac{1}{k} N^{\frac{1}{2} - \frac{a_N}{4}}\}}) q_{\ell}(z)^2.$$

We now consider separately the contributions of the two indicator functions.

- Recalling (1.17), (1.18), (2.15), the contribution of $\{\ell > \frac{1}{k}N^{1-a_N}\}$ is controlled by

$$\begin{aligned} \sum_{k=1}^N (\sigma_N^2)^k R_N^{k-1} k \sum_{\frac{1}{k}N^{1-a_N} < \ell \leq N} q_{2\ell}(0) &\leq C \sum_{k=1}^N k(\hat{\beta}^2)^k \frac{1}{\log N} \sum_{\frac{1}{k}N^{1-a_N} < \ell \leq N} \frac{1}{\ell} \\ &\leq C' \sum_{k=1}^N k(\hat{\beta}^2)^k \frac{a_N \log N + \log k}{\log N} \\ &\leq C' \left(\tilde{C}_{\hat{\beta}} a_N + \hat{C}_{\hat{\beta}} \frac{1}{\log N} \right), \end{aligned}$$

where $\tilde{C}_{\hat{\beta}} := \sum_{k=1}^{\infty} k(\hat{\beta}^2)^k$ and $\hat{C}_{\hat{\beta}} := \sum_{k=1}^{\infty} k(\log k)(\hat{\beta}^2)^k$ are finite, $\hat{\beta}$ -dependent constants. This contribution is consistent with (3.4) (recall (2.2)).

- The contribution of $\{\ell \leq \frac{1}{k}N^{1-a_N}, |z| \geq \frac{1}{k}N^{\frac{1}{2}-\frac{a_N}{4}}\}$ is given by

$$(3.24) \quad \sum_{k=1}^N (\sigma_N^2)^k R_N^{k-1} k \sum_{1 \leq \ell \leq \frac{1}{k}N^{1-a_N}} \sum_{|z| > \frac{1}{k}N^{\frac{1}{2}-\frac{a_N}{4}}} q_{\ell}(z)^2.$$

Note that we can enlarge the range of the last sum to $|z| > \vartheta \sqrt{\ell}$, with $\vartheta = N^{\frac{a_N}{4}}/\sqrt{k}$. Note that $\sup_{z \in \mathbb{Z}^2} q_{\ell}(z) \leq c/\ell$, by Gnedenko’s local limit theorem. Then, by Gaussian estimates for the simple random walk on \mathbb{Z}^2 , there is $\eta > 0$ such that

$$\sum_{|z| > \vartheta \sqrt{\ell}} q_{\ell}(z)^2 \leq \frac{c}{\ell} \mathbb{P}(|S_{\ell}| > \vartheta \sqrt{\ell}) \leq \frac{c}{\ell} e^{-\eta \vartheta^2} \quad \forall \ell \in \mathbb{N}, \forall \vartheta > 0.$$

Then we can bound (3.24) by a constant multiple of

$$(3.25) \quad \sum_{k=1}^N (\hat{\beta}^2)^k k \frac{1}{R_N} \sum_{1 \leq \ell \leq N} \frac{c}{\ell} e^{-\eta \vartheta^2} \leq C \sum_{k=1}^N (\hat{\beta}^2)^k k e^{-\eta \frac{N^{\frac{a_N}{2}}}{k}}.$$

We split the sum according to $k \leq (N^{\frac{a_N}{2}})^{1/2}$ and $k > (N^{\frac{a_N}{2}})^{1/2}$, getting the bound

$$\left\{ \sum_{k=1}^{\infty} (\hat{\beta}^2)^k k \right\} e^{-\eta (N^{\frac{a_N}{2}})^{1/2}} + \left\{ \sum_{k=1}^{\infty} \hat{\beta}^k k \right\} \hat{\beta}^{(N^{\frac{a_N}{2}})^{1/2}}.$$

Both brackets are finite, $\hat{\beta}$ -dependent constants, while the other factors are both $o(a_N)$ as $N \rightarrow \infty$, by (2.2), because $\hat{\beta} < 1$ and $N^{\frac{a_N}{2}} = \exp(\frac{1}{2}(\log N)^{\gamma}) \gg \log N$.

This completes the proof of (3.4).

4. Edwards–Wilkinson fluctuations for directed polymer. In this section, we prove Theorem 1.6, which consists of proving Propositions 2.1, 2.2, 2.3 and 2.4 as described in Section 2. The proofs are given in the following subsections.

4.1. *Proof of Proposition 2.1.* Recalling (2.8), we need to show that

$$\frac{\log N}{N^2} \sum_{x, y \in \mathbb{Z}^2} \text{Cov}[O_N(x), O_N(y)] \phi\left(\frac{x}{\sqrt{N}}\right) \phi\left(\frac{y}{\sqrt{N}}\right) \xrightarrow{N \rightarrow \infty} 0.$$

By translation invariance and Cauchy–Schwarz, it suffices to show that for any $x \in \mathbb{Z}^2$,

$$(4.1) \quad (\log N) \mathbb{E}[O_N(x)^2] \xrightarrow{N \rightarrow \infty} 0.$$

We recall that $O_N(x)$ is defined in (2.7), and in view of (2.5) we can write

$$O_N(x) = \log\left(1 + \frac{\hat{Z}_{N,\beta_N}^A(x)}{Z_{N,\beta_N}^A(x)}\right) - \frac{\hat{Z}_{N,\beta_N}^A(x)}{Z_{N,\beta_N}^A(x)}.$$

We can bound, for a suitable constant $C < \infty$,

$$(4.2) \quad |\log(1+y) - y| \leq C \cdot \begin{cases} \sqrt{\frac{|y|}{1+y}} & \text{if } -1 < y < 0, \\ y^2 & \text{if } -\frac{1}{2} \leq y \leq \frac{1}{2}, \\ |y| & \text{if } 0 < y < \infty. \end{cases}$$

The three domains are chosen to overlap on purpose: in fact, we will apply these inequalities in the domains $(-1, -a_N^{2/7})$, $[-a_N^{2/7}, a_N^{2/7}]$ and $(a_N^{2/7}, \infty)$ (recall a_N from (2.2)). We define

$$D_N^\pm := \left\{ \pm \frac{\hat{Z}_{N,\beta_N}^A(x)}{Z_{N,\beta_N}^A(x)} > a_N^{2/7} \right\}, \quad D_N := D_N^+ \cup D_N^- = \left\{ \left| \frac{\hat{Z}_{N,\beta_N}^A(x)}{Z_{N,\beta_N}^A(x)} \right| > a_N^{2/7} \right\},$$

and we bound

$$(4.3) \quad \begin{aligned} \mathbb{P}(D_N) &\leq \mathbb{P}(Z_{N,\beta_N}^A(x) < a_N^{1/7}) + \mathbb{P}(|\hat{Z}_{N,\beta_N}^A(x)| > a_N^{3/7}) \\ &\leq a_N^{1/7} \mathbb{E}[Z_{N,\beta_N}^A(x)^{-1}] + a_N^{-6/7} \mathbb{E}[\hat{Z}_{N,\beta_N}^A(x)^2] \leq (C_{2,\hat{\beta}} + C_{\hat{\beta}}) a_N^{1/7}, \end{aligned}$$

thanks to (3.15) and (3.4). Then by (4.2),

$$\begin{aligned} \frac{1}{C} \mathbb{E}[O_N(x)^2] &\leq \mathbb{E}\left[\left(\frac{\hat{Z}_{N,\beta_N}^A(x)}{Z_{N,\beta_N}^A(x)}\right)^4 \mathbb{1}_{D_N^c}\right] + \mathbb{E}\left[\left(\frac{\hat{Z}_{N,\beta_N}^A(x)}{Z_{N,\beta_N}^A(x)}\right)^2 \mathbb{1}_{D_N^+}\right] \\ &\quad + \mathbb{E}\left[\frac{|\hat{Z}_{N,\beta_N}^A(x)/Z_{N,\beta_N}^A(x)|}{1 + \hat{Z}_{N,\beta_N}^A(x)/Z_{N,\beta_N}^A(x)} \mathbb{1}_{D_N^-}\right], \end{aligned}$$

and given that

$$1 + \frac{\hat{Z}_{N,\beta_N}^A(x)}{Z_{N,\beta_N}^A(x)} = \frac{Z_{N,\beta_N}(x)}{Z_{N,\beta_N}^A(x)},$$

we can choose $p = p_{\hat{\beta}} > 2$ close to 2 as in (3.12) such that

$$\begin{aligned} \frac{1}{C} \mathbb{E}[O_N(x)^2] &\leq \mathbb{E}\left[\left(\frac{\hat{Z}_{N,\beta_N}^A(x)}{Z_{N,\beta_N}^A(x)}\right)^4 \mathbb{1}_{D_N^c}\right] + \mathbb{E}\left[\left(\frac{\hat{Z}_{N,\beta_N}^A(x)}{Z_{N,\beta_N}^A(x)}\right)^2 \mathbb{1}_{D_N^+}\right] + \mathbb{E}\left[\left|\frac{\hat{Z}_{N,\beta_N}^A(x)}{Z_{N,\beta_N}^A(x)}\right| \mathbb{1}_{D_N^-}\right] \\ &\leq a_N^{\frac{8}{7}} + \mathbb{E}[\hat{Z}_{N,\beta_N}^A(x)^p]^{\frac{2}{p}} (\mathbb{E}[Z_{N,\beta_N}^A(x)^{-\frac{2p}{p-2}} \mathbb{1}_{D_N^+}])^{1-\frac{2}{p}} + \mathbb{E}[Z_{N,\beta_N}(x)^{-2} \mathbb{1}_{D_N^-}]^{\frac{1}{2}} \\ &\leq a_N^{\frac{8}{7}} + C'_\beta a_N (\mathbb{E}[Z_{N,\beta_N}^A(x)^{-\frac{4p}{p-2}}])^{\frac{1}{2}-\frac{1}{p}} \mathbb{P}(D_N)^{\frac{1}{2}-\frac{1}{p}} \\ &\quad + \mathbb{E}[Z_{N,\beta_N}(x)^{-4}]^{\frac{1}{4}} \mathbb{P}(D_N)^{\frac{1}{4}} \\ &\leq a_N^{\frac{8}{7}} + C'_\beta a_N \mathbb{P}(D_N)^{\frac{1}{4} \wedge (\frac{1}{2}-\frac{1}{p})} \leq C'_\beta a_N^{1+\frac{1}{7}(\frac{1}{2}-\frac{1}{p})}, \end{aligned}$$

where the second last inequality holds by (3.12), (3.14) and (3.15); in the last inequality, we applied (4.3), and $C'_\beta < \infty$ is a generic constant depending only on $\hat{\beta}$. Recall from (2.2) that

$a_N = (\log N)^{\gamma-1}$. We can then choose $\gamma \in (0, \gamma^*)$ with $\gamma^* > 0$ small enough such that

$$(4.4) \quad \mathbb{E}[O_N(x)^2] \leq C'_{\hat{\beta}} a_N^{1+\frac{1}{7}(\frac{1}{2}-\frac{1}{p})} = C'_{\hat{\beta}} (\log N)^{-(1-\gamma)(1+\frac{1}{7}(\frac{1}{2}-\frac{1}{p}))} = o((\log N)^{-1}).$$

Therefore, (4.1) holds.

4.2. *Proof of Proposition 2.2.* We need to show that

$$(4.5) \quad \frac{\log N}{N^2} \sum_{x,y \in \mathbb{Z}^2} \text{Cov}[\log Z_{N,\beta_N}^A(x), \log Z_{N,\beta_N}^A(y)] \phi\left(\frac{x}{\sqrt{N}}\right) \phi\left(\frac{y}{\sqrt{N}}\right) \xrightarrow{N \rightarrow \infty} 0.$$

We recall that $Z_{N,\beta_N}^A(x)$ depends only on the disorder within set A_N^x , defined in (2.3), hence

$Z_{N,\beta_N}^A(x)$ and $Z_{N,\beta_N}^A(y)$ are independent for $|x - y| > 2N^{\frac{1}{2}-\frac{a_N}{4}}$. By Cauchy–Schwarz and (3.16), we can bound the left-hand side of (4.5) as follows:

$$\begin{aligned} & C_{2,\hat{\beta}} \frac{\log N}{N^2} \sum_{x,y \in \mathbb{Z}^2: |y-x| \leq 2N^{\frac{1}{2}-\frac{a_N}{4}}} \phi\left(\frac{x}{\sqrt{N}}\right) \phi\left(\frac{y}{\sqrt{N}}\right) \\ & \leq c C_{2,\hat{\beta}} \frac{\log N}{N^2} N^{1-\frac{a_N}{2}} |\phi|_{\infty} \sum_{x \in \mathbb{Z}^2} \left| \phi\left(\frac{x}{\sqrt{N}}\right) \right| \\ & \leq c' C_{2,\hat{\beta}} (\log N) N^{-\frac{a_N}{2}} |\phi|_{\infty} |\phi|_{L^1(\mathbb{R}^2)} = c' C_{2,\hat{\beta}} e^{\log(\log N) - \frac{1}{2}(\log N)^{\gamma}} |\phi|_{\infty} |\phi|_{L^1(\mathbb{R}^2)} \xrightarrow{N \rightarrow \infty} 0, \end{aligned}$$

where c, c' are generic constants, and the last equality holds by definition of a_N in (2.2).

4.3. *Proof of Proposition 2.3.* We need to show that

$$(4.6) \quad \begin{aligned} & \frac{\sqrt{\log N}}{N} \sum_{x \in \mathbb{Z}^2} \frac{\hat{Z}_{N,\beta_N}^A(x)}{Z_{N,\beta_N}^A(x)} \phi\left(\frac{x}{\sqrt{N}}\right) \\ & - \frac{\sqrt{\log N}}{N} \sum_{x \in \mathbb{Z}^2} (Z_{N,\beta_N}^{B \geq}(x) - 1) \phi\left(\frac{x}{\sqrt{N}}\right) \xrightarrow[N \rightarrow \infty]{L^1(\mathbb{P})} 0. \end{aligned}$$

We recall that $B_N^{\geq} := ((N^{1-9a_N/40}, N] \cap \mathbb{N}) \times \mathbb{Z}^2$; see (2.10). We define new subsets

$$(4.7) \quad B_N := ((N^{1-a_N}, N] \cap \mathbb{N}) \times \mathbb{Z}^2,$$

$$(4.8) \quad C_N^x := \{(n, z) \in \mathbb{N} \times \mathbb{Z}^2 : n \leq N^{1-a_N}, |z - x| \geq N^{\frac{1}{2}-\frac{a_N}{4}}\},$$

and we introduce new “partition functions”:

$$(4.9) \quad Z_{N,\beta_N}^{A,C}(x) := \sum_{\tau \subset \{1, \dots, N\} \times \mathbb{Z}^2: \tau \cap C_N^x \neq \emptyset} \sigma_N^{|\tau|} \mathbf{q}^{(0,x)}(\tau) \xi(\tau),$$

$$(4.10) \quad Z_{N,\beta_N}^{A,B}(x) := \sum_{\tau \subset A_N^x \cup B_N: \tau \cap B_N \neq \emptyset} \sigma_N^{|\tau|} \mathbf{q}^{(0,x)}(\tau) \xi(\tau),$$

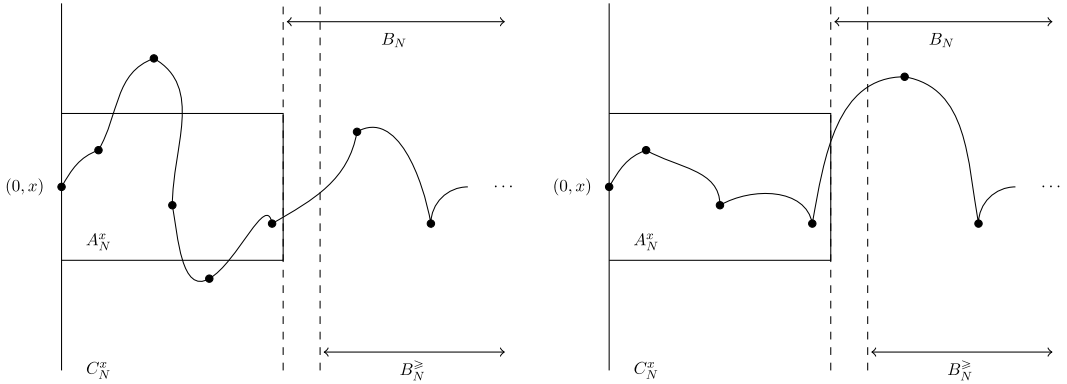
similar to the polynomial chaos expansions for Z_{N,β_N} , Z_{N,β_N}^A and $Z_{N,\beta_N}^{B \geq}$ in (2.18)–(2.19). See Figure 2 for a graphical representation of the chaos expansions.

Recall that A_N^x was defined in (2.3), and note that $(\{1, \dots, N\} \times \mathbb{Z}^2) \setminus A_N^x = C_N^x \cup B_N$.

We can then decompose $\hat{Z}_{N,\beta_N}^A(x)$, defined in (2.5), as follows:

$$(4.11) \quad \hat{Z}_{N,\beta_N}^A(x) = Z_{N,\beta_N}^{A,B}(x) + Z_{N,\beta_N}^{A,C}(x).$$

We split the sequel in three steps. The first step is:



(A) Partition function $Z_{N, \beta_N}^{A,C}(x)$. (B) Partition function $Z_{N, \beta_N}^{A,B}$.

FIG. 2. The figures depict the chaos expansions of $Z_{N, \beta_N}^{A,C}(x)$, $Z_{N, \beta_N}^{A,B}(x)$. Each term in the expansion for $Z_{N, \beta_N}^{A,C}(x)$ must include disorder from C_N^x ; while each term in the expansion for $Z_{N, \beta_N}^{A,B}(x)$ contain only disorder from $A_N^x \cup B_N$, with at least some disorder from B_N .

(1) We will show that the contribution of the term $Z_{N, \beta_N}^{A,C}$ is negligible for (4.6).

To treat the term $Z_{N, \beta_N}^{A,B}(x)$, we decompose its chaos expansion (4.10) according to the last point (t, w) of τ that lies in A_N^x and the first point (r, z) of τ that lies in B_N :

$$(4.12) \quad Z_{N, \beta_N}^{A,B}(x) = \sum_{\substack{(t,w) \in \{(0,x)\} \cup A_N^x \\ (r,z) \in B_N}} Z_{0,t, \beta_N}^A(x, w) \cdot q_{r-t}(z-w) \cdot \sigma_N \xi_{r,z} \cdot Z_{r,N, \beta_N}(z),$$

where $Z_{0,t, \beta_N}^A(x, w)$ is the ‘‘point-to-point’’ partition function from $(0, x)$ to (t, w) , defined by $Z_{0,t, \beta_N}^A(x, w) := 1$ if $(t, w) = (0, x)$ and by

$$(4.13) \quad Z_{0,t, \beta_N}^A(x, w) := \sum_{\tau \subset A_N^x \cap ([0,t] \times \mathbb{Z}^2): \tau \ni (t,w)} \sigma_N^{|\tau|} \mathbf{q}^{(0,x)}(\tau) \boldsymbol{\xi}(\tau) \quad \text{if } t > 0,$$

while $Z_{r,N, \beta_N}(z)$ is the ‘‘point-to-plane’’ partition function starting at (r, z) and running until time N :

$$(4.14) \quad Z_{r,N, \beta_N}(z) := \sum_{\tau \subset \{r+1, \dots, N\} \times \mathbb{Z}^2} \sigma_N^{|\tau|} \mathbf{q}^{(r,z)}(\tau) \boldsymbol{\xi}(\tau).$$

The next steps are:

(2) We will show that in (4.12) the contribution from $r < N^{1-9a_N/40}$ is negligible for (4.6).

(3) We will show that in (4.12) we can replace the kernel $q_{r-t}(z-w)$ by $q_r(z-x)$, that is, the transition kernel from $(0, x)$ to (r, z) , because their difference is negligible for (4.6).

Finally, note that when we restrict the sum in (4.12) to $r \geq N^{1-9a_N/40}$, that is, to $(r, z) \in B_N^{\geq}$ (recall (2.10)), and we replace $q_{r-t}(z-w)$ by $q_r(z-x)$, the right-hand side of (4.12) becomes exactly $Z_{N, \beta_N}^A(x)(Z_{N, \beta_N}^{B \geq}(x) - 1)$ (recall (2.19)). This completes the proof of (4.6).

It remains to prove the three steps stated above.

Step (1). We show that the contribution of $Z_{N, \beta_N}^{A,C}$ in (4.11) to (4.6) is negligible, that is,

$$(4.15) \quad \frac{\sqrt{\log N}}{N} \sum_{x \in \mathbb{Z}^2} \frac{Z_{N, \beta_N}^{A,C}(x)}{Z_{N, \beta_N}^A(x)} \phi\left(\frac{x}{\sqrt{N}}\right) \xrightarrow[N \rightarrow \infty]{L^2(\mathbb{P})} 0.$$

Since the chaos expansion of $Z_{N,\beta_N}^{A,C}(x)$ in (4.9) contains disorder ξ outside A_N^x , not contained in the expansion of $Z_{N,\beta_N}^A(x)$, we have that $\mathbb{E}[Z_{N,\beta_N}^{A,C}(x)/Z_{N,\beta_N}^A(x)] = 0$ thus $L^2(\mathbb{P})$ and variance computations are equivalent. We then have

$$\begin{aligned} & \mathbb{V}ar\left(\frac{\sqrt{\log N}}{N} \sum_{x \in \mathbb{Z}^2} \frac{Z_{N,\beta_N}^{A,C}(x)}{Z_{N,\beta_N}^A(x)} \phi\left(\frac{x}{\sqrt{N}}\right)\right) \\ &= \frac{\log N}{N^2} \sum_{x,y \in \mathbb{Z}^2} \mathbb{E}\left[\frac{Z_{N,\beta_N}^{A,C}(x)}{Z_{N,\beta_N}^A(x)} \cdot \frac{Z_{N,\beta_N}^{A,C}(y)}{Z_{N,\beta_N}^A(y)}\right] \phi\left(\frac{x}{\sqrt{N}}\right) \phi\left(\frac{y}{\sqrt{N}}\right). \end{aligned}$$

By Cauchy–Schwarz, we can further bound this as follows, for some constant c :

$$\begin{aligned} & \frac{\log N}{N^2} \mathbb{E}\left[\left(\frac{Z_{N,\beta_N}^{A,C}(0)}{Z_{N,\beta_N}^A(0)}\right)^2\right] \sum_{x,y \in \mathbb{Z}^2} \left|\phi\left(\frac{x}{\sqrt{N}}\right)\right| \left|\phi\left(\frac{y}{\sqrt{N}}\right)\right| \\ (4.16) \quad & \leq c \log N \cdot |\phi|_{L^1(\mathbb{R}^2)}^2 \mathbb{E}\left[\left(\frac{Z_{N,\beta_N}^{A,C}(0)}{Z_{N,\beta_N}^A(0)}\right)^2\right] \\ & \leq c \log N \cdot |\phi|_{L^1(\mathbb{R}^2)}^2 \mathbb{E}[Z_{N,\beta_N}^{A,C}(0)^{2p}]^{1/p} \mathbb{E}\left[\frac{1}{Z_{N,\beta_N}^A(0)^{2q}}\right]^{1/q}, \end{aligned}$$

where in the last step we used Hölder inequality with parameters (p, q) with $p^{-1} + q^{-1} = 1$, and p will be chosen sufficiently close to 1, to be determined below. The term $\mathbb{E}[Z_{N,\beta_N}^A(0)^{-2q}]^{1/q}$ can be uniformly bounded by the negative moment estimate (3.15).

We can use hypercontractivity (see (3.10)) to bound

$$\mathbb{E}[Z_{N,\beta_N}^{A,C}(0)^{2p}]^{1/p} \leq \sum_{\tau \subset \{1, \dots, N\} \times \mathbb{Z}^2: \tau \cap C_N^0 \neq \emptyset} (c_{2p} \sigma_N)^{2|\tau|} \mathbf{q}^{(0)}(\tau)^2.$$

The right-hand side is the second moment of the partition function (see (3.1)) except that σ_N is replaced by $c_{2p} \sigma_N$ (which corresponds asymptotically to replacing $\hat{\beta}$ by $\hat{\beta}' := c_{2p} \hat{\beta}$, see (2.15) and (1.18)) and the random walk S must satisfy $\max\{|S_n| : n < N^{1-a_N}\} > N^{1/2-a_N/4}$. In particular, recalling (3.7) and (3.6), this can be bounded by

$$(4.17) \quad \mathbb{E}\left[e^{(1+o(1))(c_{2p}\beta_N)^2 \mathcal{L}_{N^{1-a_N}}(S^{(1)}, S^{(2)})}; \max_{n \leq N^{1-a_N}} |S_n^{(i)}| > N^{\frac{1}{2} - \frac{a_N}{4}}, \text{ for } i = 1, 2\right],$$

where $S^{(1)}, S^{(2)}$ are two independent random walk copies. This is bounded via Hölder by

$$\mathbb{E}\left[e^{(1+o(1))p(c_{2p}\beta_N)^2 \mathcal{L}_{N^{1-a_N}}(S^{(1)}, S^{(2)})}\right]^{1/p} \mathbb{P}\left(\max_{n \leq N^{1-a_N}} |S_n| > N^{\frac{1}{2} - \frac{a_N}{4}}\right)^{2/q}.$$

We can now choose $p > 1$ sufficiently close to 1 so that $\sqrt{p}c_{2p}\hat{\beta} < 1$, that is, still subcritical, which is possible because $\lim_{p \rightarrow 1} c_{2p} = 1$; see (3.11). Hence the expectation above is uniformly bounded in N as shown in Section 3.1. On the other hand, standard moderate deviation estimates for the simple symmetric random walk show that

$$\mathbb{P}\left(\max_{n \leq N^{1-a_N}} |S_n| > N^{\frac{1}{2} - \frac{a_N}{4}}\right) \leq \exp(-cN^{a_N/2}) = \exp(-ce^{(\log N)^\gamma/2}),$$

where we recall that $a_N = (\log N)^{\gamma-1}$; see (2.2). Inserting these estimates in (4.16), we get

$$\mathbb{V}ar\left(\frac{\sqrt{\log N}}{N} \sum_{x \in \mathbb{Z}^2} \frac{Z_{N,\beta_N}^{A,C}(x)}{Z_{N,\beta_N}^A(x)} \phi\left(\frac{x}{\sqrt{N}}\right)\right) \leq c \log N \cdot |\phi|_{L^1(\mathbb{R}^2)}^2 \exp(-ce^{(\log N)^\gamma/2}) \xrightarrow{N \rightarrow \infty} 0.$$

Step (2). We show that in the chaos expansion (4.12) for $Z_{N,\beta_N}^{A,B}(x)$, the contribution from (r, z) with $r < N^{1-9a_N/40}$ is negligible for (4.6). The contribution we are after is

$$(4.18) \quad Z_{N,\beta_N}^{A,B^<}(x) := \sum_{\substack{(t,w) \in A_N^x \\ (r,z) \in B_N: r < N^{1-9a_N/40}}} Z_{0,t,\beta_N}^A(x, w) \cdot q_{r-t}(z - w) \cdot \sigma_N \xi_{r,z} \cdot Z_{r,N,\beta_N}(z),$$

and we want to show that

$$(4.19) \quad \mathbb{E} \left[\left(\frac{\sqrt{\log N}}{N} \sum_{x \in \mathbb{Z}^2} \phi \left(\frac{x}{\sqrt{N}} \right) \frac{Z_{N,\beta_N}^{A,B^<}}{Z_{N,\beta_N}^A}(x) \right)^2 \right] \xrightarrow{N \rightarrow \infty} 0.$$

The left-hand side of (4.19) equals

$$(4.20) \quad \frac{\log N}{N^2} \sum_{x,y \in \mathbb{Z}^2} \phi \left(\frac{x}{\sqrt{N}} \right) \phi \left(\frac{y}{\sqrt{N}} \right) \mathbb{E} \left[\frac{Z_{N,\beta_N}^{A,B^<}}{Z_{N,\beta_N}^A}(x) \cdot \frac{Z_{N,\beta_N}^{A,B^<}}{Z_{N,\beta_N}^A}(y) \right].$$

We can restrict the summation over x, y to $|x - y| > N^{\frac{1}{2} - \frac{a_N}{10}}$. Indeed, in the complementary regime, we first bound the expectation in (4.20) by Cauchy–Schwarz and obtain the bound

$$\begin{aligned} & \frac{\log N}{N^2} \mathbb{E} \left[\left(\frac{Z_{N,\beta_N}^{A,B^<}}{Z_{N,\beta_N}^A}(0) \right)^2 \right] \sum_{|x-y| \leq N^{\frac{1}{2} - \frac{a_N}{10}}} \left| \phi \left(\frac{x}{\sqrt{N}} \right) \right| \left| \phi \left(\frac{y}{\sqrt{N}} \right) \right| \\ & \leq N^{1 - \frac{a_N}{5}} \frac{\log N}{N} \mathbb{E} \left[\left(\frac{Z_{N,\beta_N}^{A,B^<}}{Z_{N,\beta_N}^A}(0) \right)^2 \right] |\phi|_\infty |\phi|_{L^1(\mathbb{R}^2)} \\ & = (\log N) e^{-\frac{1}{5}(\log N)^\gamma} \mathbb{E} \left[\left(\frac{Z_{N,\beta_N}^{A,B^<}}{Z_{N,\beta_N}^A}(0) \right)^2 \right] |\phi|_\infty |\phi|_{L^1(\mathbb{R}^2)}, \end{aligned}$$

which goes to zero as $N \rightarrow \infty$, since expectation can be bounded via Hölder with an exponent p for $Z_{N,\beta_N}^{A,B^<}(0)^2$ chosen sufficiently close to one, so that the hypercontractivity bound (3.12) can be applied, while the negative moment $\mathbb{E}[Z_{N,\beta_N}^A(0)^{-2q}]$ can be bounded by (3.15). The argument is the same as that for (4.16) and we omit the details.

To deal with (4.20) when $(x, y) \in I^> := \{x, y \in \mathbb{Z}^2 : |x - y| > N^{\frac{1}{2} - \frac{a_N}{10}}\}$, we use the chaos expansion for $Z_{N,\beta_N}^{A,B^<}$, (4.18) and write (4.20) in this case as follows (recall that $\mathbb{E}[\xi^2] = 1$):

$$(4.21) \quad \begin{aligned} & \frac{\sigma_N^2 \log N}{N^2} \sum_{x,y \in I^>} \phi \left(\frac{x}{\sqrt{N}} \right) \phi \left(\frac{y}{\sqrt{N}} \right) \sum_{\substack{(t,w) \in A_N^x \\ (s,v) \in A_N^y}} \mathbb{E} \left[\frac{Z_{0,t,\beta_N}^A(x, w)}{Z_{N,\beta_N}^A(x)} \right] \mathbb{E} \left[\frac{Z_{0,s,\beta_N}^A(y, v)}{Z_{N,\beta_N}^A(y)} \right] \\ & \times \sum_{(r,z) \in B_N: r < N^{1-9a_N/40}} q_{r-t}(z - w) q_{r-s}(z - v) \mathbb{E}[Z_{r,N,\beta_N}(z)^2], \end{aligned}$$

where the first point $(r, z) \in B_N$ in the expansion for $Z_{N,\beta_N}^A(x)$ and $Z_{N,\beta_N}^A(y)$ must match because an unmatched (r, z) gives $\mathbb{E}[\xi_{r,z}] = 0$, and we used the independence between

$$\frac{Z_{0,t,\beta_N}^A(x, w)}{Z_{N,\beta_N}^A(x)}, \quad \frac{Z_{0,s,\beta_N}^A(y, v)}{Z_{N,\beta_N}^A(y)}, \quad \text{and} \quad Z_{r,N,\beta_N}(z),$$

because they depend on disorder in the disjoint regions A_N^x, A_N^y and B_N .

We can simplify (4.21) by noticing that $\mathbb{E}[Z_{r,N,\beta_N}(z)^2]$ is independent of z and that $\sum_z q_{r-t}(z-w)q_{r-s}(z-v) = q_{2r-t-s}(w-v)$. Thus we can write it as

$$(4.22) \quad \frac{\sigma_N^2 \log N}{N^2} \sum_{x,y \in I^>} \phi\left(\frac{x}{\sqrt{N}}\right) \phi\left(\frac{y}{\sqrt{N}}\right) \sum_{\substack{(t,w) \in A_N^x \\ (s,v) \in A_N^y}} \mathbb{E}\left[\frac{Z_{0,t,\beta_N}^A(x,w)}{Z_{N,\beta_N}^A(x)}\right] \mathbb{E}\left[\frac{Z_{0,s,\beta_N}^A(y,v)}{Z_{N,\beta_N}^A(y)}\right] \\ \times \sum_{N^{1-a_N} < r < N^{1-9a_N/40}} q_{2r-t-s}(w-v) \mathbb{E}[Z_{r,N,\beta_N}(0)^2].$$

Note that $\mathbb{E}[Z_{r,N,\beta_N}(0)^2] \leq \mathbb{E}[Z_{N,\beta_N}(0)^2] \leq C_{\hat{\beta}}$ uniformly in N by (3.2). Moreover,

$$\left| \mathbb{E}\left[\frac{Z_{0,t,\beta_N}^A(x,w)}{Z_{N,\beta_N}^A(x)}\right] \right| \leq \mathbb{E}[Z_{0,t,\beta_N}^A(x,w)^2]^{1/2} \mathbb{E}\left[\frac{1}{Z_{N,\beta_N}^A(x)^2}\right]^{1/2} \leq C_{2,\hat{\beta}} \mathbb{E}[Z_{0,t,\beta_N}^A(x,w)^2]^{1/2},$$

where the constant $C_{2,\hat{\beta}}$ comes from the negative moment bound (3.15). The same bound holds with (x, t, w) replaced by (y, s, v) . Therefore, (4.22) can be bounded by

$$C_{\hat{\beta}} C_{2,\hat{\beta}} \frac{\sigma_N^2 \log N}{N^2} \sum_{x,y \in I^>} \left| \phi\left(\frac{x}{\sqrt{N}}\right) \right| \left| \phi\left(\frac{y}{\sqrt{N}}\right) \right| \\ \times \sum_{\substack{(t,w) \in A_N^x \\ (s,v) \in A_N^y}} \mathbb{E}[Z_{0,t,\beta_N}^A(x,w)^2]^{1/2} \mathbb{E}[Z_{0,s,\beta_N}^A(y,v)^2]^{1/2} \\ \times \sum_{N^{1-a_N} < r < N^{1-9a_N/40}} q_{2r-t-s}(w-v).$$

By our definitions of σ_N and β_N in (2.15) and (1.18), we have $\sigma_N^2 \log N = O(1)$. Applying Cauchy–Schwarz for the sum over (t, w) and (s, v) , we obtain the bound

$$\frac{C}{N^2} \sum_{x,y \in I^>} \left| \phi\left(\frac{x}{\sqrt{N}}\right) \right| \left| \phi\left(\frac{y}{\sqrt{N}}\right) \right| \left(\sum_{(t,w) \in A_N^x} \mathbb{E}[Z_{0,t,\beta_N}^A(x,w)^2] \sum_{(s,v) \in A_N^y} \mathbb{E}[Z_{0,s,\beta_N}^A(y,v)^2] \right)^{1/2} \\ \times \sum_{N^{1-a_N} < r < N^{1-9a_N/40}} \left(\sum_{(t,w) \in A_N^x, (s,v) \in A_N^y} q_{2r-t-s}(w-v)^2 \right)^{1/2}.$$

We next observe that $\sum_{(t,w) \in A_N^x} \mathbb{E}[Z_{0,t,\beta_N}^A(x,w)^2] = \mathbb{E}[Z_{N,\beta_N}^A(x)^2] \leq C_{\hat{\beta}}$ (see (4.13), (2.18) and (3.3)), and similarly for the sum over (s, v) . This leads to the bound

$$(4.23) \quad \frac{CC_{\hat{\beta}}}{N^2} \sum_{x,y \in I^>} \left| \phi\left(\frac{x}{\sqrt{N}}\right) \right| \left| \phi\left(\frac{y}{\sqrt{N}}\right) \right| \\ \times \sum_{N^{1-a_N} < r < N^{1-9a_N/40}} \left(\sum_{(t,w) \in A_N^x, (s,v) \in A_N^y} q_{2r-t-s}(w-v)^2 \right)^{1/2}.$$

Since $|x - y| > N^{\frac{1}{2} - \frac{a_N}{10}}$ and $|x - w|, |y - v| \leq N^{\frac{1}{2} - \frac{a_N}{4}}$, we have $|w - v| > \frac{1}{2} N^{\frac{1}{2} - \frac{a_N}{10}}$. Given $r < N^{1-9a_N/40}$, we then have

$$q_{2r-t-s}(w-v)^2 \leq c \exp\left(-\frac{|w-v|^2}{(2r-t-s)}\right) \leq c \exp(-cN^{a_N/40}) = \exp(-ce^{(\log N)^{\gamma/40}}).$$

The sums over (t, w) , (s, v) and r give only a polynomial factor in N , and hence (4.23) can be bounded by

$$\frac{c}{N^2} \sum_{x, y \in I^>} \phi\left(\frac{x}{\sqrt{N}}\right) \phi\left(\frac{y}{\sqrt{N}}\right) \cdot N^3 \exp(-ce^{(\log N)^\gamma/40}) \xrightarrow{N \rightarrow \infty} 0.$$

This proves (4.19) and completes the step.

Step (3). Let $Z_{N, \beta_N}^{A, B^\geq}(x)$ be defined as in (4.18) but with the constraint $r \geq N^{1-9a_N/40}$ instead of $r < N^{1-9a_N/40}$, that is, with the sum over $(r, z) \in B_N^\geq$ instead of B_N (recall (2.10)):

$$(4.24) \quad Z_{N, \beta_N}^{A, B^\geq}(x) := \sum_{(t, w) \in A_N^x, (r, z) \in B_N^\geq} Z_{0, t, \beta_N}^A(x, w) \cdot q_{r-t}(z - w) \cdot \sigma_N \xi_{r, z} \cdot Z_{r, N, \beta_N}(z).$$

In view of (4.6), we focus on the averaged quantity

$$(4.25) \quad \frac{\sqrt{\log N}}{N} \sum_{x \in \mathbb{Z}^2} \frac{Z_{N, \beta_N}^{A, B^\geq}(x)}{Z_{N, \beta_N}^A(x)} \phi\left(\frac{x}{\sqrt{N}}\right).$$

We will show that replacing in (4.24) the kernel $q_{r-t}(z - w)$ by $q_r(z - x)$ has a negligible effect on (4.25), in the sense that the difference tends to zero in $L^1(\mathbb{P})$.

We introduce the notation (recall (2.10))

$$B_N^\geq(x) := \{(r, z) \in B_N^\geq : |z - x| < r^{\frac{1}{2} + \frac{a_N}{80}}\}.$$

Recall that $g_t(\cdot)$ denotes the heat kernel on \mathbb{R}^2 ; see (1.12). By a refined local limit theorem for the simple random walk (see Theorem 2.3.11 in [33]), we have that for $(r, z) \in B_N^\geq(x)$,

$$q_r(z - x) = 2g_{r/2}(z - x) \exp\left(O\left(\frac{1}{r} + \frac{|z - x|^4}{r^3}\right)\right) = 2g_{r/2}(z - x) \exp(O(r^{-1 + \frac{a_N}{20}})),$$

and similarly for $(t, w) \in A_N^x$ (see (2.3)),

$$\begin{aligned} q_{r-t}(z - w) &= 2g_{(r-t)/2}(z - w) \exp\left(O\left(\frac{1}{r-t} + \frac{|z - w|^4}{(r-t)^3}\right)\right) \\ &= 2g_{(r-t)/2}(z - w) \exp(O(r^{-1 + \frac{a_N}{20}})), \end{aligned}$$

because $|z - w| \leq |z - x| + |w - x| \leq r^{\frac{1}{2} + \frac{a_N}{80}} + N^{\frac{1}{2} - \frac{a_N}{4}}$, and hence, for large N , we can bound

$$(4.26) \quad |z - w| \leq 2r^{\frac{1}{2} + \frac{a_N}{80}} \quad \text{and} \quad |r - t| \geq \frac{1}{2}r \quad \text{for } t \leq N^{1-a_N}, r \geq N^{1-9a_N/40}.$$

By a straightforward but tedious computation, there exists a positive constant c such that

$$(4.27) \quad \sup \left\{ \left| \frac{g_{r/2}(z - x)}{g_{(r-t)/2}(z - w)} - 1 \right| : r > N^{1 - \frac{9a_N}{40}}, t < N^{1-a_N}, |w - x| < N^{\frac{1}{2} - \frac{a_N}{4}}, \right. \\ \left. |z - x| < r^{\frac{1}{2} + \frac{a_N}{80}} \right\} = O(e^{-c(\log N)^\gamma})$$

as N tends to infinity, and by the local limit theorem, this bound can be transferred to the ratio $q_r(z - x)/q_{r-t}(z - w)$.

We are ready to estimate the error of replacing $q_{r-t}(z-w)$ by $q_r(z-x)$ in (4.24). We first restrict the sum on $(r, z) \in B_N^{\geq}(x)$. Then the contribution to (4.25) is

$$(4.28) \quad \frac{\sqrt{\log N}}{N} \sum_{x \in \mathbb{Z}^2} \phi\left(\frac{x}{\sqrt{N}}\right) \sum_{\substack{(t,w) \in A_N^x \\ (r,z) \in B_N^{\geq}(x)}} \frac{Z_{0,t,\beta_N}^A(x,w)}{Z_{N,\beta_N}^A(x)} \\ \times (q_{r-t}(z-w) - q_r(z-x)) \cdot \sigma_N \xi_{r,z} \cdot Z_{r,N,\beta_N}(z),$$

whose $L^1(\mathbb{P})$ norm is bounded by

$$(4.29) \quad \frac{\sqrt{\log N}}{N} \sum_{x \in \mathbb{Z}^2} \left| \phi\left(\frac{x}{\sqrt{N}}\right) \right| \mathbb{E} \left[\frac{1}{Z_{N,\beta_N}^A(x)} \right. \\ \times \left. \left| \sum_{\substack{(t,w) \in A_N^x \\ (r,z) \in B_N^{\geq}(x)}} Z_{0,t,\beta_N}^A(x,w) \left\{ 1 - \frac{q_r(z-x)}{q_{r-t}(z-w)} \right\} \right. \right. \\ \times \left. \left. q_{r-t}(z-w) \cdot \sigma_N \xi_{r,z} \cdot Z_{r,N,\beta_N}(z) \right| \right] \\ \leq \frac{\sqrt{\log N}}{N} \sum_{x \in \mathbb{Z}^2} \left| \phi\left(\frac{x}{\sqrt{N}}\right) \right| \mathbb{E} \left[\frac{1}{Z_{N,\beta_N}^A(x)^2} \right]^{1/2} \\ \times \mathbb{E} \left[\left(\sum_{\substack{(t,w) \in A_N^x \\ (r,z) \in B_N^{\geq}(x)}} Z_{0,t,\beta_N}^A(x,w) \left\{ 1 - \frac{q_r(z-x)}{q_{r-t}(z-w)} \right\} \right. \right. \\ \times \left. \left. q_{r-t}(z-w) \cdot \sigma_N \xi_{r,z} \cdot Z_{r,N,\beta_N}(z) \right)^2 \right]^{1/2}.$$

We recall that $\mathbb{E}[Z_{N,\beta_N}^A(x)^{-2}]$ is uniformly bounded by the negative moment estimate (3.15), while by orthogonality of terms in the chaos expansion and applying (4.27), the last expectation can be bounded as

$$\sum_{\substack{(t,w) \in A_N^x \\ (r,z) \in B_N^{\geq}(x)}} \mathbb{E}[Z_{0,t,\beta_N}^A(x,w)^2] \left\{ 1 - \frac{q_r(z-x)}{q_{r-t}(z-w)} \right\}^2 q_{r-t}(z-w)^2 \sigma_N^2 \mathbb{E}[Z_{r,N,\beta_N}(z)^2] \\ = O(e^{-c(\log N)^\gamma}) \sum_{\substack{(t,w) \in A_N^x \\ (r,z) \in B_N^{\geq}(x)}} \mathbb{E}[Z_{0,t,\beta_N}^A(x,w)^2] q_{r-t}(z-w)^2 \sigma_N^2 \mathbb{E}[Z_{r,N,\beta_N}(z)^2].$$

By (4.24), this last sum is bounded by $\mathbb{E}[Z_{N,\beta_N}^{A,B^{\geq}}(0)^2] \leq \mathbb{E}[Z_{N,\beta_N}(0)^2] \leq C_{\hat{\beta}}$ uniformly in N ; see (3.2). These estimates show that (4.29) is $O(\sqrt{\log N} \exp(-c(\log N)^\gamma))$, and hence converges to zero, thus the $L^1(\mathbb{P})$ norm of (4.28) converges to zero, too.

To complete the step, it remains to check that in the chaos expansion (4.24) for $Z_{N,\beta_N}^{A,B^{\geq}}(x)$, the contribution of the complementary regime $(r, z) \in B_N^{\geq} \setminus B_N^{\geq}(x)$, that is, $|z-x| \geq r^{1/2+a_N/80}$, vanishes in $L^1(\mathbb{P})$ as $N \rightarrow \infty$, and the same is true if we replace the kernel $q_{r-t}(z-w)$ by $q_r(z-x)$. Note that in this regime, by moderate deviation estimates,

$$(4.30) \quad q_r(z-x) \leq \exp\left\{-c \frac{|z-x|^2}{r}\right\} \leq \exp\{-cr^{a_N/40}\} \leq \exp\{-ce^{c(\log N)^\gamma}\},$$

and the same bound holds for $q_{r-t}(z-w)$, because $|w-x| \leq N^{\frac{1}{2}-\frac{a_N}{4}} = o(r^{1/2})$ as $N \rightarrow \infty$ (since $r \geq N^{1-9a_N/40}$), and hence $|w-x| = o(|z-x|)$ in this regime. These bounds can then be used to show that

$$\frac{\sqrt{\log N}}{N} \sum_{x \in \mathbb{Z}^2} \phi\left(\frac{x}{\sqrt{N}}\right) \sum_{\substack{(t,w) \in A_N^x, (r,z) \in B_N^{\geq} \\ |z-x| > r^{1/2+a_N/80}}} \frac{Z_{0,t,\beta_N}^A(x,w)}{Z_{N,\beta_N}^A(x)} \\ \times q_{r-t}(z-w) \cdot \sigma_N \xi_{r,z} \cdot Z_{r,N,\beta_N}(z) \xrightarrow[N \rightarrow \infty]{L^1(\mathbb{P})} 0,$$

and the same holds when $q_{r-t}(z-w)$ is replaced by $q_r(z-x)$. Indeed, we can argue as in (4.29) and then use the fact that the number of terms in the sums over $x, (t, w), (r, z)$ is only polynomial in N , while (4.30) decays faster.

4.4. *Proof of Proposition 2.4.* Recalling (2.14), we want to prove that

$$(4.31) \quad \frac{\sqrt{\log N}}{\sqrt{\pi} \hat{\beta}} \frac{1}{N} \sum_{x \in \mathbb{Z}^2} (Z_{N,\beta_N}^{B^{\geq}}(x) - 1) \phi\left(\frac{x}{\sqrt{N}}\right) \xrightarrow[N \rightarrow \infty]{d} \langle v^{(\sqrt{2}c_{\hat{\beta}})}(1/2, \cdot), \phi \rangle,$$

where $v^{(c)}(s, x)$ is the solution of the two-dimensional additive SHE as in (1.11).

The proof of (4.31) follows the same line as the proof of Theorem 2.13 in [7], which proved the convergence of the fluctuations of the polymer partition function $Z_{N,\beta_N}(x)$ as a space-time random field to the solution of the additive SHE. To see heuristically why the limit in (4.31) should be Gaussian, we can write the LHS of (4.31) as a polynomial chaos expansion (see (4.35)) where the dominant contribution (in L^2) comes from terms of finite order in the expansion because $\hat{\beta} \in (0, 1)$. Each such term is of the form $\sigma_N^k \prod_{i=1}^k q_{n_i-n_{i-1}}(x_i-x_{i-1}) \xi_{n_i,x_i}$, which due to the random walk transition kernels $q(\cdot)$, depends only on disorder ξ_{\cdot} in a neighborhood of $(n_0, x_0) := (0, x)$ that is negligible on the diffusive scale. Given such local dependence on the disorder, it is then not surprising that when averaged over $(0, x)$ on the diffusive scale with weight $\phi(x/\sqrt{N})$, we should get a Gaussian limit. The proof in [7] also shows that terms of order two and higher in the chaos expansion leads to an independent white noise in the limit, which leads to a noise coefficient $c_{\hat{\beta}} > 1$ in (1.11).

We now recall the key element in the proof of Theorem 2.13 in [7] and show how it can be adapted to our setting. The key technical tool is the following variant of Proposition 8.1 in [7], specialized to the simple random walk on \mathbb{Z}^2 (where we average in space, rather than in space-time). It will show that, in the polynomial chaos expansion of the left-hand side of (4.31), there are ‘‘building blocks’’ that converge in distribution to independent Gaussian random variables.

PROPOSITION 4.1. *For integer M and $i \in \{1, \dots, M\}$, define intervals $I_i := (N^{\frac{i-1}{M}}, N^{\frac{i}{M}}]$. A k -tuple $(i_1, \dots, i_k) \in \{1, \dots, M\}^k$ is said to belong to $\{1, \dots, M\}_{\sharp}^k$ if $|i_j - i_{j'}| \geq 2$ for all $j \neq j'$.*

For $N \in \mathbb{N}$, let $\xi = (\xi_{n,x}^{(N)})_{(n,x) \in \mathbb{N} \times \mathbb{Z}^2}$ be i.i.d. with zero mean and unit variance.

Given $N, M \in \mathbb{N}$, a k -tuple $(i_1, \dots, i_k) \in \{1, \dots, M\}_{\sharp}^k$ and a point $x \in \mathbb{Z}^2$, we define a random variable $\Theta_{i_1, \dots, i_k}^{N;M}(x)$, a multilinear polynomial of degree k in the variables ξ 's, as follows:

$$\Theta_{i_1, \dots, i_k}^{N;M}(x) := \left(\frac{M}{R_N}\right)^{\frac{k-1}{2}} \sum_{\substack{n_1 \in I_{i_1}, n_2-n_1 \in I_{i_2}, \dots, n_k-n_{k-1} \in I_{i_k} \\ n_0:=0, x_0:=x, x_1, \dots, x_k \in \mathbb{Z}^d}} \prod_{j=1}^k q_{n_j-n_{j-1}}(x_j-x_{j-1}) \prod_{i=1}^k \xi_{n_i, x_i},$$

where $q_n(x)$ is the transition kernel of the simple symmetric random walk on \mathbb{Z}^2 , and R_N is the expected overlap, defined in (1.17). For $\phi \in C_c(\mathbb{R}^2)$, we define the space-averaged version

$$\Theta_{i_1, \dots, i_k}^{N; M; \phi} := \frac{1}{N} \sum_{x \in \mathbb{Z}^2} \Theta_{i_1, \dots, i_k}^{N; M}(x) \phi\left(\frac{x}{\sqrt{N}}\right).$$

Let D_M denote the subset of $(i_1, \dots, i_k) \in \{1, \dots, M\}_{\neq}^k$ that satisfy $i_1 > \max\{i_2, \dots, i_k\}$, called dominated sequences. Then, for any fixed $M \in \mathbb{N}$ and $\phi \in C_c(\mathbb{R}^2)$, the family of random variables $(\Theta_{(i_1, \dots, i_k)}^{N; M; \phi})_{(i_1, \dots, i_k) \in D_M}$ converges in distribution as $N \rightarrow \infty$ to a family $(\zeta_{(i_1, \dots, i_k)}^\phi)_{(i_1, \dots, i_k) \in D_M}$ of independent Gaussian random variables with

$$(4.32) \quad \mathbb{E}[\zeta_{(i_1, \dots, i_k)}^\phi] = 0, \quad \text{Var}[\zeta_{(i_1, \dots, i_k)}^\phi] = 2\sigma_\phi^2 \mathbb{1}_{\{i_1 = M\}},$$

that is, the variance is nonzero only if $i_1 = M$, and is given by

$$(4.33) \quad \sigma_\phi^2 = \int_{(\mathbb{R}^2)^2} \phi(x) K_{\frac{1}{2}}(x, y) \phi(y) \, dx \, dy \quad \text{with } K_{\frac{1}{2}}(x, y) = \int_0^{\frac{1}{2}} \frac{1}{4\pi u} e^{-\frac{|x-y|^2}{4u}} \, du.$$

The proof of Proposition 4.1 in [7] is based on a variant of the fourth-moment theorem for polynomial chaos expansions, as formulated in [7], Theorem 4.2, which was obtained in [40] building on [15, 16, 41]. To check the variance, note that

$$(4.34) \quad \begin{aligned} \text{Var}[\Theta_{i_1, \dots, i_k}^{N; M; \phi}] &= \frac{1}{N^2} \sum_{x, y \in \mathbb{Z}^2} \sum_{n_1 \in I_{i_1}, x_1 \in \mathbb{Z}^2} \phi\left(\frac{x}{\sqrt{N}}\right) \phi\left(\frac{y}{\sqrt{N}}\right) q_{n_1}(x_1 - x) q_{n_1}(x_1 - y) \\ &\times \left(\frac{M}{R_N}\right)^{k-1} \sum_{\substack{n_2 - n_1 \in I_{i_2}, \dots, n_k - n_{k-1} \in I_{i_k} \\ x_2, \dots, x_k \in \mathbb{Z}^2}} \prod_{j=2}^k q_{n_j - n_{j-1}}(x_j - x_{j-1})^2, \end{aligned}$$

where the second line tends to 1 as $N \rightarrow \infty$ by the definition of R_N and I_i . We can write

$$\sum_{x_1 \in \mathbb{Z}^2} q_{n_1}(x_1 - x) q_{n_1}(x_1 - y) = q_{2n_1}(x - y) \underset{n_1 \rightarrow \infty}{=} \left(g_{n_1}(x - y) + o\left(\frac{1}{n_1}\right)\right) 2\mathbb{1}_{\{x-y \in \mathbb{Z}_{\text{even}}^2\}}$$

by the local limit theorem, where $\mathbb{Z}_{\text{even}}^2 := \{(a, b) \in \mathbb{R}^2 : a + b \text{ is even}\}$, $g_t(x)$ is as in (1.12), the factor 2 is due to random walk periodicity and we have $g_{n_1}(\cdot)$ instead of $g_{2n_1}(\cdot)$ because the random walk S_n has covariance matrix $\frac{n}{2}I$. Then, by a Riemann sum approximation, as $N \rightarrow \infty$ the first line in (4.34) is close to the integral

$$\int_{(\mathbb{R}^2)^2} \phi(x') \phi(y') \left(\int_N^{\frac{i_1}{M}-1} g_u(x' - y') \, du \right) dx' dy' \xrightarrow{N \rightarrow \infty} \begin{cases} 0 & \text{if } i_1 < M, \\ 2\sigma_\phi^2 & \text{if } i_1 = M, \end{cases}$$

with σ_ϕ^2 defined in (4.33). Also note that for $i_1 = M$, the dominant contribution comes from $n_1 \in [\varepsilon N, N]$ for ε small, and hence restricting to $n_1 \in [1, N]$, or $n_1 \in I_M = (N^{1-\frac{1}{M}}, N]$, or $n_1 \geq N^{1-9a_N/40}$ makes no difference as $N \rightarrow \infty$ (for any fixed $M \in \mathbb{N}$).

Let us show how Proposition 4.1 can be applied to prove (4.31). Recall from (2.19) that

$$(4.35) \quad Z_{N, \beta_N}^{B \geq} (x) - 1 = \sum_{k=1}^N \sigma_N^k \sum_{\substack{N^{1-9a_N/40} < n_1 < \dots < n_k \leq N \\ n_0 := 0, x_0 := x, x_1, \dots, x_k \in \mathbb{Z}^2}} \prod_{i=1}^k q_{n_i - n_{i-1}}(x_i - x_{i-1}) \xi_{n_i, x_i}.$$

For fixed $M \in \mathbb{N}$, grouping each $n_i - n_{i-1}$ according to which interval $I_j := (N^{\frac{j-1}{M}}, N^{\frac{j}{M}}]$ it belongs to, and recalling (2.15) and (1.18), we have the following approximation:

$$\begin{aligned}
 (4.36) \quad & \frac{\sqrt{\log N}}{\sqrt{\pi} \hat{\beta}} \frac{1}{N} \sum_{x \in \mathbb{Z}^2} (Z_{N, \beta_N}^{B \geq}(x) - 1) \phi\left(\frac{x}{\sqrt{N}}\right) \\
 & \approx \sum_{k=1}^M \frac{\hat{\beta}^{k-1}}{M^{(k-1)/2}} \sum_{\substack{(i_1, \dots, i_k) \in \{1, \dots, M\}_\#^k \\ i_1 = M}} \Theta_{i_1, \dots, i_k}^{N; M; \phi},
 \end{aligned}$$

where \approx means that the difference of the two sides vanishes in $L^2(\mathbb{P})$ as $N \rightarrow \infty$ followed by $M \rightarrow \infty$. The restriction $i_1 = M$ in (4.36) is due to $n_1 > N^{1-9a_N/40}$, which gives rise to a dominated sequence. The error from relaxing $n_1 > N^{1-9a_N/40}$ to $n_1 \in I_M = (N^{\frac{M-1}{M}}, N]$ is negligible in L^2 , as noted above, while the error from restricting to $(i_1, \dots, i_k) \in \{1, \dots, M\}_\#^k$ (rather than the whole $\{1, \dots, M\}^k$) is also negligible in $L^2(\mathbb{P})$, when we first send $N \rightarrow \infty$ and then $M \rightarrow \infty$, as shown in [7], Lemma 6.2.

We can then apply Proposition 4.1 to conclude that, as we let $N \rightarrow \infty$ for fixed $M \in \mathbb{N}$, the right-hand side of (4.36) converges in distribution to the same expression with $\Theta_{i_1, \dots, i_k}^{N; M; \phi}$ replaced by $\zeta_{i_1, \dots, i_k}^\phi$, that is, to a Gaussian random variable with zero mean and with variance

$$(4.37) \quad \sum_{k=1}^M \frac{(\hat{\beta}^2)^{k-1}}{M^{k-1}} 2\sigma_\phi^2 \cdot |\{(i_1, \dots, i_k) \in \{1, \dots, M\}_\#^k : i_1 = M\}|.$$

If we let $M \rightarrow \infty$, since $|\{(i_1, \dots, i_k) \in \{1, \dots, M\}_\#^k : i_1 = M\}| = M^{k-1}(1 + o(1))$, the sum in (4.37) converges to the following explicit expression, with $c_{\hat{\beta}}$ as in (1.11):

$$2\sigma_\phi^2 \sum_{k \geq 1} \hat{\beta}^{2(k-1)} = 2\sigma_\phi^2 \frac{1}{1 - \hat{\beta}^2} = (\sqrt{2}c_{\hat{\beta}})^2 \sigma_\phi^2.$$

This agrees with the variance of $\langle v^{(\sqrt{2}c_{\hat{\beta}})}(1/2, \cdot), \phi \rangle$ (see (1.13)), which proves (4.31).

5. Edwards–Wilkinson fluctuations for KPZ. In this section, we prove Theorem 1.1, which gives Edwards–Wilkinson fluctuations for the Hopf–Cole solution $h_\varepsilon(t, z) = \log u_\varepsilon(t, z)$ of the mollified KPZ (where $u_\varepsilon(t, z)$ solves the mollified SHE; see (1.5)).

The proof follows the same lines as in the directed polymer case. This is possible because $u^\varepsilon(t, z)$ admits a Feynman–Kac representation, which casts it in a form close to the directed polymer partition function of size $N = \varepsilon^{-2}t$. Indeed, by [2], Section 3 (see also [7], equation (2.27)), for fixed (t, z) we have the following equality in law:

$$\begin{aligned}
 (5.1) \quad & u^\varepsilon(t, z) \stackrel{d}{=} \mathbb{E}_{\varepsilon^{-1}z} \left[\exp \left\{ \iint_{(0, \varepsilon^{-2}t) \times \mathbb{R}^2} \left(\beta_\varepsilon j(B_s - x) \xi(s, x) \, ds \, dx - \frac{1}{2} \beta_\varepsilon^2 j(B_s - x)^2 \, ds \, dx \right) \right\} \right] \\
 & = \mathbb{E}_{\varepsilon^{-1}z} \left[\exp \left\{ \int_0^{\varepsilon^{-2}t} \int_{\mathbb{R}^2} \beta_\varepsilon j(B_s - x) \xi(s, x) \, ds \, dx - \frac{1}{2} \beta_\varepsilon^2 (\varepsilon^{-2}t) \|j\|_{L^2(\mathbb{R}^2)}^2 \right\} \right],
 \end{aligned}$$

where $B = (B_s)_{s \in [0, \infty)}$ under \mathbb{P}_x is a standard Brownian motion on \mathbb{R}^2 started at x .

We first perform a decomposition of $u^\varepsilon(t, z)$ similar to that described in Section 2, which reduces Theorem 1.1 to the four Propositions 5.1–5.4 (see Section 5.1). These are proved later (see Section 5.3) in analogy with the corresponding results for directed polymer (see Section 4), exploiting *moment bounds* analogous to those in Section 3 (see Section 5.2).

Henceforth, we set $t = 1$ and we focus on $u^\varepsilon(z) := u^\varepsilon(1, z)$.

5.1. *Decomposition, linearization and Wiener chaos.* By (5.1) and (1.21)–(1.22), $u^\varepsilon(z)$ is comparable to $Z_N(x)$, provided we identify $N = \varepsilon^{-2}$, $x = \varepsilon^{-1}z$.

As in (2.2)–(2.3), we define (for a γ^* small enough, depending only on $\hat{\beta}$ as in (2.2))

$$(5.2) \quad a_\varepsilon := \frac{1}{(\log \varepsilon^{-2})^{1-\gamma}} \quad \text{for fixed } \gamma \in (0, \gamma^*),$$

$$(5.3) \quad A_\varepsilon^z := \{(s, x) : 0 < s \leq (\varepsilon^{-2})^{1-a_\varepsilon}, |x - \varepsilon^{-1}z| < (\varepsilon^{-2})^{\frac{1-a_\varepsilon}{2}}\},$$

and we introduce a modified partition function $u_A^\varepsilon(z)$, obtained by restricting the double integral in the first line of (5.1) to the set $(s, x) \in A_\varepsilon^z$. This yields the decomposition

$$(5.4) \quad u^\varepsilon(z) = u_A^\varepsilon(z) + \hat{u}_A^\varepsilon(z),$$

where $\hat{u}_A^\varepsilon(z)$, defined by this relation, is a “remainder” which, for fixed z , can be shown to be much smaller than $u_A^\varepsilon(z)$. More precisely, as in (2.6)–(2.7), we define $O^\varepsilon(z)$ by

$$(5.5) \quad \log u^\varepsilon(z) = \log u_A^\varepsilon(z) + \frac{\hat{u}_A^\varepsilon(z)}{u_A^\varepsilon(z)} + O^\varepsilon(z),$$

and we have the following analogues of Propositions 2.1–2.2.

PROPOSITION 5.1. *Let $O^\varepsilon(\cdot)$ be defined as above, then for any $\phi \in C_c(\mathbb{R}^2)$,*

$$\sqrt{\log \varepsilon^{-1}} \int_{\mathbb{R}^2} (O^\varepsilon(z) - \mathbb{E}[O^\varepsilon(z)])\phi(z) \, dz \xrightarrow[\varepsilon \downarrow 0]{L^2(\mathbb{P})} 0.$$

PROPOSITION 5.2. *Let $u_A^\varepsilon(\cdot)$ be defined as above, then for any $\phi \in C_c(\mathbb{R}^2)$,*

$$\sqrt{\log \varepsilon^{-1}} \int_{\mathbb{R}^2} (\log u_A^\varepsilon(z) - \mathbb{E}[\log u_A^\varepsilon(z)])\phi(z) \, dz \xrightarrow[\varepsilon \downarrow 0]{L^2(\mathbb{P})} 0.$$

Next, in analogy with (2.10)–(2.11), we introduce the subset

$$(5.6) \quad B_\varepsilon^{\geq} := ((\varepsilon^{-2})^{1-9a_\varepsilon/40}, \varepsilon^{-2}) \times \mathbb{R}^2,$$

and we introduce $u_{B^{\geq}}^\varepsilon(z)$, obtained by restricting the double integral in the first line of (5.1) to the set $(s, x) \in B_\varepsilon^{\geq}$. We have the following analogues of Propositions 2.3–2.4.

PROPOSITION 5.3. *Let $u_A^\varepsilon(\cdot)$, $\hat{u}_A^\varepsilon(\cdot)$, $u_{B^{\geq}}^\varepsilon(\cdot)$ be defined as above, then for any $\phi \in C_c(\mathbb{R}^2)$,*

$$(5.7) \quad \sqrt{\log \varepsilon^{-1}} \int_{\mathbb{R}^2} \left(\frac{\hat{u}_A^\varepsilon(z)}{u_A^\varepsilon(z)} - (u_{B^{\geq}}^\varepsilon(z) - 1) \right) \phi(z) \, dz \xrightarrow[\varepsilon \downarrow 0]{L^1(\mathbb{P})} 0.$$

PROPOSITION 5.4. *Let $u_{B^{\geq}}^\varepsilon(\cdot)$ be defined as above, then for any $\phi \in C_c(\mathbb{R}^2)$,*

$$(5.8) \quad \frac{\sqrt{\log \varepsilon^{-1}}}{\sqrt{2\pi\hat{\beta}}} \int_{\mathbb{R}^2} (u_{B^{\geq}}^\varepsilon(z) - 1)\phi(z) \, dz \xrightarrow[\varepsilon \downarrow 0]{d} (v^{(c\hat{\beta})}(1, \cdot), \phi).$$

Theorem 1.1 is a direct consequence of Propositions 5.1–5.4. Regarding the centering, by (5.5) we have $\mathbb{E}[\log u^\varepsilon(z)] = \mathbb{E}[\log u_A^\varepsilon(z)] + \mathbb{E}[O^\varepsilon(z)]$, because $\hat{u}_A^\varepsilon(z)/u_A^\varepsilon(z)$ has zero mean, as we show in a moment.

By (5.1) and the definition of Wick exponential [30], Section 3.2, we have the following Wiener chaos representation for $u^\varepsilon(z)$, where we set $t_0 := 0$ and $y_0 := \varepsilon^{-1}z$:

$$(5.9) \quad u^\varepsilon(z) \stackrel{d}{=} 1 + \sum_{k \geq 1} \beta_\varepsilon^k \int \cdots \int_{\substack{0 < t_1 < \dots < t_k < \varepsilon^{-2} \\ \vec{x} \in (\mathbb{R}^2)^k}} \left(\int_{(\mathbb{R}^2)^k} \prod_{i=1}^k g_{t_i - t_{i-1}}(y_i - y_{i-1}) j(y_i - x_i) d\vec{y} \right) \times \prod_{i=1}^k \xi(t_i, x_i) dt_i dx_i,$$

where $g_t(\cdot)$ is the transition kernel of the Brownian motion.

The modified partition function $u_A^\varepsilon(z)$ admits a similar Wiener chaos expansions, with the outer integrals restricted to the set $\{(t_1, x_1), \dots, (t_k, x_k)\} \subseteq A_\varepsilon^z$. It follows that the Wiener chaos expansion of $\hat{u}_A^\varepsilon(z) := u^\varepsilon(z) - u_A^\varepsilon(z)$ contains at least one factor $\xi(t_i, x_i)$ with (t_i, x_i) outside A_ε^z , which is not present in $u_A^\varepsilon(z)$, hence $\mathbb{E}[\hat{u}_A^\varepsilon(z)/u_A^\varepsilon(z)] = 0$.

Similarly, the Wiener chaos expansions of $u_{B_\varepsilon^z}^\varepsilon(z)$ is obtained by restricting the outer integrals in (5.9) to the set $\{(t_1, x_1), \dots, (t_k, x_k)\} \subseteq B_\varepsilon^z$, that is, imposing $t_1 > (\varepsilon^{-2})^{1-9a_\varepsilon/40}$.

5.2. *Moment bounds.* We estimate positive and negative moments of $u^\varepsilon(z)$.

We start with the second moment. We prove below the following bounds for $u^\varepsilon(z)$, $u_A^\varepsilon(z)$ and $\hat{u}_A^\varepsilon(z)$, which are close analogues of (3.2), (3.3), (3.4):

$$(5.10) \quad \forall \hat{\beta} \in (0, 1) \exists C_{\hat{\beta}} < \infty \text{ such that } \forall \varepsilon > 0 : \\ \mathbb{E}[u^\varepsilon(z)^2] \leq C_{\hat{\beta}}, \quad \mathbb{E}[u_A^\varepsilon(z)^2] \leq C_{\hat{\beta}}, \quad \mathbb{E}[\hat{u}_A^\varepsilon(z)^2] \leq C_{\hat{\beta}} a_\varepsilon.$$

We can now easily deduce bounds for higher positive moments. By hypercontractivity [30], Theorem 5.1, the L^p norm of a Wiener chaos expansion like (5.9) is bounded by the L^2 norm of a modified expansion, with the k th order term multiplied by $(c_p)^k$ (i.e., $\hat{\beta}$ replaced by $c_p \hat{\beta}$), with $c_p := \sqrt{p-1}$. For $\hat{\beta} \in (0, 1)$, we can choose $p > 2$ such that $\hat{\beta} c_p < 1$, so as to apply the bounds in (5.10). This yields an analogue of (3.12):

$$(5.11) \quad \forall \hat{\beta} \in (0, 1) \exists p = p_{\hat{\beta}} \in (2, \infty) \exists C'_{\hat{\beta}} < \infty \text{ such that } \forall \varepsilon > 0 : \\ \mathbb{E}[u^\varepsilon(z)^p] \leq C'_{\hat{\beta}}, \quad \mathbb{E}[u_A^\varepsilon(z)^p] \leq C'_{\hat{\beta}}, \quad \mathbb{E}[|\hat{u}_A^\varepsilon(z)|^p] \leq C'_{\hat{\beta}} (a_\varepsilon)^{p/2}.$$

PROOF OF (5.10). We compute $\mathbb{E}[u^\varepsilon(z)^2]$ by using (5.9), applying the identity $g_t(y) \times g_t(y') = 4g_{2t}(y - y')g_{2t}(y + y')$, and switching to new variables $z_i := y_i - y'_i$, $w_i := y_i + y'_i$. This leads to the following expression (see [8], Section 8.2, for details):

$$(5.12) \quad \mathbb{E}[u^\varepsilon(z)^2] = 1 + \sum_{k \geq 1} (\beta_\varepsilon^2)^k \int \cdots \int_{\substack{0 < t_1 < \dots < t_k < \varepsilon^{-2} \\ \vec{z} \in (\mathbb{R}^2)^k, \vec{w} \in (\mathbb{R}^2)^k}} \left(\prod_{i=1}^k g_{2(t_i - t_{i-1})}(z_i - z_{i-1}) J(z_i) \cdot g_{2(t_i - t_{i-1})}(w_i - w_{i-1}) \right) \prod_{i=1}^k dt_i dz_i dw_i,$$

where $J := j * j$ and we set $z_0 := 0$, $w_0 := 2\varepsilon^{-1}z$. Integrating out w_k, w_{k-1}, \dots, w_1 , we get

$$\mathbb{E}[u^\varepsilon(z)^2] = 1 + \sum_{k \geq 1} (\beta_\varepsilon^2)^k \int \cdots \int_{\substack{0 < t_1 < \dots < t_k < \varepsilon^{-2} \\ \vec{z} \in (\mathbb{R}^2)^k}} \left(\prod_{i=1}^k g_{2(t_i - t_{i-1})}(z_i - z_{i-1}) J(z_i) \right) \prod_{i=1}^k dt_i dz_i.$$

We recall that j , hence J , has compact support. If we define

$$\bar{r}(t) := \sup_{z' \in \text{supp}(J)} \int_{\mathbb{R}^2} g_{2t}(z - z') J(z) dz,$$

we can bound

$$\begin{aligned} \mathbb{E}[u^\varepsilon(z)^2] &\leq 1 + \sum_{k \geq 1} (\beta_\varepsilon^2)^k \int \cdots \int_{0 < t_1 < \cdots < t_k < \varepsilon^{-2}} \prod_{i=1}^k \bar{r}(t_i - t_{i-1}) dt_i \\ &\leq 1 + \sum_{k \geq 1} \left\{ \beta_\varepsilon^2 \int_0^{\varepsilon^{-2}} \bar{r}(t) dt \right\}^k. \end{aligned}$$

Note that $\bar{r}(\cdot)$ is bounded for $t \geq 0$ and it satisfies $\bar{r}(t) = \frac{1}{4\pi t} + O(1)$ as $t \rightarrow \infty$, by (1.12). Recalling (1.2), we see that the bracket converges to $\hat{\beta}^2$ as $\varepsilon \rightarrow 0$, hence the series is uniformly bounded for $\hat{\beta} < 1$. This proves the first bound in (5.10).

The second bound in (5.10) follows because $\mathbb{E}[u_A^\varepsilon(z)^2] \leq \mathbb{E}[u^\varepsilon(z)^2]$, since the Wiener chaos expansion for $u_A^\varepsilon(z)$ is a subset of the expansion for $u^\varepsilon(z)$.

Finally, the third bound in (5.10) can be proved similarly to (3.4) (see Section 3.4), because $\mathbb{E}[\hat{u}_A^\varepsilon(z)^2]$ can be bounded by an expression analogous to (5.12).¹ \square

We next estimate *negative moments*, establishing the following analogues of (3.14)–(3.16):

$$\forall \hat{\beta} \in (0, 1) \forall p \in (0, \infty) \exists C_{p, \hat{\beta}} < \infty \text{ such that } \forall \varepsilon > 0 :$$

$$(5.13) \quad \mathbb{E}[u^\varepsilon(z)^{-p}] \leq C_{p, \hat{\beta}} < \infty,$$

$$(5.14) \quad \mathbb{E}[u_A^\varepsilon(z)^{-p}] \leq C_{p, \hat{\beta}} < \infty,$$

$$(5.15) \quad \mathbb{E}[|\log u_A^\varepsilon(z)|^p] \leq C_{p, \hat{\beta}} < \infty.$$

Since (5.15) follows easily from (5.14), it suffices to prove (5.13)–(5.14). These are direct corollaries of the following result, analogous to Proposition 3.1.

PROPOSITION 5.5 (Left tail for KPZ). *For $\Lambda \subseteq (0, \varepsilon^{-2}) \times \mathbb{R}^2$, denote by $u_\Lambda^\varepsilon(z)$ what we obtain by restricting the double integral in the first line of (5.1) to $(s, x) \in \Lambda$, that is,*

$$(5.16) \quad \begin{aligned} u_\Lambda^\varepsilon(z) := \mathbb{E}_{\varepsilon^{-1}z} \left[\exp \left\{ \iint_{(s,x) \in \Lambda} \left(\beta_\varepsilon j(B_s - x) \xi(s, x) ds dx \right. \right. \right. \\ \left. \left. \left. - \frac{1}{2} \beta_\varepsilon^2 j(B_s - x)^2 ds dx \right) \right\} \right]. \end{aligned}$$

For any $\hat{\beta} \in (0, 1)$ there is $c_{\hat{\beta}} \in (0, \infty)$ with the following property: for any $\varepsilon > 0$ and for any choice of subset $\Lambda \subseteq (0, \varepsilon^{-2}) \times \mathbb{R}^2$, one has

$$(5.17) \quad \forall t \geq 0: \quad \mathbb{P}(\log u_\Lambda^\varepsilon(z) \leq -t) \leq c_{\hat{\beta}} e^{-t^2/c_{\hat{\beta}}}.$$

It remains to prove Proposition 5.5. We first need to recall concentration inequalities for white noise (see Appendix C for more details).

The white noise $\xi = (\xi(s, y))_{(s,y) \in [0, \infty) \times \mathbb{R}^2}$ can be viewed as a random element of a separable Banach space E of distributions on $[0, \infty) \times \mathbb{R}^2$ (e.g., a negative Hölder space; see [10]).

¹Note that $\hat{u}_A^\varepsilon(z)$ contains at least one point (t_i, x_i) outside A_ε^z in the Wiener chaos representation (5.9). Since $j(\cdot)$ has compact support, say included in the ball $B_r := \{x \in \mathbb{R}^2 : |x| \leq r\}$, the corresponding point (t_i, y_i) in (5.9) must be close to (i.e., at distance at most r from) the point (t_i, x_i) . Then $\mathbb{E}[\hat{u}_A^\varepsilon(z)^2]$ can be bounded by an expression analogous to (5.12), with the integrals restricted to the set where at least one point $(t_i, \frac{1}{2}w_i) = (t_i, \frac{1}{2}(y_i + y'_i))$ is close to $(A_\varepsilon^z)^c$. This allows us to follow the proof in Section 3.4.

Its law μ is the Gaussian measure on E with Cameron–Martin space $H = L^2([0, \infty) \times \mathbb{R}^2)$, and the triple (H, E, μ) is a so-called *abstract Wiener space*. In this setting, sharp concentration inequalities are known to hold for (not necessarily convex) functions $f : E \rightarrow \mathbb{R}$ that are Lipschitz in the directions of H , see [34], equations (4.7) and (4.8).

We need to work with convex functions $f : E \rightarrow \mathbb{R} \cup \{-\infty, +\infty\}$ that are not globally Lipschitz. Remarkably, such functions still enjoy concentration inequalities for the *left tail* (but not, in general, for the right tail). For $x \in E$ with $|f(x)| < \infty$, denote by $|\nabla f(x)| \in [0, \infty]$ the maximal gradient of f in the directions of H , defined by

$$(5.18) \quad |\nabla f(x)| := \sup_{h \in H: \|h\|_H \leq 1} \lim_{\delta \downarrow 0} \frac{|f(x + \delta h) - f(x)|}{\delta},$$

where the limit exists by convexity. Then the following inequality holds (see Theorem C.1):

$$(5.19) \quad \mu(f \leq a - t) \mu^*(f \geq a, |\nabla f| \leq c) \leq e^{-\frac{1}{4}(t/c)^2} \quad \forall a \in \mathbb{R}, \forall t, c \in (0, \infty),$$

where μ^* is the outer measure (to avoid the issue of measurability of $|\nabla f|$).

Note that, if we fix a, c such that $\mu^*(f \geq a, |\nabla f| \leq c) > 0$, relation (5.19) gives a bound on the left tail $\mu(f \leq a - t)$ for all $t > 0$.

PROOF OF PROPOSITION 5.5. We can set $z = 0$, since the law of $u_\Lambda^\varepsilon(z)$ in (5.16) does not depend on z , and we write $u_\Lambda^\varepsilon := u_\Lambda^\varepsilon(0)$. We denote by $\mathcal{H}_\varepsilon^\xi(B)$ the argument of the exponential in (5.16), so that $u_\Lambda^\varepsilon = \mathbb{E}[\exp(\mathcal{H}_\varepsilon^\xi(B))]$. We also introduce the shorthand

$$(5.20) \quad \langle j(B), \xi \rangle := \iint_{(s,x) \in \Lambda} j(B_s - x) \xi(s, x) \, ds \, dx.$$

We start with a second moment computation:

$$(5.21) \quad \mathbb{E}[(u_\Lambda^\varepsilon)^2] = \mathbb{E}[\mathbb{E}[e^{\mathcal{H}_\varepsilon^\xi(B) + \mathcal{H}_\varepsilon^\xi(\tilde{B})}]] = \mathbb{E}[e^{\beta_\varepsilon^2 \mathcal{L}_\Lambda(B, \tilde{B})}],$$

where B, \tilde{B} are independent Brownian motions, and $\mathcal{L}_\Lambda(B, \tilde{B})$ is their overlap on Λ :

$$(5.22) \quad \mathcal{L}_\Lambda(B, \tilde{B}) := \iint_{(s,x) \in \Lambda} j(B_s - x) j(\tilde{B}_s - x) \, ds \, dx.$$

Note that u_Λ^ε is a function of the white noise ξ , so we can define

$$(5.23) \quad h_\varepsilon(\xi) := \log u_\Lambda^\varepsilon.$$

The function $h_\varepsilon(\cdot)$ is convex by Hölder’s inequality, because $\xi \mapsto \langle j(B), \xi \rangle$ is linear (more precisely, we can ensure that $h_\varepsilon(\cdot)$ is convex by choosing a suitable version of the stochastic integral $\langle j(B), \xi \rangle$; see Appendix D). Then (5.17) follows by (5.19) if we show that $\mu^*(h_\varepsilon \geq a, |\nabla h_\varepsilon| \leq c)$ is uniformly bounded from below, for $a = -\log 2$ and for suitable $c = c_{\hat{\beta}}$.

We need to evaluate the maximal gradient $|\nabla h_\varepsilon(\xi)|$; see (5.18). We define a Gibbs change of measure \mathbf{P}^ξ on the Brownian path $B = (B_s)_{s \geq 0}$ by

$$\frac{d\mathbf{P}^\xi}{d\mathbf{P}}(B) := \frac{e^{\mathcal{H}_\varepsilon^\xi(B)}}{u_\Lambda^\varepsilon}.$$

Let us fix $f \in H = L^2([0, \infty) \times \mathbb{R}^2)$. Recalling (5.20) and (5.23), we have

$$\begin{aligned} \lim_{\delta \downarrow 0} \frac{h_\varepsilon(\xi + \delta f) - h_\varepsilon(\xi)}{\delta} &= \lim_{\delta \downarrow 0} \frac{1}{\delta} \log \mathbf{E}^\xi [e^{\beta_\varepsilon \langle j(B), \delta f \rangle}] = \beta_\varepsilon \mathbf{E}^\xi [\langle j(B), f \rangle] \\ &= \beta_\varepsilon \iint_{(s,x) \in \Lambda} \mathbf{E}^\xi [j(B_s - x)] f(s, x) \, ds \, dx. \end{aligned}$$

Taking f with $\|f\|_{L^2} \leq 1$ and recalling (5.22), it follows by Cauchy–Schwarz that

$$|\nabla h_\varepsilon(\xi)|^2 \leq \beta_\varepsilon^2 \iint_{(s,x) \in \Lambda} \mathbf{E}^\xi [j(B_s - y)]^2 ds dy = \frac{\mathbb{E}[\beta_\varepsilon^2 \mathcal{L}_\Lambda(B, \tilde{B}) e^{\mathcal{H}_\varepsilon^\xi(B) + \mathcal{H}_\varepsilon^\xi(\tilde{B})}]}{(u_\Lambda^\varepsilon)^2}.$$

Then, on the event $h_\varepsilon(\xi) > a = -\log 2$, that is, $u_\Lambda^\varepsilon > \frac{1}{2}$, recalling (5.21), we can bound

$$(5.24) \quad \begin{aligned} \mathbb{E}[|\nabla h_\varepsilon(\xi)|^2 \mathbb{1}_{\{h_\varepsilon(\xi) > a\}}] &\leq 4\mathbb{E}\mathbb{E}[\beta_\varepsilon^2 \mathcal{L}_\Lambda(B, \tilde{B}) e^{\mathcal{H}_\varepsilon(B) + \mathcal{H}_\varepsilon(\tilde{B})}] \\ &\leq 4\mathbb{E}[\beta_\varepsilon^2 \mathcal{L}_\Lambda(B, \tilde{B}) e^{\beta_\varepsilon^2 \mathcal{L}_\Lambda(B, \tilde{B})}] \leq \frac{4}{\delta} \mathbb{E}[e^{(1+\delta)\beta_\varepsilon^2 \mathcal{L}_\Lambda(B, \tilde{B})}], \end{aligned}$$

for any $\delta > 0$ (by $x \leq \frac{1}{\delta} e^{\delta x}$). For any subcritical $\hat{\beta} \in (0, 1)$, we can fix $\delta = \delta_{\hat{\beta}} > 0$ small, so that $\hat{\beta}' := \hat{\beta} \sqrt{1 + \delta} < 1$ is still subcritical. By (5.21), the last expected value in (5.24) is the second moment of u_Λ^ε with $\hat{\beta}'$ instead of $\hat{\beta}$, hence it is uniformly bounded by some constant $C_{\hat{\beta}} < \infty$, by (5.10), uniformly over all subsets $\Lambda \subseteq (0, \varepsilon^{-2}) \times \mathbb{R}^2$. Summarizing,

$$(5.25) \quad \sup_{\varepsilon > 0} \mathbb{E}[|\nabla h_\varepsilon(\xi)|^2 \mathbb{1}_{\{h_\varepsilon(\xi) > a\}}] \leq C'_{\hat{\beta}} < \infty.$$

We can continue as in the directed polymer case (see Proposition 3.1), noting that

$$(5.26) \quad \begin{aligned} \mu(h_\varepsilon \geq a, |\nabla h_\varepsilon| \leq c) &= \mu(h_\varepsilon \geq a) - \mu(h_\varepsilon \geq a, |\nabla h_\varepsilon| > c) \\ &\geq \mu(h_\varepsilon \geq a) - \frac{1}{c^2} \mathbb{E}[|\nabla h_\varepsilon(\xi)|^2 \mathbb{1}_{\{h_\varepsilon(\xi) > a\}}]. \end{aligned}$$

Since $a := -\log 2$, we have $\mu(h_\varepsilon \geq a) = \mu(u_\Lambda^\varepsilon \geq \frac{1}{2}) \geq (4C_{\hat{\beta}})^{-1}$ as in (3.21). Plugging this bound together with (5.25) into (5.26), we are done by choosing $c = c_{\hat{\beta}}$ large enough. \square

5.3. *Proof of Propositions 5.1–5.4.* Propositions 5.1 and 5.2 are proved repeating almost verbatim the proofs of Propositions 2.1 and 2.2, which are the corresponding results for directed polymers. We omit the details and refer to Sections 4.1 and 4.2.

PROOF OF PROPOSITION 5.3.. We follow closely the proof of Proposition 2.3 in Section 4.3. Recall the decomposition $u^\varepsilon(z) = u_{A,C}^\varepsilon(z) + \hat{u}_A^\varepsilon(z)$; see (5.4). Then we further decompose

$$(5.27) \quad \hat{u}_A^\varepsilon(z) = u_{A,C}^\varepsilon(z) + u_{A,B}^\varepsilon(z),$$

where $u_{A,C}^\varepsilon(z), u_{A,B}^\varepsilon(z)$ are defined in analogy with $Z_{N,\beta_N}^{A,B}(x), Z_{N,\beta_N}^{A,C}(x)$ from (4.9), (4.10):

$$(5.28) \quad \begin{aligned} u_{A,C}^\varepsilon(z) &:= \sum_{k \geq 1} \beta_\varepsilon^k \int \cdots \int_{\substack{0 < t_1 < \cdots < t_k < \varepsilon^{-2}, \vec{x} \in (\mathbb{R}^2)^k \\ \{(t_1, x_1), \dots, (t_k, x_k)\} \cap C_\varepsilon^z \neq \emptyset}} \left(\int_{(\mathbb{R}^2)^k} \prod_{i=1}^k g_{t_i - t_{i-1}}(y_i - y_{i-1}) j(y_i - x_i) d\vec{y} \right) \\ &\quad \times \prod_{i=1}^k \xi(t_i, x_i) dt_i dx_i, \\ u_{A,B}^\varepsilon(z) &:= \sum_{k \geq 1} \beta_\varepsilon^k \int \cdots \int_{\substack{0 < t_1 < \cdots < t_k < \varepsilon^{-2}, \vec{x} \in (\mathbb{R}^2)^k \\ \{(t_1, x_1), \dots, (t_k, x_k)\} \subset A_\varepsilon^z \cup B_\varepsilon^z \\ \{(t_1, x_1), \dots, (t_k, x_k)\} \cap B_\varepsilon^z \neq \emptyset}} \left(\int_{(\mathbb{R}^2)^k} \prod_{i=1}^k g_{t_i - t_{i-1}}(y_i - y_{i-1}) j(y_i - x_i) d\vec{y} \right) \\ &\quad \times \prod_{i=1}^k \xi(t_i, x_i) dt_i dx_i, \end{aligned}$$

where we set $t_0 := 0, y_0 := \varepsilon^{-1}z$, we recall that A_ε^z was defined in (5.3), while $B_\varepsilon, C_\varepsilon^z$ are defined similarly to B_N, C_N^x from (4.7), (4.8) with $N = \varepsilon^{-2}$ and $x = \varepsilon^{-1}z$: more precisely, recalling a_ε from (5.2), we set

$$B_\varepsilon := ((\varepsilon^{-2})^{1-a_\varepsilon}, \varepsilon^{-2}],$$

$$C_\varepsilon^z := \{(t, x) \in \mathbb{R}^2 : 0 < t \leq (\varepsilon^{-2})^{1-a_\varepsilon}, |x - \varepsilon^{-1}z| \geq (\varepsilon^{-2})^{1-\frac{a_\varepsilon}{4}}\}.$$

The proof of Proposition 5.3, similarly to Proposition 2.3, proceeds in three steps.

The first step is to show that $u_{A,C}^\varepsilon(x)$ in (5.27) gives a negligible contribution, that is,

$$(5.28) \quad \sqrt{\log \varepsilon^{-1}} \int_{\mathbb{R}^2} \frac{u_{A,C}^\varepsilon(z)}{u_A^\varepsilon(z)} \phi(z) dz \xrightarrow[\varepsilon \rightarrow 0]{L^2(\mathbb{P})} 0.$$

The proof is identical to the case for directed polymer; see (4.15) and the following lines. The only difference is that (4.17) will be replaced by its continuum analogue, which is

$$E \left[\exp \left\{ (c_{2p} \beta_\varepsilon)^2 \int_0^{\varepsilon^{-2(1-a_\varepsilon)}} J(B_s^{(1)} - B_s^{(2)}) ds \right\}; \sup_{s \leq \varepsilon^{-2(1-a_\varepsilon)}} |B^{(i)}(s)| > \varepsilon^{-(1-\frac{a_\varepsilon}{2})}, \text{ for } i = 1, 2 \right],$$

where $c_{2p} := \sqrt{2p - 1}$ is the hypercontractivity constant for white noise, $B^{(1)}, B^{(2)}$ are two independent Brownian motions and we recall that $J(\cdot) = (j * j)(\cdot)$. The rest of the estimates follow the same lines as in the polymer case.

In view of (5.27) and (5.28), to complete the proof it remains to show that

$$(5.29) \quad \sqrt{\frac{\log \varepsilon^{-1}}{2\pi}} \left\{ \int_{\mathbb{R}^2} \frac{u_{A,B}^\varepsilon(z)}{u_A^\varepsilon(z)} \phi(z) dz - \int_{\mathbb{R}^2} (u_{B \geq}^\varepsilon(z) - 1) \phi(z) dz \right\} \xrightarrow[\varepsilon \rightarrow 0]{L^1(\mathbb{P})} 0.$$

For $u_{A,B}^\varepsilon(z)$, we can give an expression analogous to (4.12), integrating over the last point $(t, w) \in A_\varepsilon^z$ and the first point $(r, v) \in B_\varepsilon$:

$$(5.30) \quad u_{A,B}^\varepsilon(z) = \int \cdots \int_{\substack{(t,w) \in A_\varepsilon^z, w' \in \mathbb{R}^2 \\ (r,v) \in B_\varepsilon, v' \in \mathbb{R}^2}} u_A^\varepsilon(0, z; dt, dw, dw') \\ \cdot g_{r-t}(v' - w') j(v' - v) \beta_\varepsilon \xi(r, v) dr dv dv' \\ \cdot u^\varepsilon(r, v'; \varepsilon^{-2}, \cdot),$$

where $u_A^\varepsilon(0, z; dt, dw, dw')$ is the ‘‘point-to-point’’ partition function from $(0, z)$ to (t, w, w') , similar to (4.13) (the extra space variable w' is due to the convolution with $j(\cdot)$), which is defined as follows, where we set $t_0 := 0$ and $y_0 := \varepsilon^{-1}z$:

$$u_A^\varepsilon(0, z; dt, dw, dw') \\ := \sum_{k \geq 1} \beta_\varepsilon^k \left\{ \int \cdots \int_{\substack{0 < t_1 < \cdots < t_{k-1} < \varepsilon^{-2} \\ (x_1, \dots, x_{k-1}) \in (\mathbb{R}^2)^{k-1} \\ \{(t_1, x_1), \dots, (t_{k-1}, x_{k-1})\} \subset A_\varepsilon^z}} \left(\int_{(\mathbb{R}^2)^{k-1}} \prod_{i=1}^{k-1} g_{t_i - t_{i-1}}(y_i - y_{i-1}) j(y_i - x_i) \right. \right. \\ \cdot g_{t - t_{k-1}}(w' - y_{k-1}) d\vec{y} \left. \right) \prod_{i=1}^{k-1} \xi(t_i, x_i) dt_i dx_i \left. \right\} \\ \cdot j(w' - w) \xi(t, w) dt dw dw',$$

while $u^\varepsilon(r, v'; \varepsilon^{-2}, \cdot)$ is the ‘‘point-to-plane’’ partition from (r, v') until time ε^{-2} , defined by (5.9) where we set $t_0 := r, y_0 := v'$ and we replace $0 < t_1 < \cdots$ by $r < t_1 < \cdots$.

In order to prove (5.29), as in the polymer case, we need two more steps: the second step is to prove that the contribution from $r < (\varepsilon^{-2})^{1-9a_\varepsilon/40}$ to the decomposition (5.30) is negligible; the third step is to show that we can replace $g_{r-t}(z' - w')$ by $g_r(z' - \varepsilon^{-1}z)$ in (5.30), because their difference is negligible for (5.29). These steps are proved using exactly the same analysis as in the polymer case; see Section 4.3.

Finally, when we restrict the integral in (5.30) to $r \geq (\varepsilon^{-2})^{1-9a_\varepsilon/40}$, that is, to $(r, z) \in B_\varepsilon^\geq$ (recall (5.6)), and we replace $g_{r-t}(z' - w')$ by $g_r(z' - \varepsilon^{-1}z)$, the right-hand side of (5.30) becomes exactly $u_A^\varepsilon(z)(u_{B^\geq}^\varepsilon(z) - 1)$, which proves (5.29). \square

PROOF OF PROPOSITION 5.4. In principle, also this last result could be proved as in the polymer case (see Section 4.4), using a continuum analogue of Proposition 4.1. However, it is simpler to deduce it from Proposition 2.4, approximating $u_{B^\geq}^\varepsilon(z)$ in $L^2(\mathbb{P})$ by a directed polymer partition function $Z_{N, \beta_N}^{B^\geq}(x)$ with $N = \varepsilon^{-2}$, $x = \varepsilon^{-1}z$ built on the same probability space. The details are described in Section 9 in [7] (where the space-time fluctuations of $u^\varepsilon(\cdot, \cdot)$ are shown to converge to the solution of the additive SHE). \square

APPENDIX A: SCALING RELATIONS FOR KPZ

We prove a scaling relation between the solutions of the mollified KPZ equations with different parameters; see also [10], Section 2. In particular, we will verify the identity (1.8) which relates the solution of the mollified KPZ equation with the small parameter β_ε either in front of the noise or in front of the nonlinearity.

Given $\nu, \lambda, D > 0$, let $\psi^\varepsilon := \psi^{\varepsilon; \nu, \lambda, D}$ denote the solution of the mollified KPZ equation

$$(A.1) \quad \partial_t \psi^\varepsilon = \frac{\nu}{2} \Delta \psi^\varepsilon + \frac{\lambda}{2} |\nabla \psi^\varepsilon|^2 + \sqrt{D} \xi^\varepsilon, \quad x \in \mathbb{R}^2, t \geq 0, \text{ and } \psi^\varepsilon(0, \cdot) \equiv 0,$$

where $\xi^\varepsilon(t, x)$ is the mollification of the space-time white noise ξ in space with $j_\varepsilon(x) = \varepsilon^{-2}j(x/\varepsilon)$, and $j \in C_c(\mathbb{R}^2)$ is a probability density on \mathbb{R}^2 with $j(x) = j(-x)$.

PROPOSITION A.1. *Let $\psi^{\varepsilon; \nu, \lambda, D}$ be defined as above. Then for any $a > 0$, we have*

$$(A.2) \quad (\psi^{\varepsilon; \nu, \lambda, D}(t, x))_{t \geq 0, x \in \mathbb{R}^2} \stackrel{\text{dist}}{=} \left(\frac{\nu}{\lambda} \psi^{a\varepsilon; 1, 1, \beta^2}(a^2 \nu t, ax) \right)_{t \geq 0, x \in \mathbb{R}^2},$$

where $\beta^2 := \frac{\lambda^2 D}{\nu^3}$, known as the effective coupling constant; see [5].

REMARK A.2. In (A.2), setting $a = 1$, $\nu = 1$, $\lambda := \beta_\varepsilon = \hat{\beta} \sqrt{\frac{2\pi}{\log \varepsilon^{-1}}}$ and $D = 1$ gives (1.8), since the constant term C_ε in (1.3) only shifts the solution deterministically.

We need the following scaling relation for the mollified white noise ξ^ε .

LEMMA A.3. *Let ξ be the space-time white noise on $\mathbb{R} \times \mathbb{R}^2$ and let $\xi^\varepsilon := \xi * j_\varepsilon$, where $j_\varepsilon(x) = \varepsilon^{-2}j(x/\varepsilon)$. Then for any $a > 0$ and $\nu > 0$, we have*

$$(A.3) \quad \xi^\varepsilon(\nu a^2 \cdot, a \cdot) \stackrel{\text{dist}}{=} \frac{1}{a^2 \sqrt{\nu}} \xi_a^\varepsilon(\cdot, \cdot)$$

in the sense that for all $\phi \in C_c^\infty(\mathbb{R} \times \mathbb{R}^2)$,

$$(A.4) \quad \int \phi(t, x) \xi^\varepsilon(\nu a^2 t, ax) dt dx \stackrel{\text{dist}}{=} \frac{1}{a^2 \sqrt{\nu}} \int \phi(t, x) \xi_a^\varepsilon(t, x) dt dx.$$

PROOF. Since both sides of (A.4) are centered normal random variables, it suffices to check that their variances equal. Note that

$$\begin{aligned} X &:= \int_{\mathbb{R} \times \mathbb{R}^2} \phi(t, x) \xi^\varepsilon(a^2 vt, ax) dt dx \\ &= \int_{\mathbb{R} \times \mathbb{R}^2 \times \mathbb{R}^2} \phi(t, x) \varepsilon^{-2} j\left(\frac{ax - y}{\varepsilon}\right) \xi(a^2 vt, y) dt dx dy \\ &= \frac{1}{a^4 v \varepsilon^2} \int_{\mathbb{R} \times \mathbb{R}^2} \left(\int_{\mathbb{R}^2} \phi\left(\frac{\tilde{t}}{a^2 v}, \frac{\tilde{x}}{a}\right) j\left(\frac{\tilde{x} - \tilde{y}}{\varepsilon}\right) d\tilde{x} \right) \xi(\tilde{t}, \tilde{y}) d\tilde{t} d\tilde{y}. \end{aligned}$$

Therefore,

$$\begin{aligned} \text{Var}(X) &= \frac{1}{a^8 v^2 \varepsilon^4} \int_{\mathbb{R} \times \mathbb{R}^2} \left(\int_{\mathbb{R}^2} \phi\left(\frac{\tilde{t}}{a^2 v}, \frac{\tilde{x}}{a}\right) j\left(\frac{\tilde{x} - \tilde{y}}{\varepsilon}\right) d\tilde{x} \right)^2 d\tilde{t} d\tilde{y} \\ &= \frac{1}{v \varepsilon^4} \int_{\mathbb{R} \times \mathbb{R}^2} \left(\int_{\mathbb{R}^2} \phi(t, x) j\left(\frac{a(x - y)}{\varepsilon}\right) dx \right)^2 dt dy. \end{aligned}$$

On the other hand,

$$\begin{aligned} Y &:= \frac{1}{a^2 \sqrt{v}} \int_{\mathbb{R} \times \mathbb{R}^2} \phi(t, x) \xi^{\frac{\varepsilon}{a}}(t, x) dt dx \\ &= \frac{1}{a^2 \sqrt{v}} \int_{\mathbb{R} \times \mathbb{R}^2 \times \mathbb{R}^2} \phi(t, x) \frac{a^2}{\varepsilon^2} j\left(\frac{a(x - y)}{\varepsilon}\right) \xi(t, y) dt dx dy. \end{aligned}$$

Therefore,

$$\text{Var}(Y) = \frac{1}{v \varepsilon^4} \int_{\mathbb{R} \times \mathbb{R}^2} \left(\int_{\mathbb{R}^2} \phi(t, x) j\left(\frac{a(x - y)}{\varepsilon}\right) dx \right)^2 dt dy.$$

Note that the two variances agree, so we are done. \square

PROOF OF PROPOSITION A.1. For $a, b, \tilde{\varepsilon}, \beta > 0$ to be chosen later, define

$$g(t, x) := b \psi^{\tilde{\varepsilon}; 1, 1, \beta^2}(a^2 vt, ax).$$

By (A.1), we have

$$\partial_t \psi^{\tilde{\varepsilon}; 1, 1, \beta^2} = \frac{1}{2} \Delta \psi^{\tilde{\varepsilon}; 1, 1, \beta^2} + \frac{1}{2} |\nabla \psi^{\tilde{\varepsilon}; 1, 1, \beta^2}|^2 + \beta \xi^{\tilde{\varepsilon}}.$$

Therefore,

$$\begin{aligned} \frac{\partial g}{\partial t}(t, x) &= a^2 v b \frac{\partial \psi^{\tilde{\varepsilon}; 1, 1, \beta^2}}{\partial t}(a^2 vt, ax) \\ \text{(A.5)} \quad &= \frac{a^2 v b}{2} \Delta \psi^{\tilde{\varepsilon}; 1, 1, \beta^2}(a^2 vt, ax) + \frac{a^2 v b}{2} |\nabla \psi^{\tilde{\varepsilon}; 1, 1, \beta^2}(a^2 vt, ax)|^2 \\ &\quad + a^2 v b \beta \xi^{\tilde{\varepsilon}}(a^2 vt, ax) \\ &\stackrel{\text{dist}}{=} \frac{v}{2} \Delta g(t, x) + \frac{v}{2b} |\nabla g(t, x)|^2 + b \beta \sqrt{v} \xi^{\frac{\tilde{\varepsilon}}{a}}(t, x), \end{aligned}$$

where we used (A.3).

To find $a, b, \tilde{\varepsilon}$ and β such that g solves (A.1) with parameters v, λ, D , they should satisfy

$$\text{(A.6)} \quad \varepsilon = \frac{\tilde{\varepsilon}}{a}, \quad \lambda = \frac{v}{b}, \quad D = b^2 \beta^2 v.$$

Therefore, we must have

$$(A.7) \quad b = \frac{\nu}{\lambda}, \quad \tilde{\varepsilon} = a\varepsilon, \quad \beta^2 = \frac{D\lambda^2}{\nu^3},$$

while we are free to choose $a > 0$. This proves (A.2). \square

APPENDIX B: HYPERCONTRACTIVITY OF POLYNOMIAL CHAOS

We recall and refine the hypercontractivity property of polynomial chaos established in [38]. Let $(\xi_i)_{i \in \mathbb{T}}$ be i.i.d. random variables, labeled by a countable set \mathbb{T} , with

$$\mathbb{E}[\xi_i] = 0, \quad \mathbb{E}[\xi_i^2] = 1.$$

For every $k \in \mathbb{N}$, let X_k be a multilinear homogeneous polynomial of degree k in the ξ_i 's, that is,

$$(B.1) \quad X_k = \sum_{I \subseteq \mathbb{T}: |I|=k} f_k(I) \prod_{i \in I} \xi_i,$$

where $f_k(I)$ are real coefficients. For $k = 0$, let $X_0 = f_0 \in \mathbb{R}$ be a constant. Then for $k \geq 1$,

$$(B.2) \quad \mathbb{E}[X_k] = 0, \quad \mathbb{E}[X_k^2] = \sum_{I \subseteq \mathbb{T}: |I|=k} f_k(I)^2.$$

If we assume that

$$(B.3) \quad \sum_{k \in \mathbb{N}} \sum_{I \subseteq \mathbb{T}: |I|=k} f_k(I)^2 < \infty,$$

then the series $X := \sum_{k=0}^{\infty} X_k$ is easily seen to define an L^2 random variable. The next key result allows us to control higher moments of X in terms of second moments.

It is useful to allow the law of the $\xi_i = \xi_i^{(N)}$ to depend on a parameter $N \in \mathbb{N}$.

THEOREM B.1 (Hypercontractivity). *For $N \in \mathbb{N}$, let $(\xi_i = \xi_i^{(N)})_{i \in \mathbb{T}}$ be i.i.d. such that*

$$(B.4) \quad \mathbb{E}[\xi_i^{(N)}] = 0, \quad \mathbb{E}[(\xi_i^{(N)})^2] = 1, \quad \exists p_0 \in (2, \infty) : \sup_{N \in \mathbb{N}} \mathbb{E}[|\xi_i^{(N)}|^{p_0}] < \infty.$$

Then, for every $p \in (2, p_0)$, there exists a constant $c_p \in (1, \infty)$ with the following property: for any choice of coefficients $\{f_k(I)\}_{k \in \mathbb{N}, I \subseteq \mathbb{T}, |I|=k}$ satisfying (B.3), if we define X_k by (B.1), then the p th moment of the random variable $X = \sum_{k=0}^{\infty} X_k$ can be bounded as

$$(B.5) \quad \mathbb{E}\left[\left|\sum_{k=0}^{\infty} X_k\right|^p\right] \leq \left(\sum_{k=0}^{\infty} (c_p^k)^2 \mathbb{E}[X_k^2]\right)^{p/2},$$

with $\mathbb{E}[X_k^2]$ given in (B.2). The constant c_p only depends on the laws of $(\xi_i^{(N)})$ and satisfies

$$(B.6) \quad \lim_{p \downarrow 2} c_p = 1.$$

Except for relation (B.6), which we prove below, this theorem was proved in [38] as an extension of the corresponding result in the Gaussian framework; see [30]. In fact, [38], Proposition 3.16, gave the following explicit bound on c_p :

$$c_p \leq \tilde{c}_p = 2\sqrt{p-1} \sup_{N \in \mathbb{N}} \frac{\mathbb{E}[|\xi_i^{(N)}|^p]^{1/p}}{\mathbb{E}[|\xi_i^{(N)}|^2]^{1/2}} = 2\sqrt{p-1} \sup_{N \in \mathbb{N}} \mathbb{E}[|\xi_i^{(N)}|^p]^{1/p},$$

and note that $\lim_{p \downarrow 2} \tilde{c}_p = 2$. This extra factor 2 is the byproduct of a nonoptimal symmetrization argument in the proof in [38]. We now prove (B.6).

PROOF OF EQUATION (B.6). By [38], Section 3.2, relation (B.5) holds with constant c_p if the law of the random variable $\xi = \xi_i$ in (B.1) is $(2, p, 1/c_p)$ -hypercontractive, that is,

$$\forall a \in \mathbb{R} : \left\| a + \frac{1}{c_p} \xi \right\|_p \leq \|a + \xi\|_2,$$

where $\|\cdot\|_p := \mathbb{E}[|\cdot|^p]^{1/p}$ denotes the L^p norm. Since we allow the law of $\xi = \xi^{(N)}$ to depend on $N \in \mathbb{N}$, it follows that we can characterize c_p as follows:

$$(B.7) \quad c_p = \inf \left\{ c > 1 : \left\| a + \frac{1}{c} \xi^{(N)} \right\|_p \leq \|a + \xi^{(N)}\|_2 \quad \forall a \in \mathbb{R}, \forall N \in \mathbb{N} \right\}.$$

For simplicity, we split the proof in two steps.

Step 1. We first consider the case of a fixed law for the random variable ξ (independent of $N \in \mathbb{N}$) satisfying (B.4). In view of (B.7), we can rephrase our goal $\lim_{p \downarrow 2} c_p = 1$ as follows:

$$(B.8) \quad \forall c > 1 \exists p > 2 : \left\| a + \frac{1}{c} \xi \right\|_p \leq \|a + \xi\|_2 \quad \forall a \in \mathbb{R}.$$

We will first show that given $c > 1$, we can find $\bar{p} = \bar{p}_{c,p_0} > 2$ and $K = K_{c,p_0} > 0$, such that for all $p \in (2, \bar{p}]$ and $|a| > K$, the inequality in (B.8) holds. We will then find $p \in (2, \bar{p}]$ such that the inequality in (B.8) also holds for all $|a| \leq K$.

We first need an elementary estimate: for any $p_0 \in (2, \infty)$ there exists $C = C_{p_0} < \infty$ such that, for all $p \in [2, p_0]$ and $x \in \mathbb{R}$,

$$(B.9) \quad |1 + x|^p = 1 + px + \frac{p(p-1)}{2} x^2 + R(x) \quad \text{with } |R(x)| \leq C(|x|^3 \wedge |x|^{p_0}).$$

This follows by Taylor’s formula for $|x| \leq \frac{1}{2}$ (say) and by the triangle inequality for $|x| > \frac{1}{2}$.

We may assume that $p_0 \in (2, 3]$ in (B.4) (just replace p_0 by $p_0 \wedge 3$). Then for every $\delta \in \mathbb{R}$ with $|\delta| \leq 1$ we can bound

$$|R(\delta\xi)| \leq C(|\xi|^3 \wedge |\xi|^{p_0})|\delta|^{p_0} \leq C(1 + |\xi|^{p_0})|\delta|^{p_0}.$$

Since $\mathbb{E}[\xi] = 0$, it follows by (B.4) and (B.9) that for every $\delta \in \mathbb{R}$ with $|\delta| \leq 1$,

$$(B.10) \quad \mathbb{E}[|1 + \delta\xi|^p] = 1 + \frac{p(p-1)}{2} \delta^2 \mathbb{E}[\xi^2] + r(\delta) \quad \text{with } |r(\delta)| \leq C' \delta^{p_0},$$

where $C' = C'_{p_0} := C(1 + \mathbb{E}[|\xi|^{p_0}])$.

Then, as $|\delta| \rightarrow 0$,

$$\|1 + \delta\xi\|_p = 1 + \frac{p-1}{2} \delta^2 \mathbb{E}[\xi^2] + O(|\delta|^{p_0}),$$

uniformly for $p \in [2, p_0]$. This implies that as $|a| \rightarrow \infty$

$$(B.11) \quad \frac{\|a + \frac{1}{c} \xi\|_p}{\|a + \xi\|_2} = \frac{\|1 + \frac{1}{ca} \xi\|_p}{\|1 + \frac{1}{a} \xi\|_2} = \frac{1 + \frac{(p-1)\mathbb{E}[\xi^2]}{2c^2|a|^2} + O(\frac{1}{|a|^{p_0}})}{1 + \frac{\mathbb{E}[\xi^2]}{2|a|^2} + O(\frac{1}{|a|^{p_0}})}$$

$$= 1 + \left\{ \frac{p-1}{c^2} - 1 \right\} \frac{\mathbb{E}[\xi^2]}{2|a|^2} + O\left(\frac{1}{|a|^{p_0}}\right).$$

Given $c > 1$, we can take $\bar{p} = \bar{p}_{c,p_0} := \min\{1 + c, p_0\} > 2$ so that uniformly in $p \in (2, \bar{p}]$, the term in bracket is bounded by $c^{-1} - 1 < 0$. Then the RHS of (B.11) is < 1 for large $|a|$, say

for $|a| > K$, where $K = K_{p_0,c} < \infty$ only depends on c and p_0 . This proves the inequality in (B.8) for all $p \in (2, \bar{p}]$ and $|a| > K$.

To complete the proof, we now show that there exists $p \in (2, \bar{p}]$ such that the inequality in (B.8) holds for $|a| \leq K$. If this is false, then for any sequence $p_n \in (2, p_0]$ with $p_n \downarrow 2$, we can find $a_n \in [-K, K]$ such that

$$(B.12) \quad \left\| a_n + \frac{1}{c} \xi \right\|_{p_n} > \|a_n + \xi\|_2 \quad \forall n \in \mathbb{N}.$$

Extracting subsequences, we may assume that $a_n \rightarrow a \in [-K, K]$. Since $f(p, a) := \|a + \frac{1}{c} \xi\|_p$ is a continuous function of $(p, a) \in [2, p_0] \times [-K, K]$ (by dominated convergence), we may take the limit of (B.12) as $n \rightarrow \infty$ and get

$$(B.13) \quad \left\| a + \frac{1}{c} \xi \right\|_2 \geq \|a + \xi\|_2,$$

which is a contradiction, since $\|a + \frac{1}{c} \xi\|_2 = \sqrt{a^2 + \frac{1}{c^2} \mathbb{E}[\xi^2]} < \|a + \xi\|_2$ (recall that $c > 1$).

Step 2. Next, we allow the law of $\xi = \xi^{(N)}$ to depend on $N \in \mathbb{N}$. In view of (B.7), our goal $\lim_{p \downarrow 2} c_p = 1$ can be rephrased as follows:

$$(B.14) \quad \forall c > 1 \exists p > 2: \left\| a + \frac{1}{c} \xi^{(N)} \right\|_p \leq \|a + \xi^{(N)}\|_2 \quad \forall a \in \mathbb{R}, \forall N \in \mathbb{N}.$$

We follow the same proof as in Step 1. We just need to check the uniformity in $N \in \mathbb{N}$.

Relation (B.10) still holds with ξ replaced by $\xi^{(N)}$, where we stress that $C' = C'_{p_0} < \infty$ because we assume that $\sup_{N \in \mathbb{N}} \mathbb{E}[|\xi^{(N)}|^{p_0}] < \infty$; see (B.4). Then (B.11) holds as $|a| \rightarrow \infty$, uniformly for $p \in [2, p_0]$ and also for $N \in \mathbb{N}$. This proves that (B.14) holds if we restrict $|a| \leq K$, for a suitable $K = K_{p_0,c}$ depending only on $c > 1$ and p_0 .

It remains to fix $c > 1$, $K < \infty$ and prove that (B.14) holds, for some $p > 2$ and for every $|a| \leq K$. Arguing again by contradiction, assume now that there are sequences $p_n \in (2, p_0]$, $a_n \in [-K, K]$, $N_n \in \mathbb{N}$, with $p_n \downarrow 2$, such that

$$(B.15) \quad \left\| a_n + \frac{1}{c} \xi^{(N_n)} \right\|_{p_n} > \|a_n + \xi^{(N_n)}\|_2 \quad \forall n \in \mathbb{N}.$$

Extracting subsequences, we may assume that $a_n \rightarrow a \in [-K, K]$, and also that $\xi^{(N_n)}$ converges in law to a random variable ξ (the sequence is tight, by (B.4)). Since $|a_{N_n} + \frac{1}{c} \xi^{(N_n)}|^{p_n}$ are uniformly integrable, again by (B.4), we can take the limit of relation (B.15) and we get precisely (B.13), which we already showed to be a contradiction. \square

APPENDIX C: GAUSSIAN CONCENTRATION IN THE CONTINUUM

We prove a Gaussian concentration result, based on [34, 35], which can be viewed as a “one-sided version” of [21], Theorem 2 (cf. (C.6) below with equation (4) in [21]).

Given a probability measure μ on a measurable space (E, \mathcal{E}) , we denote by μ_* and μ^* the inner and outer measures: $\mu_*(A) := \sup\{\mu(A') : A' \subseteq A, A' \in \mathcal{E}\}$ and $\mu^*(A) = 1 - \mu_*(A^c)$.

THEOREM C.1. *Let μ be a Gaussian measure on a separable Banach space E , with Cameron–Martin space H .² Let $f : E \rightarrow \mathbb{R} \cup \{-\infty, +\infty\}$ be convex. For $x \in E$ with*

²This means that H is a separable Hilbert space, continuously embedded as a dense subset of the separable Banach space E , and μ is a probability on E that can be described as follows: given any complete orthonormal set $(h_n)_{n \in \mathbb{N}}$ in H and given i.i.d. $N(0, 1)$ random variables $(Z_n)_{n \in \mathbb{N}}$, the sequence of random elements $X_N := \sum_{n=1}^N Z_n h_n$ converges a.s. in the space E , and μ is the law on E of the limit $X := \sum_{n \in \mathbb{N}} Z_n h_n$. The triple (H, E, μ) is called an *abstract Wiener space*. We refer to [34] for more details.

$|f(x)| < \infty$, define the maximal gradient $|\nabla f(x)| \in [0, \infty]$ in the directions of H by (5.18). Then

$$(C.1) \quad \mu(f \leq a - t)\mu^*(f \geq a, |\nabla f| \leq c) \leq e^{-\frac{1}{4}(t/c)^2} \quad \forall a \in \mathbb{R}, \forall t, c \in (0, \infty).$$

(The outer measure μ^* appears in (C.1) to avoid the issue of measurability of $|\nabla f|$.)

Let us denote by $\mathcal{K} := \{h \in H : \|h\|_H \leq 1\}$ the unit ball in the Cameron–Martin space H . Given a subset $A \subseteq E$, we define its enlargement $A + r\mathcal{K} := \{x + rh : x \in A, h \in \mathcal{K}\}$. We recall the classical concentration property established by Borell [34], Theorem 4.3:

$$(C.2) \quad \forall A \subseteq E \text{ with } 0 < \mu(A) < 1, \text{ setting } a := \Phi^{-1}(\mu(A)), \\ \mu_*(A + r\mathcal{K}) \geq \Phi(a + r) \quad \forall r \geq 0,$$

where $\Phi(x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-t^2/2} dt$ is the standard Gaussian distribution function.

The proof of Theorem C.1 is based on the following lemma of independent interest, which follows from (C.2). It is close to [35], Corollary 1.4 (see also [9], Appendix B.1).

LEMMA C.2. For any measurable subset $A \subseteq E$, the following inequality holds:

$$(C.3) \quad \mu(A)(1 - \mu_*(A + r\mathcal{K})) \leq e^{-\frac{1}{4}r^2} \quad \forall r \geq 0.$$

PROOF. We may assume $0 < \mu(A) < 1$ (otherwise (C.3) holds trivially) and we apply (C.2):

$$(C.4) \quad 1 - \mu_*(A + r\mathcal{K}) \leq 1 - \Phi(r + a) \leq e^{-\frac{1}{2}((r+a)^+)^2} \quad \forall r \geq 0,$$

where $x^+ := \max\{x, 0\}$ and we used the basic bound $1 - \Phi(x) \leq e^{-x^2/2}$ for $x \geq 0$.

Consider first the case $\mu(A) \geq \frac{1}{2}$: then $a = \Phi^{-1}(\mu(A)) \geq 0$ and $(r + a)^+ \geq r$, so (C.3) follows by (C.4) (just bound $\mu(A) \leq 1$). Henceforth, we take $\mu(A) < \frac{1}{2}$, so $a < 0$. Note that

$$(C.5) \quad \mu(A) = \Phi(a) = 1 - \Phi(|a|) \leq e^{-\frac{1}{2}|a|^2}.$$

Fix $r \geq 0$. If $|a| \geq r$, then (C.3) follows by (C.5) (just bound $1 - \mu_*(A + r\mathcal{K}) \leq 1$). If $|a| < r$, then $(r + a)^+ = (r - |a|)^+ = r - |a|$ and relations (C.4)–(C.5) yield

$$\mu(A)(1 - \mu_*(A + r\mathcal{K})) \leq e^{-\frac{1}{2}\{|a|^2 + (r - |a|)^2\}} \leq \sup_{x \in \mathbb{R}} e^{-\frac{1}{2}(x^2 + (r - x)^2)} = e^{-\frac{1}{4}r^2}. \quad \square$$

PROOF OF THEOREM C.1. Fix $x, x' \in E$ such that $h := x' - x \in H$. The function $g : [0, 1] \rightarrow \mathbb{R}$ defined by $g(s) := f((1 - s)x + sx')$ is convex (since f is convex), hence

$$f(x') - f(x) = g(1) - g(0) \leq g'(1-) := \lim_{\varepsilon \downarrow 0} \frac{g(1) - g(1 - \varepsilon)}{\varepsilon} = \lim_{\varepsilon \downarrow 0} \frac{f(x') - f(x' - \varepsilon h)}{\varepsilon}.$$

Recalling (5.18), we have shown that

$$(C.6) \quad f(x') - f(x) \leq \|\nabla f(x')\| \|x' - x\|_H.$$

Let us now set

$$A := \{f \leq a - t\}, \quad B := \{f \geq a, |\nabla f| \leq c\}.$$

In view of Lemma C.2, to prove (C.1) it suffices to show that for any $r < \frac{t}{c}$ we have $B \subseteq (A + r\mathcal{K})^c$, that is, $A + r\mathcal{K} \subseteq B^c$. So we fix $x \in A, h \in H$ with $\|h\|_H < \frac{t}{c}$ and we show that $x' := x + h \in B^c$. Either $f(x') < a$, and then $x' \notin B$, or $f(x') \geq a$, and then (by $x \in A$)

$$|\nabla f(x')| \geq \frac{f(x') - f(x)}{\|x' - x\|_H} \geq \frac{a - (a - t)}{\|h\|_H} > \frac{t}{t/c} = c,$$

hence again $x' \notin B$. This completes the proof. \square

APPENDIX D: STOCHASTIC INTEGRAL AS A LINEAR FUNCTION

We formulate a linearity result for the stochastic integral with respect to the white noise $\xi = (\xi(z))_{z \in \mathbb{R}^d}$ on \mathbb{R}^d , which is needed in the proof of Proposition 5.5. Recall that the white noise can be realized as a random element of a separable Banach space E of distributions on \mathbb{R}^d (e.g., a negative Hölder space), equipped with its Borel σ -algebra. Denoting by μ the law of the white noise on E , we will use the probability space (E, μ) as a canonical construction of ξ . We also set $H = L^2(\mathbb{R}^d)$.

For any $h \in H$, the stochastic integral $\langle h, \xi \rangle := \int_{\mathbb{R}^d} h(z)\xi(z) dz \sim N(0, \|h\|_H^2)$ is a random variable in $L^2(E, \mu)$, so it is not canonically defined for any given $\xi \in E$. The following results guarantees the existence of a convenient version of $\langle h, \xi \rangle$.

THEOREM D.1. *It is possible to define $\langle h, \xi \rangle$ as a jointly measurable of $(h, \xi) \in H \times E$, with the following properties:*

- $\langle h, \xi \rangle$ is a version of the stochastic integral $\int_{\mathbb{R}^d} h(z)\xi(z) dz$, for every $h \in H$.
- For any probability measure ν on H , there is a measurable vector space $V_\nu \subseteq E$ with

$$\mu(V_\nu) = 1, \quad V_\nu + H = V_\nu,$$

such that the following property holds:

$$(D.1) \quad \forall \xi, \xi' \in V_\nu : \langle h, \alpha\xi + \alpha'\xi' \rangle = \alpha\langle h, \xi \rangle + \alpha'\langle h, \xi' \rangle < \infty$$

for ν -a.e. $h \in H, \forall \alpha, \alpha' \in \mathbb{R}$.

REMARK D.2. Given any probability ν on H , we can define $f : E \rightarrow \mathbb{R} \cup \{+\infty\}$ by

$$(D.2) \quad f(\xi) := \log \int_H e^{\langle h, \xi \rangle} \nu(dh).$$

This function is convex when restricted to the vector space V_ν of Theorem D.1, by (D.1) and Hölder’s inequality. If we redefine $f(\xi) := +\infty$ for $\xi \notin V_\nu$, we obtain a version of f (recall that $\mu(V_\nu) = 1$) which is convex on the whole space E .

This applies, in particular, to the function $h_\varepsilon(\xi) := \log u_\Lambda^\varepsilon$ in the proof of Proposition 5.5 (see (5.23)), with $u_\Lambda^\varepsilon = u_{\Lambda, \xi}^\varepsilon(0)$ defined in (5.16). In this case, $\mathbb{R}^d = \mathbb{R}^{1+2}$ and ν is the law of the process $(\beta_\varepsilon j(B_s - x))_{(s,x) \in [0, \varepsilon^{-2}] \times \mathbb{R}^2} \in L^2(\mathbb{R}^{1+2})$.

PROOF OF THEOREM D.1. Fix a probability density $\varrho \in C_c^\infty(\mathbb{R}^d)$ and set $\varrho_\varepsilon(z) := \varepsilon^{-d}\varrho(\varepsilon^{-1}z)$. Also fix a smooth cutoff function $\chi : \mathbb{R}^d \rightarrow [0, 1]$ with $\chi(x) \equiv 1$ for $|x| \leq 1$ and $\chi(x) \equiv 0$ for $|x| \geq 2$, and set $\chi_\varepsilon(z) := \chi(\varepsilon z)$. For any $h \in H = L^2(\mathbb{R}^d)$, we define $h_\varepsilon \in C_c^\infty(\mathbb{R}^d)$ by

$$h_\varepsilon(z) := \chi_\varepsilon(z)(\varrho_\varepsilon * h)(z).$$

Since $\lim_{\varepsilon \downarrow 0} \|h_\varepsilon - h\|_H = 0$, we can find $(\varepsilon_n = \varepsilon_n^h)_{n \in \mathbb{N}}$ such that $\|h_{\varepsilon_n} - h\|_H \leq 2^{-n}$. (We can ensure that ε_n^h is measurable in h , for example, $\varepsilon_n^h := \max\{\varepsilon \in \{\frac{1}{k} : k \in \mathbb{N}\} : \|h_\varepsilon - h\|_H \leq 2^{-n}\}$.)

For every $n \in \mathbb{N}$, we have $h_{\varepsilon_n} \in C_c^\infty(\mathbb{R}^d)$; hence the map

$$(D.3) \quad (h, \xi) \mapsto \langle h, \xi \rangle_n := \langle h_{\varepsilon_n}, \xi \rangle$$

is canonically defined for any distribution $\xi \in E$, and is jointly measurable in $(h, \xi) \in H \times E$. By the Itô isometry of the stochastic integral and Borel–Cantelli, for any fixed $h \in H$ we have $\lim_{n \rightarrow \infty} \langle h, \xi \rangle_n = \langle h, \xi \rangle$ for μ -a.e. $\xi \in E$. We can finally define the measurable map

$$\langle h, \xi \rangle := \begin{cases} \lim_{n \rightarrow \infty} \langle h, \xi \rangle_n & \text{if the limit exists in } \mathbb{R}, \\ +\infty & \text{otherwise.} \end{cases}$$

For every $n \in \mathbb{N}$, the maps $\xi \mapsto \langle h, \xi \rangle_n$ are linear, hence the limit map $\xi \mapsto \langle h, \xi \rangle$ is linear, too, whenever it is finite. More precisely, for every $h \in H$ and $\xi, \zeta \in E$:

$$(D.4) \quad \begin{aligned} \langle h, \xi \rangle < \infty, \quad \langle h, \zeta \rangle < \infty &\implies \\ \langle h, \alpha\xi + \beta\zeta \rangle = \alpha\langle h, \xi \rangle + \beta\langle h, \zeta \rangle < \infty &\quad \forall \alpha, \beta \in \mathbb{R}. \end{aligned}$$

By construction, for every $h \in H$ we have $\langle h, \xi \rangle \in L^2(E, \mu)$, so $\langle h, \xi \rangle < \infty$ for μ -a.e. $\xi \in E$. If we now fix a probability ν on H , and we define the measurable subset $V_\nu \subseteq E$ by

$$V_\nu := \{ \xi \in E : \langle h, \xi \rangle < \infty \text{ for } \nu\text{-a.e. } h \in H \},$$

it follows by Fubini's theorem that $\mu(V_\nu) = 1$. Note that $V_\nu + H = V_\nu$, because $\langle h, g \rangle < \infty$ for all $h, g \in H$. Finally, relation (D.4) implies (D.1), which shows that V_ν is a vector space. \square

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